Vitamin supplementation for preventing miscarriage (Review)

Balogun, Olukunmi O.; da Silva Lopes, Katharina; Ota, Erika; Takemoto, Yo; Rumbold, Alice; Takegata, Mizuki; Mori, Rintaro

Published in:
Cochrane Database of Systematic Reviews

DOI:
10.1002/14651858.CD004073.pub4

Published: 06/05/2016

Document Version
Publisher's PDF, also known as Version of record

Link to publication

Citation for published version (APA):

General rights
Copyright and moral rights for the publications made accessible in the public portal are retained by the authors and/or other copyright owners and it is a condition of accessing publications that users recognise and abide by the legal requirements associated with these rights.

• Users may download and print one copy of any publication from the public portal for the purpose of private study or research.
• You may not further distribute the material or use it for any profit-making activity or commercial gain
• You may freely distribute the URL identifying the publication in the public portal

Take down policy
If you believe that this document breaches copyright please contact us providing details, and we will remove access to the work immediately and investigate your claim.
Vitamin supplementation for preventing miscarriage (Review)

Balogun OO, da Silva Lopes K, Ota E, Takemoto Y, Rumbold A, Takegata M, Mori R

Vitamin supplementation for preventing miscarriage.
DOI: 10.1002/14651858.CD004073.pub4.

www.cochranelibrary.com
## Table of Contents

<table>
<thead>
<tr>
<th>Section</th>
<th>Page</th>
</tr>
</thead>
<tbody>
<tr>
<td>HEADER</td>
<td>1</td>
</tr>
<tr>
<td>ABSTRACT</td>
<td>1</td>
</tr>
<tr>
<td>PLAIN LANGUAGE SUMMARY</td>
<td>2</td>
</tr>
<tr>
<td>SUMMARY OF FINDINGS FOR THE MAIN COMPARISON</td>
<td>4</td>
</tr>
<tr>
<td>BACKGROUND</td>
<td>6</td>
</tr>
<tr>
<td>OBJECTIVES</td>
<td>7</td>
</tr>
<tr>
<td>METHODS</td>
<td>7</td>
</tr>
<tr>
<td>RESULTS</td>
<td>11</td>
</tr>
<tr>
<td>............................... Figure 1. ........................................</td>
<td>13</td>
</tr>
<tr>
<td>............................... Figure 2. ........................................</td>
<td>14</td>
</tr>
<tr>
<td>ADDITIONAL SUMMARY OF FINDINGS</td>
<td>19</td>
</tr>
<tr>
<td>DISCUSSION</td>
<td>24</td>
</tr>
<tr>
<td>............................... Figure 3. ........................................</td>
<td>25</td>
</tr>
<tr>
<td>............................... Figure 4. ........................................</td>
<td>26</td>
</tr>
<tr>
<td>............................... Figure 5. ........................................</td>
<td>27</td>
</tr>
<tr>
<td>AUTHORS’ CONCLUSIONS</td>
<td>28</td>
</tr>
<tr>
<td>ACKNOWLEDGEMENTS</td>
<td>28</td>
</tr>
<tr>
<td>REFERENCES</td>
<td>29</td>
</tr>
<tr>
<td>CHARACTERISTICS OF STUDIES</td>
<td>41</td>
</tr>
<tr>
<td>DATA AND ANALYSES</td>
<td>118</td>
</tr>
<tr>
<td>WHAT’S NEW</td>
<td>124</td>
</tr>
<tr>
<td>HISTORY</td>
<td>125</td>
</tr>
<tr>
<td>CONTRIBUTIONS OF AUTHORS</td>
<td>125</td>
</tr>
<tr>
<td>DECLARATIONS OF INTEREST</td>
<td>126</td>
</tr>
<tr>
<td>SOURCES OF SUPPORT</td>
<td>126</td>
</tr>
<tr>
<td>DIFFERENCES BETWEEN PROTOCOL AND REVIEW</td>
<td>126</td>
</tr>
<tr>
<td>INDEX TERMS</td>
<td>128</td>
</tr>
</tbody>
</table>
Vitamin supplementation for preventing miscarriage

Olukunmi O Balogun1, Katharina da Silva Lopes1, Erika Ota2, Yo Takemoto3, Alice Rumbold4, Mizuki Takegata1, Rintaro Mori1

1Department of Health Policy, National Center for Child Health and Development, Tokyo, Japan. 2Global Health Nursing, St. Luke’s International University, Graduate School of Nursing Sciences, Tokyo, Japan. 3National Research Institute for Child Health and Development, Tokyo, Japan. 4The Robinson Research Institute, The University of Adelaide, Adelaide, Australia

Contact address: Erika Ota, Global Health Nursing, St. Luke’s International University, Graduate School of Nursing Sciences, 10-1 Akashi-cho, Chuo-Ku, Tokyo, 104-0044, Japan. e-i@umin.ac.jp.

Editorial group: Cochrane Pregnancy and Childbirth Group.
Publication status and date: New search for studies and content updated (no change to conclusions), published in Issue 5, 2016.
Review content assessed as up-to-date: 6 November 2015.


Copyright © 2016 The Cochrane Collaboration. Published by John Wiley & Sons, Ltd.

ABSTRACT

Background

Miscarriage is a common complication of pregnancy that can be caused by a wide range of factors. Poor dietary intake of vitamins has been associated with an increased risk of miscarriage; therefore supplementing women with vitamins either prior to or in early pregnancy may help prevent miscarriage.

Objectives

The objectives of this review were to determine the effectiveness and safety of any vitamin supplementation, on the risk of spontaneous miscarriage.

Search methods

We searched the Cochrane Pregnancy and Childbirth Group Trials Register (6 November 2015) and reference lists of retrieved studies.

Selection criteria

All randomised and quasi-randomised trials comparing supplementation during pregnancy with one or more vitamins with either placebo, other vitamins, no vitamins or other interventions. We have included supplementation that started prior to conception, periconceptionally or in early pregnancy (less than 20 weeks’ gestation).

Data collection and analysis

Three review authors independently assessed trials for inclusion, extracted data and assessed trial quality. We assessed the quality of the evidence using the GRADE approach. The quality of evidence is included for numerical results of outcomes included in the ‘Summary of findings’ tables.

Main results

We included a total of 40 trials (involving 276,820 women and 278,413 pregnancies) assessing supplementation with any vitamin(s) starting prior to 20 weeks’ gestation and reporting at least one primary outcome that was eligible for the review. Eight trials were cluster-randomised and contributed data for 217,726 women and 219,267 pregnancies in total.
Approximately half of the included trials were assessed to have a low risk of bias for both random sequence generation and adequate concealment of participants to treatment and control groups.

**Vitamin C supplementation**

There was no difference in the risk of total fetal loss (risk ratio (RR) 1.14, 95% confidence interval (CI) 0.92 to 1.40, seven trials, 18,949 women; *high-quality evidence*); early or late miscarriage (RR 0.90, 95% CI 0.65 to 1.26, four trials, 13,346 women; *moderate-quality evidence*); stillbirth (RR 1.31, 95% CI 0.97 to 1.76, seven trials, 21,442 women; *moderate-quality evidence*) or adverse effects of vitamin supplementation (RR 1.16, 95% CI 0.39 to 3.41, one trial, 739 women; *moderate-quality evidence*) between women receiving vitamin C with vitamin E compared with placebo or no vitamin C groups. No clear differences were seen in the risk of total fetal loss or miscarriage between women receiving any other combination of vitamin C compared with placebo or no vitamin C groups.

**Vitamin A supplementation**

No difference was found in the risk of total fetal loss (RR 1.01, 95% CI 0.61 to 1.66, three trials, 1640 women; *low-quality evidence*); early or late miscarriage (RR 0.86, 95% CI 0.46 to 1.62, two trials, 1397 women; *low-quality evidence*) or stillbirth (RR 1.29, 95% CI 0.57 to 2.91, three trials, 1640 women; *low-quality evidence*) between women receiving vitamin A plus iron and folate compared with placebo or no vitamin A groups. There was no evidence of differences in the risk of total fetal loss or miscarriage between women receiving any other combination of vitamin A compared with placebo or no vitamin A groups.

**Multivitamin supplementation**

There was evidence of a decrease in the risk for stillbirth among women receiving multivitamins plus iron and folic acid compared iron and folate only groups (RR 0.92, 95% CI 0.85 to 0.99, 10 trials, 79,851 women; *high-quality evidence*). Although total fetal loss was lower in women who were given multivitamins without folic acid (RR 0.49, 95% CI 0.34 to 0.70, one trial, 907 women); and multivitamins with or without vitamin A (RR 0.60, 95% CI 0.39 to 0.92, one trial, 1074 women), these findings included one trial each with small numbers of women involved. Also, they include studies where the comparison groups included women receiving either vitamin A or placebo, and thus require caution in interpretation.

We found no difference in the risk of total fetal loss (RR 0.96, 95% CI 0.93 to 1.00, 10 trials, 94,948 women; *high-quality evidence*) or early or late miscarriage (RR 0.98, 95% CI 0.94 to 1.03, 10 trials, 94,948 women; *moderate-quality evidence*) between women receiving multivitamins plus iron and folic acid compared with iron and folate only groups.

There was no evidence of differences in the risk of total fetal loss or miscarriage between women receiving any other combination of multivitamins compared with placebo, folic acid or vitamin A groups.

**Folic acid supplementation**

There was no evidence of any difference in the risk of total fetal loss, early or late miscarriage, stillbirth or congenital malformations between women supplemented with folic acid with or without multivitamins and/or iron compared with no folic acid groups.

**Antioxidant vitamins supplementation**

There was no evidence of differences in early or late miscarriage between women given antioxidant compared with the low antioxidant group (RR 1.12, 95% CI 0.24 to 5.29, one trial, 110 women).

**Authors’ conclusions**

Taking any vitamin supplements prior to pregnancy or in early pregnancy does not prevent women experiencing miscarriage. However, evidence showed that women receiving multivitamins plus iron and folic acid had reduced risk for stillbirth. There is insufficient evidence to examine the effects of different combinations of vitamins on miscarriage and miscarriage-related outcomes.

**Plain Language Summary**

**Vitamin supplementation for preventing miscarriage**

**What is the issue?**

Miscarriage occurs frequently among pregnant women but it is often difficult to know the factors responsible. Poor diet, without enough vitamins, has been associated with an increased risk of women losing their baby in early pregnancy. Does vitamin supplementation
taken by women before pregnancy and during pregnancy decrease the risk of spontaneous miscarriage? Does supplementation improve maternal, birth and infant outcomes, and are there any side effects?

Why is this important?

Vitamin supplementation is commonly recommended for pregnant women and women planning to conceive. Considering the widespread use of vitamin supplementation before and during pregnancy, it is important to study the relation between vitamin supplementation and early pregnancy outcomes, particularly since the causes of miscarriage are unknown and the nutritional status of a mother can affect her baby's development.

What evidence did we find?

This review included 40 randomised controlled trials involving 276,820 women and 278,413 pregnancies. Supplementing women with any vitamins does not reduce the number of women who have miscarriages. However, the risk for stillbirth was reduced among women receiving multivitamins plus iron and folic acid compared with iron and folate only groups. Although total fetal loss was lower in women who were given multivitamins without folic acid and multivitamins with or without vitamin A, these findings included one trial each with small numbers of women involved. Also, they include studies where the comparison groups included women receiving either vitamin A or placebo, and thus require caution in interpretation.

What does this mean?

Taking vitamin supplements before pregnancy or in early pregnancy may be beneficial; but this review did not show sufficient evidence that taking vitamin supplements prevents miscarriage.
### Summary of Findings for the Main Comparison

**Vitamin C plus vitamin E versus control for preventing miscarriage**

- **Population:** pregnant women
- **Setting:** Australia, Brazil, Canada, Scotland, UK, USA
- **Intervention:** vitamin C and vitamin E
- **Comparison:** placebo

<table>
<thead>
<tr>
<th>Outcomes</th>
<th>Anticipated absolute effects* (95% CI)</th>
<th>Relative effect (95% CI)</th>
<th>n of participants (studies)</th>
<th>Quality of the evidence (GRADE)</th>
<th>Comments</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Risk with placebo</td>
<td>Risk with Vitamin C plus vitamin E</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Total fetal loss</td>
<td>Study population</td>
<td>RR 1.14 (0.92 to 1.40)</td>
<td>18,949 (7 RCTs)</td>
<td>⊕⊕⊕⊕ HIGH ¹</td>
<td></td>
</tr>
<tr>
<td></td>
<td>17 per 1000</td>
<td>20 per 1000 (16 to 24)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>Moderate</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>14 per 1000</td>
<td>16 per 1000 (13 to 20)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Early or late miscarriage</td>
<td>Study population</td>
<td>RR 0.90 (0.65 to 1.26)</td>
<td>13,346 (4 RCTs)</td>
<td>⊕⊕⊕⊕ MODERATE ²</td>
<td></td>
</tr>
<tr>
<td></td>
<td>9 per 1000</td>
<td>9 per 1000 (7 to 12)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>Moderate</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>8 per 1000</td>
<td>8 per 1000 (6 to 11)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Stillbirth</td>
<td>Study population</td>
<td>RR 1.31 (0.97 to 1.76)</td>
<td>21,442 (7 RCTs)</td>
<td>⊕⊕⊕⊕ MODERATE ²</td>
<td></td>
</tr>
</tbody>
</table>

Copyright © 2016 The Cochrane Collaboration. Published by John Wiley & Sons, Ltd.
### Any adverse effects of vitamin supplementation sufficient to stop supplementation

<table>
<thead>
<tr>
<th>Study population</th>
<th>RR 1.16 (0.39 to 3.41)</th>
<th>739 (1 RCT)</th>
<th>MODERATE 23</th>
</tr>
</thead>
<tbody>
<tr>
<td>16 per 1000</td>
<td>19 per 1000 (6 to 56)</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

- **The risk in the intervention group** (and its 95% confidence interval) is based on the assumed risk in the comparison group and the **relative effect** of the intervention (and its 95% CI).

**CI:** Confidence interval; **RR:** Risk ratio

### GRADE Working Group grades of evidence

- **High quality:** We are very confident that the true effect lies close to that of the estimate of the effect.
- **Moderate quality:** We are moderately confident in the effect estimate: The true effect is likely to be close to the estimate of the effect, but there is a possibility that it is substantially different.
- **Low quality:** Our confidence in the effect estimate is limited: The true effect may be substantially different from the estimate of the effect.
- **Very low quality:** We have very little confidence in the effect estimate: The true effect is likely to be substantially different from the estimate of effect.

---
1. Not effective but 95% CI is narrow and precise.
2. Non significant with wide 95% CI.
3. Small sample size.
BACKGROUND

Description of the condition

Miscarriage can be caused by a wide range of factors, and determining the aetiology is often difficult given the variety of underlying mechanisms potentially responsible. Consideration of the timing of the miscarriage is also important, as diverse causes of miscarriage manifest at different periods of gestation. The most common causes include abnormal chromosomal rearrangements, endocrinological disorders and uterine abnormalities (Garrido-Gimenez 2015). Early miscarriages are mostly associated with chromosomal abnormalities, defective placental development and maternal disease conditions; while late miscarriages are more likely due to structural problems of the uterus and/or cervix such as cervical incompetence. Women experiencing recurrent miscarriage often have an underlying medical condition such as autoimmune disease, i.e. systemic lupus erythematosus and antiphospholipid syndrome, or other blood clotting disorders such as hyperhomocysteinaemia (high levels of homocysteine in the blood) or another thrombophilia (Preston 1996). Other risk factors for miscarriage include higher maternal age at conception, multiple pregnancies and a history of previous miscarriage (Baba 2011; Garcia-Enguadanos 2002; Hure 2012). Behavioural factors including alcohol consumption (Maconochie 2007), smoking (Baba 2011; Hure 2012), use of illicit drugs (Garcia-Enguadanos 2002), and exposure to non-steroidal anti-inflammatory drugs (NSAIDs) around the time of conception are also suggested causes of miscarriage (Li 2003; Nielsen 2001). While several factors may promote miscarriage, for a great proportion of women, no cause can be found.

In clinical practice, surgical and non-surgical interventions are used in the management of miscarriage. Bed rest, commonly prescribed for preventing miscarriage, is lacking proven benefit (Aleman 2005). Similarly, there is currently insufficient evidence on the benefit provided from the use of uterine muscle relaxant drugs (Lede 2005), Chinese herbal medicine (Li 2012), hormones (Devaselvan 2010; Haas 2013; Lim 2013; Morley 2013), and immunotherapy (Wong 2014).

Description of the intervention

Vitamins are essential nutrients required in the body for numerous functions such as to ensure normal metabolism, physical growth and development as well as to prevent diseases. Based on evidence from observational studies, vitamin supplementation has been advocated for the prevention of miscarriage (Hasan 2009; Maconochie 2007), most commonly folate and B vitamins. Due to consistent associations between pregnancy complications and decreased antioxidant defence and infections, it has been suggested that vitamin supplementation during pregnancy might provide protection against adverse pregnancy outcomes and may influence the risk of spontaneous miscarriage in women.

How the intervention might work

Vitamins are either water soluble - such as vitamin C and the B group vitamins (including folate) or fat soluble such as vitamins A, D, E and K. They may be obtained directly from the diet or in the form of dietary supplements of either individual vitamin or multivitamin preparations. Multivitamins contain a range of vitamins and minerals, usually in doses similar to the recommended dietary intakes.

The rationale for vitamin supplementation for the prevention of miscarriage is based on epidemiological studies linking healthy dietary patterns with reduced risk for miscarriage (Hasan 2009; Maconochie 2007). Several studies have reported an association between certain vitamin deficiencies and adverse reproductive outcomes (George 2002; Guerra-Shinohara 2010; Hübner 2008; Reznikoff-Etiévant 2002). Nutritional mechanisms underlying this association include homocysteine metabolism and oxidative stress.

Homocysteine is an amino acid that is involved in several key metabolic processes, vital to ongoing cellular activity of the living organism. The metabolism of homocysteine is facilitated by B vitamins and folate. The concentration of homocysteine in the blood is determined by various dietary factors, including folate, vitamin B6 and vitamin B12. Disturbance of maternal and fetal homocysteine metabolism has been associated with various obstetric conditions including miscarriage (Hague 2003; Nelen 2000), and hyperhomocysteinaemia is considered a risk factor for recurrent early pregnancy loss. Therefore, supplementation with B vitamins and folate may influence the risk of spontaneous miscarriage in women with recurrent miscarriage. Moreover, low serum vitamin B12 concentrations have been reported in women with recurrent miscarriage (Hübner 2008). Evidence on the effect of vitamin supplementation, particularly folic acid, on risk of miscarriage is still conflicting; however the few studies that have adjusted for confounding support a protective effect.

Oxidative stress is caused by an imbalance between pro-oxidants and antioxidants. Pro-oxidants act either by generating reactive oxygen species (ROS) or by inhibiting antioxidant systems. In living cells, ROS are formed continuously both from biochemical processes occurring in the body and in reaction to external factors. Excessive ROS production may however, overpower the body’s natural antioxidant defence system, creating an environment unsuitable for normal female reproductive processes (Al-Gubory 2010). A recent review of evidence from experimental and observational studies suggests that oxidative stress is an important cause in spontaneous and recurrent miscarriage (Agarwal 2012b; Al-Gubory 2010; Gupta 2007). Adequate maternal antioxidant status before and during pregnancy could prevent and control oxidative stress. Therefore, intake of antioxidant vitamins...
such as vitamin C and vitamin E may be an important factor to reduce the risk of miscarriage. In a population-based case-control study, vitamin supplementation (including vitamin C), and eating fresh fruits and vegetables daily were associated with reduced risk of miscarriage (Maconochie 2007). Another observational study demonstrated an association between the risk of spontaneous early miscarriage and dietary factors; poor intake of green vegetables, fruit and dairy products coupled with a high intake of fat was associated with a high risk of spontaneous miscarriage (Di Cintio 2001). There is limited information available about the impact of vitamins on the risk of early versus late miscarriage. However, dietary factors could theoretically influence structural abnormalities such as cervical incompetence. There is a growing body of research investigating the relationship between nutrition and placental development, fetal growth, pregnancy outcomes and adult diseases (McMillen 2008; Meher 2015; Wu 2004). Therefore, adequate maternal nutrition, particularly vitamin intake, may be an important factor in preventing spontaneous miscarriage. Currently, little information is available about the most appropriate vitamin type or combination. Similarly, many commercially available vitamin preparations contain a range of combinations of vitamins. Therefore, this review will cover all vitamin types.

**Why it is important to do this review**

Vitamin supplementation is frequently recommended for pregnant women and women planning to conceive. The documented benefits of supplementation relate mainly to the lowered risk of congenital anomalies such as neural tube defects (Hovdenak 2012; MRC Vitamin Study Research Group 1991). Given the widespread vitamin supplementation before and during pregnancy, studying the relationship between this common exposure and early pregnancy outcomes is of great value, particularly since the causes of miscarriage are unknown and this exposure is known to affect specific developmental processes. This is an update of a Cochrane review first published in 2005 and previously updated in 2011. The previously updated review included 28 trials involving 96,674 women (98,267 pregnancies (Rumbold 2011)). Based on available evidence, Rumbold 2011 concluded that taking any vitamin supplements prior to pregnancy or in early pregnancy does not prevent women from experiencing miscarriage or stillbirth. However, there is insufficient evidence to examine the effects of different combinations of vitamins on miscarriage. In the current review, we examined the effect of different vitamin combinations on the risk of miscarriage. The scope of the current update has been restricted to look at miscarriage and miscarriage-related outcomes.

**OBJECTIVES**

1. To determine the effectiveness and safety of any vitamin supplementation taken by women prior to conception, periconceptionally and in early pregnancy on the risk of spontaneous miscarriage.

2. If vitamins are effective, to determine which of these agents are best and to compare vitamins with other interventions.

**METHODS**

**Criteria for considering studies for this review**

**Types of studies**

All randomised trials (including individual- and cluster-randomised) and quasi-randomised trials comparing one or more vitamins with either placebo, other vitamins, no vitamins or other interventions, prior to conception, periconceptionally or in early to mid pregnancy. Cross-over trials were not included.

**Types of participants**

Pregnant women (less than 20 weeks’ gestation) or women in the reproductive age group planning on becoming pregnant in the near future, regardless of whether they are at low or high risk of having a miscarriage. No restrictions were placed on the age of participants or past obstetric history.

**Types of interventions**

Comparisons of specific vitamin(s), alone or in combination with other agents with either placebo, other vitamin(s), no vitamin(s) or other interventions for the prevention of miscarriage, either in areas where there is an inadequate dietary intake or where there is a presumed adequate intake of that vitamin(s). The review authors deemed it important to include any supplementation trials, where supplementation began prior to 20 weeks’ gestation, and where at least one miscarriage-related outcome as specified in the review was reported, even if the intervention was not specifically for the prevention of miscarriage. We excluded trials where the onset of supplementation occurred definitely after 20 weeks’ gestation or where it was reported that the majority of women commenced supplementation after 20 weeks’ gestation. We included trials where the onset of supplementation occurred both prior to and after 20 weeks’ gestation, and when it could not be established whether the majority of the women started supplementation prior to 20 weeks’ gestation.
**Types of outcome measures**
The scope of the current update has been restricted to look at miscarriage and miscarriage-related outcomes.

**Primary outcomes**
1. Total fetal loss, defined as the combined numbers of early miscarriage (spontaneous pregnancy loss less than 12 weeks' gestation), late miscarriage (spontaneous pregnancy loss greater than or equal to 12 and less than 24 weeks), and stillbirth (pregnancy loss at greater than or equal to 24 weeks).
2. Early or late miscarriage.

To overcome wide variation in the definitions of miscarriage and stillbirth between studies, we included the combined outcome 'total fetal loss' in the review.

**Secondary outcomes**
1. Stillbirth.
2. Congenital malformations.
3. Adverse effects of vitamin supplementation sufficient to stop supplementation, such as manifestations of hypervitaminosis, headache, nausea, vomiting, diarrhoea

**Search methods for identification of studies**
The following methods section of this review is based on a standard template used by the Cochrane Pregnancy and Childbirth Group.

**Electronic searches**
We searched the Cochrane Pregnancy and Childbirth Group's Trials Register by contacting the Trials Search Co-ordinator (6 November 2015).

The Register is a database containing over 21,000 reports of controlled trials in the field of pregnancy and childbirth. For full search methods used to populate the Pregnancy and Childbirth Group's Trials Register including the detailed search strategies for CENTRAL, MEDLINE, Embase and CINAHL; the list of hand-searched journals and conference proceedings, and the list of journals reviewed via the current awareness service, please follow this link to the editorial information about the Cochrane Pregnancy and Childbirth Group in The Cochrane Library and select the 'Specialized Register' section from the options on the left side of the screen.

Briefly, the Cochrane Pregnancy and Childbirth Group's Trials Register is maintained by the Trials Search Co-ordinator and contains trials identified from:
1. monthly searches of the Cochrane Central Register of Controlled Trials (CENTRAL);
2. weekly searches of MEDLINE (Ovid);
3. weekly searches of Embase (Ovid);
4. monthly searches of CINAHL (EBSCO);
5. handsearches of 30 journals and the proceedings of major conferences;
6. weekly current awareness alerts for a further 44 journals plus monthly BioMed Central email alerts.

Search results are screened by two people and the full text of all relevant trial reports identified through the searching activities described above is reviewed. Based on the intervention described, each trial report is assigned a number that corresponds to a specific Pregnancy and Childbirth Group review topic (or topics), and is then added to the Register. The Trials Search Co-ordinator searches the Register for each review using this topic number rather than keywords. This results in a more specific search set which has been fully accounted for in the relevant review sections (Included studies; Excluded studies; Studies awaiting classification; Ongoing studies).

[We carried out additional author searching in an earlier version of this review (Rumbold 2005), see Appendix 1 for details]

**Searching other resources**
We searched the reference lists of retrieved studies. We did not apply any language or date restrictions.

**Data collection and analysis**
For methods used in the previous version of this review, see Rumbold 2011.

For this update, the following methods were used for assessing the 90 reports that were identified as a result of the updated search. The following methods section of this review is based on a standard template used by the Cochrane Pregnancy and Childbirth Group.

**Selection of studies**
Two review authors independently assessed all the potential studies identified as a result of the search strategy for inclusion. Disagreements were resolved through discussion and, when required, we consulted a third person. We created a study flow diagram to map out the number of records identified, included and excluded.

**Data extraction and management**
We collected data from the selected studies using a pre-designed data extraction form. For eligible studies, two review authors extracted the data using the agreed form. We resolved discrepancies through discussion. If discrepancies could not be resolved, we consulted a third review author. We entered data into Review Manager software (RevMan 2014) and checked for accuracy. When information regarding any of the above was unclear, we attempted to contact authors of the original reports to provide further details.
**Assessment of risk of bias in included studies**

Two review authors independently assessed risk of bias for each study using the criteria outlined in the *Cochrane Handbook for Systematic Reviews of Interventions* (Higgins 2011). We resolved any disagreement by discussion or by involving a third assessor.

(1) **Random sequence generation (checking for possible selection bias)**

We describe for each included study the method used to generate the allocation sequence in sufficient detail to allow an assessment of whether it should produce comparable groups.

We assessed the method as:
- low risk of bias (any truly random process, e.g. random number table; computer random number generator);
- high risk of bias (any non-random process, e.g. odd or even date of birth; hospital or clinic record number);
- unclear risk of bias.

(2) **Allocation concealment (checking for possible selection bias)**

We describe for each included study the method used to conceal allocation to interventions prior to assignment and assessed whether intervention allocation could have been foreseen in advance of, or during recruitment, or changed after assignment.

We assessed the methods as:
- low risk of bias (e.g. telephone or central randomisation; consecutively numbered sealed opaque envelopes);
- high risk of bias (open random allocation; unsealed or non-opaque envelopes, alternation; date of birth);
- unclear risk of bias.

(3.1) **Blinding of participants and personnel (checking for possible performance bias)**

We describe for each included study the methods used, if any, to blind study participants and personnel from knowledge of which intervention a participant received. We considered that studies were at low risk of bias if they were blinded, or if we judged that the lack of blinding would be unlikely to affect results. We assessed blinding separately for different outcomes or classes of outcomes.

We assessed the methods as:
- low, high or unclear risk of bias for participants;
- low, high or unclear risk of bias for personnel.

(3.2) **Blinding of outcome assessment (checking for possible detection bias)**

We describe for each included study the methods used, if any, to blind outcome assessors from knowledge of which intervention a participant received. We assessed blinding separately for different outcomes or classes of outcomes.

We assessed methods used to blind outcome assessment as:
- low, high or unclear risk of bias.

(4) **Incomplete outcome data (checking for possible attrition bias due to the amount, nature and handling of incomplete outcome data)**

We describe for each included study, and for each outcome or class of outcomes, the completeness of data including attrition and exclusions from the analysis. We state whether attrition and exclusions were reported and the numbers included in the analysis at each stage (compared with the total randomised participants), reasons for attrition or exclusion where reported, and whether missing data were balanced across groups or were related to outcomes. Where sufficient information was reported, or supplied by the trial authors, we re-included missing data in the analyses which we undertook.

We assessed methods as:
- low risk of bias (e.g. no missing outcome data; missing outcome data balanced across groups);
- high risk of bias (e.g. numbers or reasons for missing data imbalanced across groups; ‘as treated’ analysis done with substantial departure of intervention received from that assigned at randomisation);
- unclear risk of bias.

(5) **Selective reporting (checking for reporting bias)**

We describe for each included study how we investigated the possibility of selective outcome reporting bias and what we found.

We assessed the methods as:
- low risk of bias (where it is clear that all of the study’s pre-specified outcomes and all expected outcomes of interest to the review have been reported);
- high risk of bias (where not all the study’s pre-specified outcomes have been reported; one or more reported primary outcomes were not pre-specified; outcomes of interest are reported incompletely and so cannot be used; study fails to include results of a key outcome that would have been expected to have been reported);
- unclear risk of bias.

(6) **Other bias (checking for bias due to problems not covered by (1) to (5) above)**

We describe for each included study any important concerns we have about other possible sources of bias.

We assessed whether each study was free of other problems that could put it at risk of bias:
- low risk of other bias;
- high risk of other bias;
- unclear whether there is risk of other bias.
Overall risk of bias
We made judgements about whether studies were at high risk of bias, according to the criteria given in the Handbook (Higgins 2011). With reference to (1) to (6) above, we assessed the likely magnitude and direction of the bias and whether we considered it was likely to impact on the findings. We explored the impact of the level of bias through undertaking sensitivity analyses - see Sensitivity analysis.

Assessment of the quality of the evidence using the GRADE approach
For this update, we assessed the quality of the evidence using the GRADE approach as outlined in the GRADE handbook in order to assess the quality of the body of evidence relating to the following outcomes in the comparisons: 1) vitamin C and vitamin E versus placebo, 2) vitamin A plus iron plus folate versus iron plus folate and 3) multivitamin plus iron plus folate versus iron plus folate.
1. Total fetal loss, defined as the combined numbers of early miscarriage (spontaneous pregnancy loss less than 12 weeks’ gestation), late miscarriage (spontaneous pregnancy loss greater than or equal to 12 and less than 24 weeks), and stillbirth (pregnancy loss at greater than or equal to 24 weeks).
2. Early or late miscarriage.
4. Adverse effects of vitamin supplementation sufficient to stop supplementation.
We used the GRADEpro Guideline Development Tool to import data from Review Manager 5.3 (RevMan 2014) in order to create ‘Summary of findings’ tables. A summary of the intervention effect and a measure of quality for each of the above outcomes was produced using the GRADE approach. The GRADE approach uses five considerations (study limitations, consistency of effect, imprecision, indirectness and publication bias) to assess the quality of the body of evidence for each outcome. The evidence was downgraded from ‘high quality’ by one level for serious (or by two levels for very serious) limitations, depending on assessments for risk of bias, indirectness of evidence, serious inconsistency, imprecision of effect estimates or potential publication bias.

Measures of treatment effect
Dichotomous data
For dichotomous data, we present results as summary risk ratio (RR) with 95% confidence intervals (CI).

Continuous data
For continuous data, we planned to use the mean difference (MD) if outcomes were measured in the same way between trials and, if appropriate, the standardised mean difference (SMD) to combine trials that measured the same outcome, but used different methods.

Unit of analysis issues
Where trials recruited women prior to becoming pregnant, we reported the denominators for each trial as all women randomised; or where there was accurate information about the number of women in each trial who became pregnant, we reported the denominators as the number of women randomised with confirmed pregnancy.
We included all included trials in the initial analysis which we performed by any vitamin to include all vitamin combinations and then by individual vitamin type.

Cluster-randomised trials
We included cluster-randomised trials in the analyses along with individually-randomised trials. We used the effect estimates and uncertainty range from the cluster trials to perform the meta-analysis using the generic inverse variance approach for the meta-analysis of dichotomous outcomes where trials using cluster-randomisation techniques were included (Alderson 2004).

Dealing with missing data
For included studies, we noted levels of attrition. We explored the impact of including studies with high levels of missing data in the overall assessment of treatment effect by using sensitivity analysis. For all outcomes, we carried out analyses, as far as possible, on an intention-to-treat basis, i.e. we attempted to include all participants randomised to each group in the analyses, and all participants were analysed in the group to which they were allocated, regardless of whether or not they received the allocated intervention. The denominator for each outcome in each trial was the number randomised minus any participants whose outcomes were known to be missing.

Assessment of heterogeneity
We applied tests of heterogeneity between trials to assess the significance of any differences between trials in the analyses. We regarded heterogeneity as substantial if the $I^2$ was greater than 30% and either the $\tau^2$ was greater than zero, or there was a low $P$ value (less than 0.10) in the Chi² test for heterogeneity.

Assessment of reporting biases
We investigated reporting biases (such as publication bias) using funnel plots. We assessed funnel plot asymmetry visually. If asymmetry was suggested by a visual assessment, we performed exploratory analyses to investigate it.
Data synthesis

We carried out statistical analysis using the Review Manager software (RevMan 2014). We used fixed-effect meta-analysis for combining data where it was reasonable to assume that studies were estimating the same underlying treatment effect: i.e. where trials were examining the same intervention, and the trials’ populations and methods were judged sufficiently similar. If there was clinical heterogeneity sufficient to expect that the underlying treatment effects differed between trials, or if substantial statistical heterogeneity was detected, we used random-effects meta-analysis to produce an overall summary, if an average treatment effect across trials was considered clinically meaningful. We treated the random-effects summary as the average of the range of possible treatment effects and we discuss the clinical implications of treatment effects differing between trials. If the average treatment effect was not clinically meaningful, we did not combine trials. Where we used random-effects analyses, we have presented the results as the average treatment effect with 95% confidence intervals, and the estimates of Tau^2 and I^2.

Subgroup analysis and investigation of heterogeneity

Had we identified substantial heterogeneity, we planned to investigate it using subgroup analyses and sensitivity analyses and to consider whether an overall summary was meaningful, and if it was, to use random-effects analysis to produce it.

In this update, we were not able to carry out the following subgroup analyses:

1. the dose of vitamin(s) (below or above the recommended dietary intake); and the duration of vitamin usage, based on time of trial entry: before pregnancy, < 12 weeks' gestation, between 12-20 weeks' gestation or 'mixed', which included women enrolled before and after 20 weeks' gestation;

2. their risk of spontaneous miscarriage (high risk defined as the presence of medical conditions associated with miscarriage such as hyperhomocysteinaemia, thrombophilia, antiphospholipid syndrome, systemic lupus erythematosus; low risk defined as none of the above conditions); their risk of recurrent miscarriage (high risk defined as two or more previous consecutive spontaneous miscarriages, and/or the presence of medical conditions associated with miscarriage such as hyperhomocysteinaemia, thrombophilia, antiphospholipid syndrome, systemic lupus erythematosus; low risk defined as none of the above conditions);

3. low or adequate dietary vitamin intake at trial entry (low intake defined as less than the recommended daily intake for each vitamin in that setting, as measured by dietary questionnaire).

We would have included all outcomes in the subgroup analysis. We planned to assess subgroup differences by interaction tests available within RevMan (RevMan 2014) and to report the results of subgroup analyses quoting the Chi^2 statistic and P value, and the interaction test I^2 value.

Sensitivity analysis

We carried out sensitivity analysis to explore the effects of trial quality and type of randomisation on the primary outcomes related to fetal loss (total fetal loss and early or late miscarriage). We included only trials with ‘adequate’ rating on allocation concealment. We considered these trials to be of high quality. We also carried out sensitivity analysis by excluding cluster-randomised trials and comparing the results of cluster-randomised trials with the individually-randomised trials.

RESULTS

Description of studies

See tables Characteristics of included studies and Characteristics of excluded studies for details of individual studies.

Included studies

For the 2016 update, we included a total of 40 trials (involving 276,820 women) assessing supplementation with specific vitamin(s) starting prior to 20 weeks' gestation. Many of the trials assessed interventions not specifically for the prevention of miscarriage, however, the authors included any supplementation trials, where supplementation began prior to 20 weeks' gestation, and where at least one miscarriage-related outcome as specified in the review was reported.

Participants

The demographic and obstetric characteristics of the women varied widely between the trials (see table Characteristics of included studies). The 40 included trials contributed data for analysis from 276,820 women. Eight of the 40 included studies were cluster-randomised trials including 217,726 women (Bhutta 2009; Katz 2000; Summit 2008; Sunawang 2009; West 2011; West 2014; Zagre 2007; Zeng 2008). Two of the trials from the previous version of this review (one cluster (Katz 2000), and one small trial (Roberfroid 2008)) included women who were pregnant more than once in the study period; resulting in data contributing to 59,146 pregnancies for the individual trials and 219,267 pregnancies from the cluster trials. Five trials enrolled women prior to conception (Czeizel 1994; Hemmi 2003; ICMR 2000; Kirke 1992; MRC 1991) and asked women to continue taking the supplements until the second or third missed menstrual period. One trial (Katz 2000) supplemented women from before conception, through pregnancy and postpartum for a total of 3.5 years postpartum. Another eight trials enrolled women in the first trimester (Briscoe 1959; Hans 2010; Rumiris 2006; Tofail 2008; West 2011; West 2014; Wibowo 2012; Zagre 2007), and 24 trials in early
to mid pregnancy (Bhutta 2009; Chappell 1999; Fawzi 1998; Fawzi 2007; Fleming 1968; Fleming 1986; Kumwenda 2002; McCance 2010; Osrin 2005; People’s League 1942; Poston 2006; Prawirohartono 2011; Roberfroid 2008; Roberts 2010; Rumbold 2006; Rush 1980; Schmidt 2001; Spinnato 2007; Steyn 2003; Sunawang 2009; Van den Broek 2006; Villar 2009; Xu 2010; Zeng 2008). Some of the trials enrolling women in early to mid pregnancy included women enrolled at or after 20 weeks’ gestation (Chappell 1999; Fawzi 1998; Fawzi 2007; Fleming 1968; Fleming 1986; Kumwenda 2002; McCance 2010; Osrin 2005; People’s League 1942; Roberfroid 2008; Rumbold 2006; Rush 1980; Schmidt 2001; Spinnato 2007; Steyn 2003; Van den Broek 2006; Villar 2009; Zeng 2008). One trial (Summit 2008), enrolled 41,839 women at ‘any gestational age’, although more than 70% of the women were enrolled in the first or second trimester. Two trials (Fawzi 1998; Kumwenda 2002) involved vitamin A supplementation in women seropositive for the Human Immunodeficiency Virus (HIV); one trial (Poston 2006) involved only women with clinical risk factors for pre-eclampsia, while one trial (McCance 2010), limited the eligibility to women with type 1 diabetes. Roberts 2010 involved only nulliparous women. Baseline characteristics of women enrolled in the intervention group and control group were comparable in all the trials except two (Xu 2010; Zagre 2007). In Xu 2010, there was a slightly higher proportion of women with multiple pregnancies in the placebo group; while in Zagre 2007, women in the control group tended to be poorer and less educated, while women in the intervention group had larger households and used more preventive measures against malaria.

The trials were conducted in both resource-rich and resource-poor countries including the United States (Briscoe 1959; Roberts 2010; Rush 1980), Canada (Xu 2010), the United Kingdom (Chappell 1999; McCance 2010; People’s League 1942; Poston 2006), Hungary (Czeizel 1994), Tanzania (Fawzi 1998; Fawzi 2007), Nigeria (Fleming 1968; Fleming 1986), Burkino Faso (Roberfroid 2008), Japan (Hemmi 2003), India (ICMR 2000), Nepal (Katz 2000; Osrin 2005), the Republic of Ireland (Kirke 1992), Uganda (Hans 2010), Bangladesh (West 2011; West 2014; Tofail 2008), China (Zeng 2008), Niger (Zagre 2007), Pakistan (Bhutta 2009), Australia (Rumbold 2006), Brazil (Spinnato 2007), Mexico (Xu 2010), Malawi (Kumwenda 2002; Van den Broek 2006), Indonesia (Prawirohartono 2011; Rumiris 2006; Schmidt 2001; Sunawang 2009; Summit 2008; Wibowo 2012), and South Africa (Steyn 2003). One trial involved 33 international centres (MRC 1991) and another trial was a multi-country study involving India, Peru, South Africa and Vietnam (Villar 2009).

**Interventions**

The 40 trials assessed a range of vitamin supplements, alone or in combination with other supplements. The vitamins included vitamin A, alone or with iron, folic acid, multivitamins, or β-carotene (Fawzi 1998; Katz 2000; Kumwenda 2002; Prawirohartono 2011; Schmidt 2001; Van den Broek 2006; West 2011); vitamin C with or without multivitamins or vitamin E (Briscoe 1959; Chappell 1999; Hans 2010; Hemmi 2003; McCance 2010; Poston 2006; Roberts 2010; Rumbold 2006; Spinnato 2007; Steyn 2003; Villar 2009; Xu 2010); folic acid with or without multivitamins and/or iron (Czeizel 1994; Fleming 1968; Fleming 1986; ICMR 2000; Kirke 1992; MRC 1991); antioxidant vitamins (Wibowo 2012); multivitamins with/without folic acid, vitamin A, vitamin E or iron and folic acid (Bhutta 2009; Czeizel 1994; Fawzi 1998; Fawzi 2007; ICMR 2000; Kirke 1992; MRC 1991; Osrin 2005; Roberfroid 2008; Rumiris 2006; Rush 1980; Sunawang 2009; Summit 2008; Tofail 2008; West 2014; Zagre 2007; Zeng 2008); and multivitamins alone (People’s League 1942; Rush 1980). The doses of vitamins were similar for the vitamin C supplementation trials (range 400 mg to 1000 mg). However, they varied widely between trials for the folic acid (range 0.3 mg to 10 mg), multivitamins and vitamin A trials (range 5000 international units (IU) to 23,300 IU). The components of MMN supplementation were different among the trials but all of them contained iron and folate in the MMN supplements. All supplements were taken orally from the enrolment until delivery or up to 3.5 years postpartum.

**Outcomes**

**Main outcomes**

Thirty-six trials reported pregnancy loss as miscarriage or stillbirth. The outcome ‘total fetal loss’ included both miscarriage or stillbirth, and overcame problems with different definitions of miscarriage and stillbirth. There was no consistency amongst trials with regards to the definition of miscarriage. For some trials, miscarriage was considered to occur up until 26 or 28 weeks’ gestation, while other studies reported miscarriage as pregnancy loss prior to 20 weeks’ gestation. Other studies did not specify their definition of miscarriage or stillbirth.

**Other outcomes**

There was no consistency amongst trials with regards to the definition of stillbirth. In some trials, stillbirth was considered as pregnancy loss greater than or equal to 20 weeks’ gestation, while some trials considered stillbirth as pregnancy loss beyond 24 weeks’ gestation. Thirty trials reported stillbirth as an outcome. Only one trial (Spinnato 2007) reported on adverse effects of vitamin sufficient to stop supplementation, while congenital malformations was reported in nine trials (Czeizel 1994; Kirke 1992; McCance 2010; MRC 1991; Osrin 2005; Poston 2006; Spinnato 2007; Villar 2009; Xu 2010).

**Excluded studies**

We excluded 48 trials, of which 16 trials reported no clinically relevant data, or data in a format suitable for inclusion (Christian
Seven trials did not clearly report the gestational age when supplementation was started (Biswas 1984; Fletcher 1971; Hampel 1974; Lumeng 1976; Schuster 1984; Trigg 1976; Young 2015), and for two trials, the majority of women were enrolled after 20 weeks and did not report outcomes separately for women starting supplementation prior to 20 weeks (Ferguson 1955; Giles 1971). Thirteen trials (Baumslag 1970; Blot 1981; Chanarin 1968; Colman 1974; Courtsoudis 1999; Dawson 1962; Edelstein 1968; Feyi-Waboso 2005; Hankin 1966; Kaestel 2005; Marya 1981; Metz 1965; Owen 1966) reported supplementation after 20 weeks' gestation. One trial (Ross 1985) did not specify the contents of the supplements; in five trials all women were given a vitamin supplement (Hekmatdoost 2011; Hunt 1984; Huybregts 2009; Shu 2002; Wehby 2012); one trial was a food intervention (Potdar 2014) and two were non-randomised (Smithells 1981; Ulrich 1999).

Three other trials (Beazley 2005; Chaudhuri 1969; Rivas-Echeverria 2000) supplemented women for the prevention of pre-eclampsia, and did not report any outcomes related to pregnancy loss. These trials are covered in the Cochrane review 'Antioxidants for preventing pre-eclampsia' (Rumbold 2008).

Risk of bias in included studies

Figure 1 and Figure 2 illustrate that the trials were of variable quality. Two trials (Fleming 1968; People’s League 1942) used quasi-random allocation methods involving alternate allocation of participants. Similarly, eight trials (Bhutta 2009; Katz 2000; Summit 2008; Sunawang 2009; West 2011; West 2014; Zagre 2007; Zeng 2008) used cluster randomisation.
Figure 2. 'Risk of bias' summary: review authors’ judgements about each risk of bias item for each included study

<table>
<thead>
<tr>
<th>Study</th>
<th>Random sequence generation</th>
<th>Allocation concealment</th>
<th>Blinding of participants and personnel</th>
<th>Blinding of outcome assessment</th>
<th>Incomplete outcome data</th>
<th>Selective reporting</th>
<th>Other bias</th>
</tr>
</thead>
<tbody>
<tr>
<td>Bissell 2009</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Bisson 1999</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Czepiel 1994</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Fand 1990</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Fand 1997</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Fleming 1998</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Fleming 1996</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Hens 1990</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Hens 2003</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>ICMF 2009</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Isavvina 2004</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Hiz 2000</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Kilege 1992</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Kanwana 2002</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Khi 2010</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Kho 1990</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>WCR 1991</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Osen 2005</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>People’s League 1942</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Peth 2006</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Peczynkiewicz 2011</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Pinkney 2000</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Roberts 2010</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Rumins 2004</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Rush 1990</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Sinden 2009</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Sinden 2007</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Olche 2003</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Dunn 2008</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Dunwoody 2009</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Taha 2008</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Van de Velde 2006</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Van 2009</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Woot 2014</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Woot 2014</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Woot 2012</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Ju 2010</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Zap 2007</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Zong 2008</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>
Allocation

Sequence generation: 25 trials adequately randomised the participants to the intervention and control groups and were judged to be of low risk of bias (Bhutta 2009; Chappell 1999; Hans 2010; Kirke 1992; Kumwenda 2002; McCance 2010; MRC 1991; Osrin 2005; Poston 2006; Prawirohardono 2011; Roberfroid 2008; Roberts 2010; Rumbold 2006; Rumiris 2006; Spinnato 2007; Steyn 2003; Summit 2008; Tofail 2008; Van den Broek 2006; Villar 2009; West 2011; West 2014; Wibowo 2012; Xu 2010; Zeng 2008). The method used for random sequence generation was unclear in 13 trials (Briscoe 1959; Czeizel 1994; Fawzi 1998; Fawzi 2007; Fleming 1986; Hemmi 2003; ICMR 2000; Jauniaux 2004; Katz 2000; Rush 1980; Schmidt 2001; Sunawang 2009; Zagre 2007), because methodological details were not reported or not clearly described. The remaining two trials Fleming 1968 and People’s League 1942 used quasi-randomised methods and were rated as high risk of bias for sequence generation.

Allocation concealment: 20 trials were assessed to have a low risk of bias for adequate concealment of participants to treatment and control groups (Bhutta 2009; Chappell 1999; Fawzi 2007; Kirke 1992; Kumwenda 2002; McCance 2010; MRC 1991; Roberfroid 2008; Rumbold 2006; Rumiris 2006; Spinnato 2007; Steyn 2003; Summit 2008; Tofail 2008; Van den Broek 2006; Villar 2009; West 2011; West 2014; Zagre 2007; Zeng 2008). In 18 trials the method for allocation concealment was not described or not clearly described (Briscoe 1959; Czeizel 1994; Fawzi 1998; Fleming 1986; Hans 2010; Hemmi 2003; ICMR 2000; Jauniaux 2004; Katz 2000; Osrin 2005; Poston 2006; Prawirohardono 2011; Roberts 2010; Rush 1980; Schmidt 2001; Sunawang 2009; Wibowo 2012; Xu 2010). In two trials, allocation was not concealed and therefore judged as high risk of bias (Fleming 1968; People’s League 1942).

Blinding

Participants and personnel: 28 trials were assessed as having a low risk of performance bias and reported blinding of participants and personnel to the treatment allocation (Bhutta 2009; Briscoe 1959; Chappell 1999; Fawzi 1998; Fawzi 2007; Fleming 1968; Fleming 1986; Katz 2000; Kirke 1992; McCance 2010; MRC 1991; Poston 2006; Prawirohardono 2011; Roberfroid 2008; Roberts 2010; Rumbold 2006; Rumiris 2006; Spinnato 2007; Summit 2008; Tofail 2008; Van den Broek 2006; Villar 2009; West 2011; West 2014; Wibowo 2012; Xu 2010; Zagre 2007; Zeng 2008). Another 11 trials were judged to have an unclear risk of bias because no or not enough information were provided (Czeizel 1994; Hans 2010; Hemmi 2003; ICMR 2000; Jauniaux 2004; Kumwenda 2002; Osrin 2005; People’s League 1942; Rush 1980; Schmidt 2001; Steyn 2003). Sunawang 2009 was rated as high risk of bias as personnel (but not participants) were aware of participants’ allocation due to the different appearance of the supplements.

Outcome assessment: blinding of outcome assessors was clearly stated in 14 trials (Bhutta 2009; Chappell 1999; Fawzi 2007; Fleming 1968; Fleming 1986; McCance 2010; Prawirohardono 2011; Roberts 2010; Summit 2008; Tofail 2008; Villar 2009; West 2011; West 2014; Xu 2010) and unclear in 23 trials (Briscoe 1959; Czeizel 1994; Fawzi 1998; Hans 2010; Hemmi 2003; ICMR 2000; Jauniaux 2004; Katz 2000; Kumwenda 2002; MRC 1991; People’s League 1942; Poston 2006; Roberfroid 2008; Rumbold 2006; Rumiris 2006; Rush 1980; Schmidt 2001; Spinnato 2007; Steyn 2003; Van den Broek 2006; Wibowo 2012; Zagre 2007; Zeng 2008). In Kirke 1992; Osrin 2005 and Sunawang 2009 outcome assessors were not blinded to the allocation code and the trials were rated as high risk of detection bias.

Incomplete outcome data

Loss to follow-up ranged from no loss at all in Briscoe 1959; Rumbold 2006; Rumiris 2006; Steyn 2003 to over 20% in Fleming 1968; ICMR 2000; Summit 2008 and Van den Broek 2006. Incomplete outcome data was judged low risk of bias in 21 trials and high in nine trials. Ten trials were rated as unclear risk of bias due to missing information about loss of follow-up.

Selective reporting

Eighteen trials were considered to have a low risk of reporting bias. Another 20 trials were assessed as unclear risk of bias because of unavailability of trial registration or protocol (Bhutta 2009; Rush 1980), variations between the protocol and the publication (Poston 2006), due to insufficient details provided about methods or selective reporting (Briscoe 1959; Fleming 1968; Hans 2010; Hemmi 2003; ICMR 2000; Kirke 1992; MRC 1991; People’s League 1942; Roberfroid 2008; Steyn 2003; Villar 2009; Zagre 2007), variations of information between serial publications (Czeizel 1994; Fawzi 1998; Katz 2000; Schmidt 2001), or the trial was stopped before completion (Jauniaux 2004). The remaining two trials were at high risk of reporting bias; in Fleming 1986, not all outcomes were reported for all treatment groups and in McCance 2010, fewer outcomes were stated in the trial registration compared with the report.

Other potential sources of bias

In 18 trials, other sources of bias were not detected and these trials were rated as low risk of bias. Fourteen trials provided only limited methodological details to exclude other sources of bias and were judged as unclear (Briscoe 1959; Czeizel 1994; Fawzi 1998; Fleming 1968; Fleming 1986; ICMR 2000; Jauniaux 2004;
Effects of interventions

See: Summary of findings for the main comparison Vitamin C and vitamin E versus placebo for preventing miscarriage; Summary of findings 2 Vitamin A plus iron plus folate versus iron plus folate for preventing miscarriage; Summary of findings 3 Multivitamin plus iron plus folate versus iron plus folate for preventing miscarriage

See Summary of findings for the main comparison; Summary of findings 2; Summary of findings 3 for each of the main comparisons. The quality of evidence is included for numerical results of outcomes included in the 'Summary of findings' tables. We have included 40 trials, involving 276,820 women and 278,413 pregnancies. One trial (Jauniaux 2004) contributed no outcome data because it was stopped early and withdrawn. Thus, 39 trials contributed data to our analyses.

Vitamin C supplementation

The trials involving vitamin C supplementation included the following interventions: vitamin C plus multivitamins versus placebo plus multivitamins (Briscoe 1959; Hans 2010), vitamin C and vitamin E supplementation versus placebo (Chappell 1999; McCance 2010; Poston 2006; Roberts 2010; Rumbold 2006; Spinnato 2007; Xu 2010), and vitamin C alone versus no supplement or placebo (Hemmi 2003; Steyn 2003).

Primary outcomes

There was no difference in the risk of total fetal loss between women receiving:

1. vitamin C with vitamin E (risk ratio (RR) 1.14, 95% confidence interval (CI) 0.92 to 1.40, seven trials, 18,949 women; Analysis 1.1; high-quality evidence);
2. vitamin C alone (RR 1.28, 95% CI 0.58 to 2.83, two trials, 224 women; Analysis 2.1);
3. vitamin C with multivitamins (RR 1.32, 95% CI 0.63 to 2.77, one trial, 406 women; Analysis 3.1);
4. vitamin C with or without multivitamins versus multivitamins (excluding vitamin A) or placebo (Fawzi 1998), and vitamin A plus iron and folic acid versus iron and folic acid (Kumwenda 2002; Schmidt 2001; Van den Broek 2006).
5. vitamin A with or without multivitamin (RR 0.80, 95% CI 0.53 to 1.21, one trial, 1074 women; Analysis 8.1); compared with placebo or no vitamin A groups. The heterogeneity in Analysis 5.1 seemed to have been contributed by combining two cluster-randomised trials and one individually-randomised trial. Heterogeneity was no longer present when the individually-randomised trial was excluded. However, this did not change the conclusion of no significant difference between vitamin A and no vitamin A groups. There was no evidence of a difference in the risk for early or late miscarriage between women receiving:

1. vitamin A plus iron and folate (RR 0.86, 95% CI 0.46 to 1.62, two trials, 1397 women; Analysis 4.2; low-quality evidence);
2. vitamin A alone (RR 0.98, 95% CI 0.92 to 1.04, one trial, 39,668 women; Analysis 5.2);
3. beta-carotene alone (RR 1.00, 95% CI 0.94 to 1.06, one trial, 39,860 women; Analysis 6.2);
4. vitamin A with or without multivitamin (RR 0.76, 95% CI 0.37 to 1.55, one trial, 1075 women; Analysis 8.2); compared with placebo or no vitamin A groups.

Secondary outcomes
There was no difference in the risk for stillbirth for the following supplementation treatments:

1. vitamin A plus iron and folate (RR 1.29, 95% CI 0.57 to 2.91, three trials, 1640 women; Analysis 4.3; low-quality evidence);
2. vitamin A alone (RR 0.95, 95% CI 0.86 to 1.06, one trial, 39,668 women; Analysis 5.3);
3. beta-carotene alone (RR 1.09, 95% CI 0.98 to 1.20, one trial, 39,860 women; Analysis 6.3);
4. vitamin A with or without multivitamin (RR 1.04, 95% CI 0.60 to 1.79, one trial, 1075 women; Analysis 8.3); compared with placebo or no vitamin A groups. Congenital malformations and adverse effects of vitamin supplementation were not reported by trials included in these analyses.

Multivitamin supplementation
The trials involving multivitamin supplementation included the following interventions: multivitamins with or without folic acid versus no multivitamins or folic acid (Czeizel 1994; ICMR 2000; MRC 1991); multivitamins with or without folic acid versus folic acid (Kirke 1992; MRC 1991; Zeng 2008); multivitamins with or without vitamin A versus vitamin A or placebo (Fawzi 1998); multivitamins versus control (People’s League 1942); multivitamins with vitamin E versus multivitamins without vitamin E or control (Rush 1980); multivitamins with iron and folic acid versus iron and folic acid (Bhutta 2009; Fawzi 2007; Osirin 2005; Roberfroid 2008; Rumiris 2006; Sunawang 2009; Summit 2008; Tofail 2008; West 2014; Zagre 2007).

Primary outcomes
The risk for total fetal loss was reduced in women supplemented with:

1. multivitamin without folic acid (RR 0.49, 95% CI 0.34 to 0.70, one trial, 907 women; Analysis 10.1);
2. multivitamin with/without vitamin A (RR 0.60, 95% CI 0.39 to 0.92, one trial, 1074 women; Analysis 15.1); compared with placebo or no multivitamins groups. There was no difference in the risk of total fetal loss for the following interventions:

1. multivitamin plus iron and folic acid compared iron and folate only groups (RR 0.96, 95% CI 0.93 to 1.00, 10 trials, 94,948 women; Analysis 9.1; high-quality evidence);
2. multivitamins alone or in combination with other vitamins or micronutrients compared with placebo or no multivitamins groups (Analysis 11.1; Analysis 12.1; Analysis 13.1; Analysis 14.1; Analysis 16.1; Analysis 17.1; Analysis 18.1 (random effects; three trials)). Similarly, we found no difference in the risk for early or late miscarriage between women receiving the following interventions:

1. multivitamin plus iron and folic acid compared iron and folate only groups (RR 0.98, 95% CI 0.94 to 1.03, 10 trials, 94,948 women; Analysis 9.2; moderate-quality evidence);
2. multivitamins alone or in combination with other vitamins or micronutrients compared with placebo or no multivitamins groups (Analysis 10.2; Analysis 11.2; Analysis 12.2; Analysis 13.2; Analysis 14.2; Analysis 17.2; Analysis 18.2 (random effects; three trials)). The heterogeneity in Analysis 18.1 and Analysis 18.2 seemed to have been contributed by ICMR 2000, which included women who had previously given birth to a child with an open neural tube defect. When this trial was excluded, the heterogeneity was no longer present.

Secondary outcomes
There was evidence of a decrease in the risk for stillbirth among women receiving multivitamin plus iron and folic acid compared iron and folate only groups (RR 0.92, 95% CI 0.85 to 0.99, 10 trials, 79,851 women; Analysis 9.3; high-quality evidence). There was no difference in the risk of:

1. stillbirth (Analysis 10.3; Analysis 11.3; Analysis 12.3; Analysis 13.3; Analysis 14.3; Analysis 16.2; Analysis 17.3; Analysis 18.3); or
2. congenital malformation (Analysis 10.4; Analysis 11.4; Analysis 12.4; Analysis 13.4; Analysis 14.4; Analysis 18.4) between women receiving multivitamins alone or in combination with other vitamins or micronutrients compared with placebo or no multivitamins groups. There were no data available to conduct any analysis for adverse effects of vitamin supplementation.
Folic acid supplementation

The trials involving folic acid supplementation included the following interventions: folic acid with or without multivitamins compared with no folic acid or multivitamins (Czeiél 1994; ICMR 2000; MRC 1991); folic acid with or without multivitamins compared with multivitamins (Kirke 1992; MRC 1991); folic acid and iron compared with iron (Fleming 1968); folic acid and iron compared with no iron or folic acid (Fleming 1986).

Primary outcomes

We found no difference in the risk of:

1. total fetal loss (Analysis 19.1 (random effects; three trials); Analysis 20.1; Analysis 21.1; Analysis 22.1; Analysis 23.1; Analysis 24.1; Analysis 25.1; Analysis 26.1); or
2. early or late miscarriage (Analysis 19.2 (random effects; three trials); Analysis 20.2; Analysis 21.2; Analysis 22.2; Analysis 23.2; Analysis 24.2; Analysis 25.2; Analysis 26.2); between women supplemented with folic acid with or without multivitamins and/or iron compared with no folic acid groups. The heterogeneity found seemed to have been contributed by ICMR 2000, which included women who had previously given birth to a child with an open neural tube defect. Excluding this trial removed the heterogeneity but did not change the conclusion of no difference between the treatment groups.

Secondary outcomes

There was no difference in the risk of:

1. stillbirth (Analysis 19.3; Analysis 20.3; Analysis 21.3; Analysis 22.3 (random effects); Analysis 23.3; Analysis 24.3 (random effects); Analysis 25.3); or
2. congenital malformations (Analysis 19.4; Analysis 20.4; Analysis 21.4; Analysis 22.4 (random effects); Analysis 23.4; Analysis 24.4 (random effects)) between women supplemented with folic acid with or without multivitamins and/or iron; compared no folic acid groups. There were no data available to conduct any analysis for adverse effects of vitamin supplementation.

Antioxidant vitamins supplementation

The trial involving antioxidant vitamins supplementation included the following interventions: antioxidant with multivitamins compared with multivitamins with low antioxidant content (Wibowo 2012).

Primary outcomes

In the one trial involving 110 women (Wibowo 2012), there was no evidence of differences between women given antioxidant with multivitamins compared multivitamins with low antioxidant group on early or late miscarriage (RR 1.12, 95% CI 0.24 to 5.29, one trial, 110 women, Analysis 27.1). No other primary or secondary outcomes were reported by this trial.

Subgroup analyses by dose of vitamins and duration of vitamin usage

Subgroup analyses by dose of vitamin(s) (below or above the recommended dietary intake) were complicated by the limited number of studies in each vitamin group, and by the use of multivitamin supplements. For many of the vitamin types and for those reporting pregnancy loss outcomes, all of the trials supplemented women with amounts that were above the recommended dietary intake. Similarly, the duration of vitamin usage was complicated by the fact that many of the trials had wide recruitment periods, and one trial (Katz 2000) supplemented women up until three years postpartum. We have not performed subgroup analyses based on vitamin dosage or time of trial entry.

Subgroup analyses by women’s risk of spontaneous or recurrent miscarriage

Information enabling women to be classified at high or low risk of either spontaneous miscarriage or recurrent miscarriage was not clearly stated in any of the trials included in this update. Based on the inclusion criteria, one trial (Rumbold 2006) included women at low risk of miscarriage. One trial (Briscoe 1959) included women who had experienced recurrent miscarriage as well as women at high risk of miscarriage (more than two previous miscarriages and/or bleeding in the pregnancy) and low-risk women (two or less previous miscarriages and no bleeding in the pregnancy). After classifying women into these groups, the number of women in the high-risk group was too small to permit any meaningful comparisons and we have therefore not performed subgroup analyses.

Subgroup analyses by dietary intake of vitamins

Seven trials (Bhutta 2009; Fleming 1968; Kumwenda 2002; People’s League 1942; Schmidt 2001; Steyn 2003; West 2011) reported information about women’s nutritional status or the percentage of women who were dietary deficient at trial entry for the vitamin of interest. Other trials reported that they were being undertaken in countries where the population was at high risk of multiple micronutrient deficiencies (Osirin 2005; Prawiroharto 2011; Roberfroid 2008; Summit 2008; Villar 2009), or there was a high prevalence of anaemia (Bhutta 2009; Fleming 1986; Sunawang 2009; Zagre 2007; Zeng 2008), but provided no specific information on nutritional status of participants. Two trials (Rumiris 2006; Wibowo 2012) included women with ‘low antioxidant status’. There were not enough trials within each vitamin group to assess the role of supplementation in women with dietary deficient intakes of the individual vitamins and results were not reported separately for women with a low dietary vitamin intake; therefore, we could not perform subgroup analyses.
Sensitivity analyses

We carried out sensitivity analysis to explore the effects of trial quality and type of randomisation on the primary outcomes related to fetal loss (total fetal loss and early or late miscarriage). We included only trials with 'adequate' rating on allocation concealment, but found that restricting to only trials with adequate allocation concealment made very little difference to the results for the primary outcomes. Effect of type of randomisation was explored by excluding cluster-randomised trials and restricting the analyses to individually-randomised trials. We found no difference between women supplemented with multivitamins compared with controls for total fetal loss or early or late miscarriage when the analyses were restricted to individually-randomised trials only. These sensitivity analyses indicate that the analyses for the effects of multivitamins on outcomes related to fetal loss and early or late miscarriage are no different when only individually-randomised trials are included.
## ADDITIONAL SUMMARY OF FINDINGS

### Vitamin A plus iron plus folate versus iron plus folate for preventing miscarriage

**Population:** pregnant women  
**Settings:** Indonesia, Malawi  
**Intervention:** vitamin A plus iron plus folate  
**Comparison:** iron plus folate

<table>
<thead>
<tr>
<th>Outcomes</th>
<th>Anticipated absolute effects* (95% CI)</th>
<th>Relative effect (95% CI)</th>
<th>N of participants (studies)</th>
<th>Quality of the evidence (GRADE)</th>
<th>Comments</th>
</tr>
</thead>
<tbody>
<tr>
<td>Total fetal loss (including miscarriages or combined miscarriages and stillbirths) - Vitamin A + iron + folate versus iron + folate</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Study population</td>
<td></td>
<td>RR 1.01 (0.61 to 1.66)</td>
<td>1640 (3 RCTs)</td>
<td>⊕⊕⃝⃝ LOW</td>
<td>12</td>
</tr>
<tr>
<td>37 per 1000</td>
<td>37 per 1000 (22 to 61)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Moderate</td>
<td>59 per 1,000</td>
<td>60 per 1,000 (36 to 98)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Early or late miscarriage - Vitamin A + iron + folate versus iron + folate</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Study population</td>
<td></td>
<td>RR 0.86 (0.46 to 1.62)</td>
<td>1397 (2 RCTs)</td>
<td>⊕⊕⃝⃝ LOW</td>
<td>12</td>
</tr>
<tr>
<td>31 per 1000</td>
<td>26 per 1000 (14 to 50)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Moderate</td>
<td>50 per 1,000</td>
<td>43 per 1,000 (23 to 80)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Stillbirth - Vitamin A + iron + folate versus iron + folate</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Study population</td>
<td></td>
<td>RR 1.29 (0.57 to 2.91)</td>
<td>1640 (3 RCTs)</td>
<td>⊕⊕⃝⃝ LOW</td>
<td>12</td>
</tr>
<tr>
<td></td>
<td>13 per 1000</td>
<td>16 per 1000 (7 to 37)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>----------------------</td>
<td>-------------</td>
<td>-----------------------</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>Moderate</td>
<td>21 per 1,000</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td></td>
<td>27 per 1,000 (12 to 61)</td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

Any adverse effects | See comments. | No studies reported this outcome.

* The risk in the intervention group (and its 95% confidence interval) is based on the assumed risk in the comparison group and the relative effect of the intervention (and its 95% CI).

CI: Confidence interval; RR: Risk ratio

**GRADE Working Group grades of evidence**

High quality: We are very confident that the true effect lies close to that of the estimate of the effect

Moderate quality: We are moderately confident in the effect estimate: The true effect is likely to be close to the estimate of the effect, but there is a possibility that it is substantially different

Low quality: Our confidence in the effect estimate is limited: The true effect may be substantially different from the estimate of the effect

Very low quality: We have very little confidence in the effect estimate: The true effect is likely to be substantially different from the estimate of effect

1. High and unclear risk of attrition bias.
2. Wide 95% CI.
## Multivitamin plus iron plus folate versus iron plus folate for preventing miscarriage

**Population:** pregnant women  
**Settings:** Bangladesh, Barkino Faso, Indonesia, Nepal, Niger, Pakistan, Tanzania  
**Intervention:** vitamin A plus iron plus folate  
**Comparison:** iron plus folate

<table>
<thead>
<tr>
<th>Outcomes</th>
<th>Anticipated absolute effects* (95% CI)</th>
<th>Relative effect (95% CI)</th>
<th>No. of participants (studies)</th>
<th>Quality of the evidence (GRADE)</th>
<th>Comments</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Total fetal loss (including miscarriages or combined miscarriages and stillbirths) - Multivitamins + iron + folic acid versus iron + folic acid</strong></td>
<td>Study population</td>
<td>RR 0.96 (0.93 to 1.00)</td>
<td>94,948 (10 RCTs)</td>
<td>⊕⊕⊕⊕ HIGH</td>
<td></td>
</tr>
<tr>
<td>Risk with placebo</td>
<td>136 per 1000</td>
<td>130 per 1000 (126 to 136)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Risk with Multivitamin</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Moderate</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Risk with placebo</td>
<td>218 per 1,000</td>
<td>209 per 1,000 (202 to 218)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Risk with Multivitamin</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td><strong>Early or late miscarriage - Multivitamin + iron + folic acid versus iron + folic acid</strong></td>
<td>Study population</td>
<td>RR 0.98 (0.94 to 1.03)</td>
<td>94948 (10 RCTs)</td>
<td>⊕⊕⊕ MODERATE</td>
<td></td>
</tr>
<tr>
<td>Risk with placebo</td>
<td>84 per 1000</td>
<td>83 per 1000 (79 to 87)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Risk with Multivitamin</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Moderate</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Risk with placebo</td>
<td>134 per 1,000</td>
<td>132 per 1,000 (126 to 138)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Risk with Multivitamin</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td><strong>Stillbirth - Multivitamin + iron + folic acid versus iron + folic acid</strong></td>
<td>Study population</td>
<td>RR 0.92 (0.85 to 0.99)</td>
<td>79,851 (10 RCTs)</td>
<td>⊕⊕⊕⊕ HIGH</td>
<td></td>
</tr>
</tbody>
</table>

*Anticipated absolute effects are calculated as the number of events per 1000 or 1,000 participants, depending on the study population size.
<p>| | | |</p>
<table>
<thead>
<tr>
<th></th>
<th></th>
<th></th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>29 per 1000</td>
<td>26 per 1000</td>
</tr>
<tr>
<td></td>
<td>(24 to 28)</td>
<td>(24 to 28)</td>
</tr>
<tr>
<td>Moderate</td>
<td>46 per 1,000</td>
<td>43 per 1,000</td>
</tr>
<tr>
<td></td>
<td>(39 to 46)</td>
<td>(39 to 46)</td>
</tr>
</tbody>
</table>

Any adverse effects | See comments | No studies reported this outcome

*The risk in the intervention group (and its 95% confidence interval) is based on the assumed risk in the comparison group and the relative effect of the intervention (and its 95% CI).

CI: Confidence interval; RR: Risk ratio

**GRADE Working Group grades of evidence**

**High quality:** We are very confident that the true effect lies close to that of the estimate of the effect

**Moderate quality:** We are moderately confident in the effect estimate: The true effect is likely to be close to the estimate of the effect, but there is a possibility that it is substantially different

**Low quality:** Our confidence in the effect estimate is limited: The true effect may be substantially different from the estimate of the effect

**Very low quality:** We have very little confidence in the effect estimate: The true effect is likely to be substantially different from the estimate of effect

---

1 Publication bias detected by funnel plot.

2 Wide confidence interval crossing the line of no effect.
**DISCUSSION**

**Summary of main results**

The purpose of this review was to determine the effectiveness and safety of any vitamin supplementation taken by women pre- or periconceptionally on the risk of miscarriage. In this updated version of the review, we included 40 studies involving 59,094 women from individually-randomised trials plus a further 217,726 women from eight cluster-randomised controlled trials. The results did not provide sufficient evidence to support the use of single vitamin supplementation for preventing total fetal loss or early or late miscarriage. However, stillbirth was significantly lower in women given multivitamin supplementation plus iron and folic acid compared to iron and folic acid alone. Although there was evidence of decreased risk for total fetal loss among women receiving multivitamins without folic acid compared with no multivitamin/folic acid and multivitamin supplementation with/without vitamin A compared with vitamin A or placebo; these findings occurred in analyses involving one trial each with small numbers of women involved. Also, they include studies where the comparison groups included women receiving either vitamin A or placebo, and thus require caution in interpretation.

**Overall completeness and applicability of evidence**

There was considerable consistency in reported total fetal loss (including miscarriages or combined miscarriages and stillbirths) among included studies with no difference in the rates of miscarriage and stillbirth across treatment groups. While this may suggest the true effect of vitamin supplementation on risk of miscarriage, most of the studies included in this review did not originally set out to examine the effect of vitamin supplementation on the risk of miscarriage. Our review included trials that randomised women prior to conception; however, in some cases, not all women enrolled in these trials fell pregnant during the study period. Some of the trials reported outcomes only for women falling pregnant, whereas other trials did not distinguish between women who were never pregnant and women who may have been pregnant but were lost to follow-up. The outcomes in this review relating to pregnancy outcomes are not relevant for the women who never became pregnant during the study period. In this review, where trials provided accurate information about the number of women who joined the study and became pregnant in the time period, we included this number in the totals, rather than the number of women who may have been randomised. Where it was not clear about the exact number of women with a confirmed pregnancy, we included all women who had been randomised. This may therefore mean that a certain proportion of women in the denominator were never pregnant during the study period. By including these women who were never pregnant in the totals, the review assumes that if these women had become pregnant, they would not have had a miscarriage, which is unlikely to be entirely correct. Including these women creates the potential to underestimate any treatment effects observed.

Similarly, for one large trial (Katz 2000) and one smaller trial (Roberfroid 2008), some women were pregnant more than once during the study period. In these trials, the denominators reported are the total number of pregnancies during the study period, not the total number of women randomised, which incorrectly assumes that each data point included is independent from the next. This has the potential to either underestimate or overestimate the results, depending on whether the women contributing data for more than one pregnancy may be more or less susceptible to experiencing miscarriage or stillbirth. One way to overcome this may be to summarise the data for each woman so that there is only one set of data points for each woman; however, we were unable to do this for these particular studies.

**Quality of the evidence**

Some of the trials included in the review were at high risk of bias, either due to poor or unclear allocation concealment or large losses to follow-up. The data were also complicated by differing definitions of miscarriage. For some trials, miscarriage was considered to occur up until 26 or 28 weeks’ gestation, while other studies reported miscarriage as pregnancy loss prior to 20 weeks’ gestation, and stillbirth as pregnancy loss greater or equal to 20 weeks’ gestation. Other studies did not specify their definition of miscarriage or stillbirth. In addition to the problems with differing definitions, the timing of the onset of vitamin supplementation for some of the included trials occurred in mid-pregnancy, which may limit the impact of supplementation on the risk of miscarriage. The review attempted to overcome these issues by using the outcome ‘total fetal loss’, which included either miscarriage or stillbirth. We assessed the quality of the evidence using GRADE and judged the evidence for vitamin C and vitamin E compared with control as high quality for total fetal loss, and moderate quality for early or late miscarriage, stillbirth, and adverse effects, which was downgraded due to wide 95% confidence intervals (CIs) (Summary of findings for the main comparison). For vitamin A plus iron plus folic acid versus iron plus folic acid trials were judged to have low quality of evidence for total fetal loss, early or late miscarriage, and stillbirth due to design limitation and wide 95% CI (Summary of findings 2). No studies reported any adverse effects for this comparison. For multivitamin plus iron plus folic acid versus iron plus folic acid trials were judged to be high quality for total fetal death and stillbirth, moderate quality for early or late miscarriage, downgraded due to publication bias suspected by funnel plot, or wide CI crossing the line of no effect (Summary of findings 3). No studies reported any adverse effects for this comparison.

In order to determine the effect of publication bias, we undertook funnel plots for comparisons with 10 or more studies (Figure 3;
Figure 4: Figure 5) for the comparisons of multivitamins plus iron and folic acid versus iron and folic acid. Asymmetry was suggested by visual assessment of Figure 4 for early or late miscarriage.

Figure 3. Funnel plot of comparison: Multivitamin plus iron and folic acid versus iron and folic acid, outcome: Total fetal loss.
Figure 4. Funnel plot of comparison: Multivitamin plus iron and folic acid versus iron and folic acid, outcome: Early or late miscarriage.
Potential biases in the review process

We took steps to minimise the introduction of bias during the review process. All relevant trials were identified including published abstracts from conference proceedings, English and non-English publications. A pro forma translation sheet was used to extract relevant information from non-English articles. At least two review authors independently assessed each trial, performed data extraction, and assessment of risk of bias for each of the included trials. Our assessment of previously identified ongoing trials that remained unpublished were limited to trial published protocols or the records of the initial communication between our authors and the authors of the unpublished trials.

Agreements and disagreements with other studies or reviews

There are several Cochrane reviews evaluating the effect of single vitamin supplementation during pregnancy on maternal, fetal, neonatal and infant outcomes. Benefits or hazards of vitamin supplementation in pregnancy on total fetal loss and miscarriage have not been or insufficiently investigated. However, our results on secondary outcomes are consistent with finding in the particular publications.

In the analysis by vitamin type, vitamin C supplementation alone or in combination with vitamin E or multivitamins did not show any effect on total fetal loss, miscarriage, or the secondary outcomes stillbirth, congenital malformation and adverse effects. A review focusing on vitamin C supplementation alone or in combination with other separate supplements on pregnancy outcomes, did not observe effects on stillbirth or congenital malformations which is consistent with our results (Rumbold 2015).

Supplementing women with vitamin A alone or in combination with iron and folic acid or multivitamins was not associated with changes in fetal loss or miscarriage as well as stillbirth. These findings are consistent with the Cochrane review ‘Vitamin A supplementation during pregnancy for maternal and newborn outcomes’ (McCauley 2015), which found no difference in the rate of stillbirth for women receiving vitamin A alone compared with placebo/no treatment or vitamin A with other micronutrients compared with micronutrient supplementation without vitamin A.

In the analysis comparing multivitamin alone or in combination with other vitamins, we found a positive effect of multivitamin supplementation without folic acid compared with no multivi-
Vitamin/folic acid as well as multivitamin with/without vitamin A compared with vitamin A alone or placebo on total fetal loss. However, these findings resulted from only one study, respectively. Stillbirth was significantly reduced for women receiving multivitamin plus iron and folic acid. This result is consistent with findings in a review assessing the effect of multiple-micronutrient supplementation during pregnancy on maternal, fetal and infant health outcomes (Haider 2015). Here they also reported a significant reduction in the risk of stillbirth. Miscarriage (loss before 28 weeks) was not affected by this intervention.

Folic acid supplementation with or without multivitamin compared to no folic acid/multivitamin or multivitamin alone did not reduce the risk of total fetal loss, miscarriage, stillbirth or congenital malformations. This in accordance with a review evaluating the effectiveness of oral folic acid supplementation during pregnancy on maternal health and pregnancy outcomes (Lassi 2013). The authors did not observe any effect of folic acid supplementation on stillbirth. Even though miscarriage was included as a secondary outcome, none of the included studies reported on miscarriage. In addition, another review assessed the effects and safety of periconceptional oral folate supplementation for preventing birth defects (De-Regil 2015). There was no effect of folate versus no intervention, placebo or other micronutrients without folate on miscarriage or stillbirth. They investigated the effect of folate supplementation on several congenital malformations and found a 69% reduction in the risk of neural tube defects.

Antioxidant vitamin supplementation had no effect on early or late miscarriage. The effectiveness and safety of any antioxidant supplementation during pregnancy on the risk of various pregnancy outcomes is explored in the Cochrane review ‘Antioxidants for preventing pre-eclampsia’ (Rumbold 2008). Our results are in accordance with the results form this review where any antioxidant supplementation compared to control or placebo had no effect on miscarriage or stillbirth.

**Authors’ Conclusions**

**Implications for practice**

There is no evidence to support the prophylactic use of single vitamins to prevent either early or late miscarriages. Supplementing women with multivitamin with or without iron and/or folic acid or vitamin A, may decrease the risk of total fetal loss and stillbirth. Even though there is a positive effect of multivitamin supplementation on pregnancy outcomes, there was insufficient evidence to examine the effect of different combinations of vitamins on miscarriage and miscarriage-related outcomes. Our findings suggest, that no particular vitamin decreases the risk of miscarriage or stillbirth, but the combination of various vitamins may have the potential to positively influence pregnancy outcomes. This could be due to an overall improvement in maternal nutrition and health status, making women more resistant to infections during pregnancy. However, this needs to be investigated further before recommendations on routine multivitamin supplementation to prevent miscarriage can be given.

**Implications for research**

The impact of different combinations of vitamins (i.e. individual vitamins or multivitamin preparations with or without vitamin A and folic acid) on miscarriage and miscarriage-related outcomes is unclear. Any future studies of vitamin supplementation should be high quality and focus on women at high risk of miscarriage. Considerations should include timing of the intervention and trials should assess the most appropriate vitamin type and dosage; to see whether it is beneficial without causing any harms to the mother or fetus and include assessments of any psychological effects and long-term follow-up of mothers and infants. Further, the data in the current review were complicated by differing definitions of miscarriage and so this may be an important issue to consider in any future trials.

**Acknowledgements**

For previous versions of the review, we thank Simon Gates for statistical advice regarding inclusion of cluster-randomised trials, Lelia Duley for helpful comments on the format of the review and Sonja Henderson for assisting with review administration.

We also thank Ning Pan, Caroline Crowther and Philippa Middleton for their contribution as authors on previous versions of the review.

We thank the following for their translation help: Izabella Brzostowicz and Agnieszka Kimball for Chelchowska 2004; Rebecca Gainey and Kate Bartos for Lira 1989; and Lucia Bartos for Météni 1996.

As part of the pre-publication editorial process, this review has been commented on by three peers (an editor and two referees who are external to the editorial team), a member of the Pregnancy and Childbirth Group’s international panel of consumers and the Group’s Statistical Adviser.

This project was supported by the National Institute for Health Research, via Cochrane Infrastructure funding to Cochrane Pregnancy and Childbirth. The views and opinions expressed therein are those of the authors and do not necessarily reflect those of the Systematic Reviews Programme, NIHR, NHS or the Department of Health.

Rintaro Mori’s institution receives government funding from the Clinical Research Program for Child Health and Development, AMED, Japan to provide support for the PCG Satellite in Japan.
REFERENCES

References to studies included in this review

Bhutta 2009  [published data only]

Briscoe 1959  [published data only]

Chappell 1999  [published data only]

Czeizel 1994  [published data only]


Ménéki J, Dudás I, Czeizel AE. Periconceptional multivitamin administration may result in higher frequency of twin pregnancies. *Orvosi Hetilap* 1996;137:2401–5.

Fawzi 1998  [published data only]


Fawzi 2007  [published data only]
Vitamin supplementation for preventing miscarriage (Review)

Copyright © 2016 The Cochrane Collaboration. Published by John Wiley & Sons, Ltd.

Fleming 1968 [published data only]

Fleming 1986 [published data only]

Hans 2010 [published data only]

Hemmi 2003 [published data only]

ICMR 2000 [published data only]

Jauniaux 2004 [published data only]

Katz 2000 [published data only]

Kirke 1992 [published data only]

Kumwenda 2002 [published data only]

McCance 2010 [published data only]
* Weisgerber T, Gandley R. Types of pregnancy associated hypertension according to haptoglobin phenotype. Reproductive Sciences 2011;18(3 Suppl 1):69A.
Vitamin supplementation for preventing miscarriage (Review)

Copyright © 2016 The Cochrane Collaboration. Published by John Wiley & Sons, Ltd.


MRC 1991 [published data only]

Osrin 2005 [published data only]

Osrin 2005 [published data only]

Poston 2006 [published data only]

Prawirohartono 2011 [published data only]

Robefroid 2008 [published data only]


Roberts 2010 [published data only]

Hauth J, for the Eunice Kennedy Shriver National Institute of Child Health and Human Development Maternal-Fetal Medicine Units Network. Vitamin C and E supplementation to prevent preterm premature rupture of


Rush 1980 [published data only]

Rush 2006 [published data only]

Schmidt 2001 [published data only]


Steyn 2003 [published data only]


Summit 2008 [published data only]

Sunawang 2009 *(published data only)*

Tofail 2008 *(published data only)*

Van den Broek 2006 *(published data only)*

Villar 2009 *(published data only)*


**West 2011 [published data only]**


**West 2014 [published data only]**


**Wibowo 2012 [published data only]**


**Xu 2010 [published data only]**


**Zagre 2007 [published data only]**


**Zeng 2008 [published data only]**


Zeng L, Yan H, Cheng Y, Dibley M. Modifying effects of maternal nutrition status on the response to multiple...


**References to studies excluded from this review**

Baumslag 1970 [published data only]


Biswas 1984 [published data only]


Blot 1981 [published data only]


Chanarin 1968 [published data only]


Chelchowska 2004 [published data only]


Christian 2003 [published data only]


Colman 1974 [published data only]


Correia 1982 [published data only]


Coutsoudis 1999 [published data only]


Dawson 1962 [published data only]

Vitamin supplementation for preventing miscarriage (Review)

Copyright © 2016 The Cochrane Collaboration. Published by John Wiley & Sons, Ltd.

Edelstein 1968 [published data only]

Ferguson 1955 [published data only]

Feyi-Waboso 2005 [published data only]

Fletcher 1971 [published data only]

Giles 1971 [published data only]

Hampel 1974 [published data only]

Hankin 1966 [published data only]

Hekmatdoost 2011 [published data only]
Hekmatdoost A. Comparison of the effect of folic acid and 5-methyltetrahydrofolate (5MTHF) on serum folate and homocysteine levels, and abortion rates in women suffering from recurrent abortion. *IRCT Iranian Registry of Clinical Trials* (www.irct.ir) (accessed 8 July 2011) 2011.

Hunt 1984 [published data only]

Huybregts 2009 [published data only]


Hibbard 2009 [published data only]

Hibbard 2010 [published data only]

Laurence 1981 [published data only]


Lin 2010 [published data only]
Lin JH, Yang YK, Liu H, Lin QD, Zhang WY. Effect of antioxidants on amelioration of high-risk factors inducing

Lira 1989 *(published data only)*


Lumeng 1976 *(published data only)*


Marya 1981 *(published data only)*


Meirinho 1987 *(published data only)*


Metz 1965 *(published data only)*


Mock 2002 *(published data only)*


Moldenhauer 2002 *(published data only)*


Owen 1966 *(published data only)*


Potdar 2014 *(published data only)*


Ross 1985 *(published data only)*


Schuster 1984 *(published data only)*


Semba 2001 *(published data only)*


Shu 2002 *(published data only)*


Smithells 1981 *(published data only)*


Suharno 1993 *(published data only)*


Tanumihardjo 2002 *(published data only)*

Tanumihardjo SA. Vitamin A and iron status are improved by vitamin A and iron supplementation in pregnant Indonesian women. *Journal of Nutrition* 2002;132:1009–12.

Taylor 1982 *(published data only)*

Vitamin supplementation for preventing miscarriage (Review)

Copyright © 2016 The Cochrane Collaboration. Published by John Wiley & Sons, Ltd.

Thauvin 1992  {published data only}

Trigg 1976  {published data only}

Ulrich 1999  {published data only}


Vutyavanich 1995  {published data only}

Villamor 2002  {published data only}

Wethington 1999  {published data only}

Young 2015  {published data only}

References to studies awaiting assessment

Adu-Afarwuah 2015  {published data only}

Agarwal 2012a  {published data only}

Frenzel 1956  {published data only}

Prado 2015  {published data only}

References to ongoing studies

Johns 2004  {published data only}

Additional references

Agarwal 2012b

Al-Gubory 2010

Haas 2009

Haas 2013

Hague 2003

Haider 2015

Hasan 2009

Higgins 2011

Hovdenak 2012

Hure 2012

Hübner 2008

Lassi 2013

Lede 2005

Vitamin supplementation for preventing miscarriage (Review)
of twin pregnancies. Multivitamin administration may result in higher frequency of miscarriage. 


Preston 1996


RevMan 2014 [Computer program]


Reznikoff-Etiévant 2002


Rivas-Echeverria 2000


Rumbold 2008


Rumbold 2015


Rumbold 2014


Wong 2014


Wu 2004


References to other published versions of this review

Rumbold 2005


Rumbold 2011

Rumbold A, Middleton P, Pan N, Crowther CA. Vitamin supplementation for preventing miscarriage. Cochrane
Vitamin supplementation for preventing miscarriage (Review)

Copyright © 2016 The Cochrane Collaboration. Published by John Wiley & Sons, Ltd.
### Characteristics of included studies  
*ordered by study ID*

#### Bhutta 2009

<table>
<thead>
<tr>
<th>Methods</th>
<th>Randomisation and allocation concealment: random allocation of the entire population of the urban and rural areas (population, 110,000; 20,400 households) into 28 discrete clusters (16 rural and 12 urban) on the basis of household characteristics, socioeconomic criteria, and geographic location. Each cluster was allocated to a community health worker who distributed the supplements on a cluster-based allocation strategy of supplements (“either iron-folic acid or multiple micronutrients”). Distribution of the sealed, coded supplement bottles were independently controlled by the pharmacy at Aga Khan University. Blinding of outcome assessment: medical officers, community health workers, social scientists, and data collection team remained blinded to the supplementation allocation. Documentation of exclusion: 373 women (16.5%) were excluded. Use of placebo control: no placebo given, women in the control group were given IFA.</th>
</tr>
</thead>
<tbody>
<tr>
<td>Participants</td>
<td>2378 women from community settings in urban and rural Sindh (Pakistan) less than 16 weeks of gestation. Eligible women were women with a confirmed pregnancy at less than 16 weeks of gestation. Women who did not have a confirmed pregnancy on ultrasound scanning or women who were clearly advanced beyond 24 weeks of gestation were excluded.</td>
</tr>
<tr>
<td>Interventions</td>
<td>Multiple micronutrients comprised 30 mg of iron (ferrous fumarate) and 400 mcg of folic acid along with 800 mcg of retinol (retinyl acetate), 200 IU of vitamin D (ergocalciferol), 10 mg of vitamin E (α-tocopherol acetate), 70 mg of ascorbic acid, 1.4 mg of vitamin B1 (thiamine mononitrate), 18 mg of niacin (niacinamide) 1.4 mg of vitamin B2, 1.9 mg of vitamin B6 (pyridoxine), 2.6 mcg of vitamin B12 (cyanocobalamin), 15 mg of zinc (zinc gluconate), 2 mg of copper, 65 mcg of selenium, and 150 mcg of iodine. Intervention was timed to start at less than 16 weeks' gestation. Comparison was iron (60 mg) and folic acid (400 mcg).</td>
</tr>
</tbody>
</table>
| Outcomes | Maternal outcomes:  
1. Blood Hb level as well as serum ferritin, zinc, and vitamin A.  
2. Physical and clinical examination, including a morbidity assessment and measurement of fundal height.  
3. Height, weight, and mid-upper-arm circumference.  
Infant outcomes:  
1. Birthweight.  
2. Gestational age.  
| Notes | Women’s risk of spontaneous and recurrent miscarriage is unclear.  
Womens’ BMI, Hb, ferritin, zinc, and serum retinol at admission are reported.  
Sample-size calculation reported by 2 methods: 1. based on a potential 5% gain in birthweight, 2. based to estimate a difference in birthweight of 150 g between the 2 groups.  
No intention-to-treat analyses performed.  
Compliance: community health worker performed a tablet count every fourth nightly. |
visit. Proportion of tablets consumed 75.65 in the intervention group and 76.7% in the control group. Location: urban population (Bilal Colony, Karachi) and rural villages (Kot Diji district, rural Sindh), Pakistan. Timeframe: unclear.

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>“We randomly allocated the entire population of the urban and rural area.” Pg S497</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>“Cluster-based allocation strategy of supplements (either iron-folic acid or multiple micronutrients) by the community health workers was implemented. The allocation of either iron-folic acid or multiple micronutrient supplements and the distribution of the sealed, coded supplement bottles were independently controlled by the pharmacy at Aga Khan University, which maintained the allocation codes by individual community health workers.” Pg S498</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>Low risk</td>
<td>All pregnant women were allocated a unique code and a uniquely labelled and numerically coded specific supplement supply for the duration of pregnancy. Blinding is unlikely to have been broken. Pg S498</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>Low risk</td>
<td>“The field staff (medical officers, community health workers, social scientists, and data collection team) remained completely blinded as to the supplement allocation.” Pg S498</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>Low risk</td>
<td>Approximately 16.5% of attrition with balanced number and similar reason for each group. S500 Figure1</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Unclear risk</td>
<td>No information about trial registration.</td>
</tr>
<tr>
<td>Other bias</td>
<td>High risk</td>
<td>The distribution of study participants across the urban and rural areas is unclear from the text and no adjustments were made for cluster design</td>
</tr>
</tbody>
</table>
### Briscoe 1959

#### Methods

Randomisation and allocation concealment: unclear, no methodological details given, dubious as the number of women allocated to the treatment group was more than double that allocated to the placebo group. "Unselected patients were each given 200 capsules... these were given a code, unknown to us and contained either an inert powder or 100 mg each of ascorbic acid and hesperidin."

Blinding of outcome assessment: women and study investigators did not know the treatment codes.

Documentation of exclusion: none reported.

Use of placebo control: placebo given; however, all women received an additional multivitamin supplement.

#### Participants

406 women were recruited in the study. Eligible women were "unselected patients" in private obstetrics care, who were less than or equal to 10 weeks' pregnant, and were eligible regardless of whether they were currently bleeding or the number of previous pregnancies. Women greater than 10 weeks' gestation were excluded. 406 women were randomised to either vitamin C (n = 303) or placebo (n = 103), no losses to follow-up were reported. 77 women in the study had more than 2 previous miscarriages and/or bleeding in the pregnancy, and 329 had 2 or fewer miscarriages and no bleeding in the pregnancy.

#### Interventions

All women were given 200 tablets, containing either 100 mg each of ascorbic acid and hesperidin or placebo (an inert powder). The study lasted for 7 weeks. For the first 2 weeks, women were asked to take 8 tablets daily (i.e. daily 800 mg each of vitamin C and hesperidin or placebo). For the following 5 weeks, women took 4 tablets daily (i.e. daily 400 mg each of vitamin C and hesperidin or placebo). All women received a multiple vitamin supplement containing 50 mg vitamin C.

#### Outcomes

1. Spontaneous miscarriage.
2. Spontaneous miscarriage in women with 2 or fewer previous miscarriages and no bleeding in the current pregnancy.
3. Spontaneous miscarriage in women with more than 2 previous miscarriages and/or bleeding in the current pregnancy.

#### Notes

Women's risk of spontaneous and recurrent miscarriage is unclear, as there is no information about concurrent medical conditions or other risk factors for miscarriage. 9 of the 406 women were classified as experiencing recurrent miscarriage. No information is available about women's nutritional status. No sample-size calculation reported.

Intention-to-treat analyses performed (no losses to follow-up reported). Compliance: no compliance information reported.

Location: Philadelphia, USA.

Timeframe: unclear.

---

### Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors' judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

Vitamin supplementation for preventing miscarriage (Review)
### Briscoe 1959  (Continued)

<table>
<thead>
<tr>
<th>Outcome</th>
<th>Risk Assessment</th>
<th>Risk Rating</th>
<th>Comment</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Unclear risk</td>
<td>No methodological details given.</td>
<td></td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Unclear risk</td>
<td>No methodological details given.</td>
<td></td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>Low risk</td>
<td>Women and study investigators did not know the treatment allocation</td>
<td></td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>Unclear risk</td>
<td>Double-blind study, but it is unclear who was blinded and if the code was broken before or after outcome assessment pg289</td>
<td></td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>Low risk</td>
<td>No losses to follow-up reported.</td>
<td></td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Unclear risk</td>
<td>Limited information about selection bias, stated that ‘unselected patients’ were included</td>
<td></td>
</tr>
<tr>
<td>Other bias</td>
<td>Unclear risk</td>
<td>Limited methodological details provided including patient compliance</td>
<td></td>
</tr>
</tbody>
</table>

### Chappell 1999

**Methods**

Randomisation and allocation concealment: a computer-generated randomisation list using blocks of 10 was given to the hospital pharmacy departments. Researchers allocated the next available number to participants and women collected the trial tablets from the pharmacy department.

Blinding of outcome assessment: women, caregivers and researchers were blinded to the treatment allocation until recruitment, data collection and laboratory analyses were complete.

Documentation of exclusion: 123 (43.5%) women were excluded, of which 70 women were withdrawn because their second Doppler scan was normal. Pregnancy outcome data were reported for all women randomised.

Use of placebo control: placebo control.

**Participants**

283 women were recruited into the study. Inclusion criteria: abnormal Doppler waveform in either uterine artery at 18-22 weeks’ gestation or a history in the preceding pregnancy of pre-eclampsia necessitating delivery before 37 weeks’ gestation, eclampsia or the syndrome of HELLP.

Exclusion criteria: heparin or warfarin treatment, abnormal fetal-anomaly scan or multiple pregnancy.

Women were randomised at 18-22 weeks’ gestation; however, women with a previous history who were identified at an earlier stage were randomised at 16 weeks’ gestation. Women with abnormal Doppler waveform analysis returned for a second scan at 24 weeks’ gestation, those with a normal waveform at this time stopped treatment and were withdrawn from the study. The remaining women who had persistently abnormal...
waveforms, and those with a previous history or pre-eclampsia remained in the study and were seen every 4 weeks through the rest of pregnancy. 1512 women underwent Doppler screening, 273 women had abnormal waveforms and of these, 242 women consented to the study. An additional 41 women who had a history of pre-eclampsia consented. 283 women were randomised to either the vitamin C and E group (n = 141) or the placebo group (n = 142), 72 women had normal Doppler scans at 24 weeks' gestation and 24 women did not return for a second scan and were withdrawn. A further 27 women withdrew from the trial after 24 weeks' gestation for various reasons. In total, 160 women completed the trial protocol until delivery, 79 in the vitamin C and E group and 81 in the placebo group. Pregnancy outcome data were presented for all women randomised (n = 283) as well as only for those women completing the trial protocol (n = 160).

### Interventions

Women randomised to the vitamin C and E group received tablets containing 1000 mg vitamin C daily and capsules containing 400 IU vitamin E daily. Women randomised to the placebo group received tablets containing microcrystalline cellulose and soya bean oil, that were identical in appearance to the vitamin C tablets and vitamin E capsules. After 24 weeks' gestation women were seen every 4 weeks, and blood samples were taken at each visit.

### Outcomes

1. Ratio of PAI-1 to PAI-2.
2. Incidence of pre-eclampsia.
3. Placental abruption.
4. Spontaneous preterm delivery (< 37 weeks).
5. Intrauterine death.
6. Small-for-gestational-age infants (on or below the 10th centile).
7. Mean systolic and diastolic blood pressure before delivery.
8. Gestational age at delivery (median, IQR).
10. Birthweight centile (median, IQR).

### Notes

Women's risk of spontaneous and recurrent miscarriage is unclear, women were at high risk of pre-eclampsia.

No information is available about women's nutritional status.

Sample-size calculation reported, based on a 30% reduction in PAI-1.

Intention-to-treat analyses performed.

Compliance: "within the treated group, plasma ascorbic acid concentration increased by 32% from baseline values and plasma alpha-tocopherol increased by 54%".

Location: London, UK.

Timeframe: unclear.

### Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors' judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>Computer-generated random number list.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>Random number list used blocks of 10 and was held by the pharmacy department</td>
</tr>
</tbody>
</table>
### Chappell 1999 (Continued)

<table>
<thead>
<tr>
<th>Blinding of participants and personnel (performance bias)</th>
<th>Low risk</th>
<th>Women, caregivers and researchers were blinded until the analyses were completed</th>
</tr>
</thead>
<tbody>
<tr>
<td>All outcomes</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Blinding of outcome assessment (detection bias)</th>
<th>Low risk</th>
<th>“The code was revealed to the researchers once recruitment, data collection, and laboratory analyses were complete” pg811</th>
</tr>
</thead>
<tbody>
<tr>
<td>All outcomes</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Incomplete outcome data (attrition bias)</th>
<th>Low risk</th>
<th>123 (43.5%) women were excluded, of which, 70 women were withdrawn because their second Doppler scan was normal. Data were reported for all women randomised</th>
</tr>
</thead>
<tbody>
<tr>
<td>All outcomes</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Selective reporting (reporting bias)</th>
<th>Low risk</th>
<th>Data reported for all outcomes in methods.</th>
</tr>
</thead>
<tbody>
<tr>
<td>All outcomes</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Other bias</th>
<th>Low risk</th>
<th>The study appears to be free of other sources of bias.</th>
</tr>
</thead>
<tbody>
<tr>
<td>All outcomes</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

### Czeizel 1994

<table>
<thead>
<tr>
<th>Methods</th>
<th>Randomisation and allocation concealment: unclear, “women agreed to their allocation on the basis of a random table”</th>
</tr>
</thead>
<tbody>
<tr>
<td>Blinding of outcome assessment: unclear, women were aware of the “blind use of one of two kinds of tablets”, but no other details given</td>
<td></td>
</tr>
<tr>
<td>Documentation of exclusion: 49 women (1%) were lost to follow-up and excluded</td>
<td></td>
</tr>
<tr>
<td>Use of placebo control: “trace element control” given.</td>
<td></td>
</tr>
<tr>
<td>All outcomes</td>
<td></td>
</tr>
</tbody>
</table>

| Participants                                             | 7765 women were recruited into the study. Women participating in the HOFPP who volunteered to take part, were not currently pregnant, and who conceived within 12 months of ceasing contraception. In the first 2 years of the HOFPP, women were also required to be aged < 35 years, and not to have had a previous pregnancy except a prior induced abortion. 7905 women were approached, of which 140 refused participation, 7765 were randomised and 5502 women had a confirmed pregnancy and were allocated to either multivitamins (n = 2819) or control (n = 2683). 49 women of the 5502 confirmed pregnancies were lost to follow-up |
|----------------------------------------------------------|----------|----------------------------------------------------------------------------------------------------------------------------------|
| All outcomes                                             |          |                                                                                                                                |

<table>
<thead>
<tr>
<th>Interventions</th>
<th>Women were provided with multivitamin or trace element ‘control’ from at least 28 days before conception continuing until at least the second missed menstrual period.</th>
</tr>
</thead>
<tbody>
<tr>
<td>The multivitamin with folic acid contained 6000 IU vitamin A, 1.6 mg vitamin B1, 1.8 mg vitamin B2, 2.6 mg vitamin B6, 4.0 mcg vitamin B12, 100 mcg vitamin C, 500 IU vitamin D, 15 mg vitamin E, 19 mg nicotinamide, 10 mg calcium pantothenate, 0.2 mg biotin, 0.8 mg folic acid, 125 mg calcium, 125 mg phosphorus, 100 mg magnesium, 60 mg iron, 1 mg copper, 1 mg manganese, 7.5 mg zinc. The trace element control contained 7.5 mg vitamin C, 1 mg copper, 1 mg manganese and 7.5 mg zinc</td>
<td></td>
</tr>
<tr>
<td>----------------------------------------------------------</td>
<td>----------------------------------------------------------------------------------------------------------------------------------</td>
</tr>
<tr>
<td>All outcomes</td>
<td></td>
</tr>
</tbody>
</table>
Outcomes

1. NTDs and other birth defects.
2. Miscarriage.
3. Ectopic pregnancy.
4. Termination of pregnancy.
5. Live births.
7. Multiple gestation.
8. Subgroup data are available on menstrual cycle, first trimester symptoms and sexual activity.

Notes

Women's risk of spontaneous and recurrent miscarriage is unclear.
Information on their dietary status is unknown.
No sample-size calculation reported.
Partial intention-to-treat analyses performed.
Compliance: compliance was assessed by questioning, checking the tick-off on the basal temperature chart and counting of unused tablets. 70% of women in the multivitamin group and 71% in the control group took the full course of the supplements, with an additional 20% and 21% in the multivitamin and control groups respectively receiving a partial course of supplementation.
Location: Hungary.
The denominators used for this trial are the number of women randomised and with a confirmed pregnancy (i.e. 2819 for the multivitamin group and 2683 for the control group).

Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Unclear risk</td>
<td>Methodological details unclear.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Unclear risk</td>
<td>Methodological details unclear, 'women agreed to their allocation on the basis of a random table'</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>Unclear risk</td>
<td>Women were aware of the 'blind use of one of two kinds of tablets', but no other details given</td>
</tr>
<tr>
<td>All outcomes</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>Unclear risk</td>
<td>No details are given if outcome assessment was blinded.</td>
</tr>
<tr>
<td>All outcomes</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>Low risk</td>
<td>49 women (1%) excluded, partial intention-to-treat analyses performed</td>
</tr>
<tr>
<td>All outcomes</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>
Czeizel 1994  (Continued)

<table>
<thead>
<tr>
<th>Selective reporting (reporting bias)</th>
<th>Unclear risk</th>
<th>Denominators vary with serial publications.</th>
</tr>
</thead>
<tbody>
<tr>
<td>Other bias</td>
<td>Unclear risk</td>
<td>Limited methodological details provided.</td>
</tr>
</tbody>
</table>

Fawzi 1998

Methods

Randomisation and allocation concealment: block randomisation using blocks of 20, eligible women were “assigned the next numbered bottle of regimen”. The study used a 2 by 2 factorial design and women were randomised to 1 of 4 groups. Tablets were indistinguishable and packaged in identically coded bottles.

Blinding of outcome assessment: women and study investigators were unaware of the treatment allocation, no information given about blinding of outcome assessors.

Documentation of exclusion: 64 women (6%) were lost to follow-up and excluded.

Use of placebo control: placebo given.

Participants

1085 women were recruited into the study. Pregnant women between 12 and 27 weeks’ gestation who were HIV-1 infected, living in Dar es Salaam and intended to stay there for at least 1 year were eligible for the study. Women not HIV-1 positive or moving out of Dar es Salaam were excluded. 13,879 pregnant women consented to be HIV-1 tested, of which 1806 were positive, and 1085 were randomised. Of these, 3 women were not pregnant and 7 women died before delivery and were excluded from the trial. Of the remaining 1075 women, 54 women (5%) were lost to follow-up by the time of delivery, leaving birth outcomes reported for 1021 women. Women were randomised to 1 of 4 groups: vitamin A (n = 269), multivitamins excluding vitamin A (n = 269); multivitamins including vitamin A, all formulated in 2 tablets; or placebo (n = 267).

Interventions

Women were randomised to 1 of 4 groups:

1. vitamin A (30 mg beta-carotene plus 5000 IU preformed vitamin A);
2. multivitamins excluding vitamin A (20 mg vitamin B1, 20 mg vitamin B2, 25 mg vitamin B6, 100 mg niacin, 50 mcg vitamin B12, 500 mg vitamin C, 30 mg vitamin E, 0.8 mg folic acid);
3. multivitamins including vitamin A, all formulated in 2 tablets; or
4. placebo.

All women received 400 mg ferrous sulphate and 5 mg folic acid daily, as well as 500 mg chloroquine phosphate weekly. At delivery, all women taking vitamin A were to receive an additional oral dose of 200,000 IU vitamin A and the others an extra dose of a placebo. Pill counts were conducted at each visit and new tablets were given out at each visit.

Outcomes

1. Miscarriage, defined as delivery before 28 weeks’ gestation.
2. Stillbirth, defined as delivery of a dead baby at or after 28 weeks’ gestation.
3. Fetal death, defined as either miscarriage or stillbirth.
4. Low birthweight, defined as birthweight less than 2500 g.
5. Very low birthweight, defined as birthweight less than 2000 g.
6. Preterm delivery, defined as delivery before 37 weeks.
7. Severe preterm birth, defined as delivery before 34 weeks.
8. Small-for-gestational age, defined as birthweight less than the 10th percentile for gestational age.
Fawzi 1998  *(Continued)*

<table>
<thead>
<tr>
<th>Notes</th>
<th>Women’s risk of spontaneous and recurrent miscarriage was unclear, although may be increased due to their HIV-1 positive status. Women’s nutritional status is also unclear. Figures change with serial publications, particularly for secondary outcomes, and results are not reported separately for the individual 4 groups. Results are reported as: any multivitamins, multivitamin, any vitamin A or no vitamin A. Sample-size calculation performed allowing for 20% loss to follow-up. Intention-to-treat analyses performed. Compliance: compliance assessed by the percentage of prescribed tablets absent from the returned bottles, and in plasma vitamin A concentrations in a subset of 100 women. Median compliance assessed using pill counts was 90% by the time of delivery. Location: Tanzania. Timeframe: April 1995 to July 1997.</th>
</tr>
</thead>
</table>

**Risk of bias**

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Unclear risk</td>
<td>Block randomisation using blocks of 20.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Unclear risk</td>
<td>Women assigned the ‘next numbered bottle of regimen’.</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>Low risk</td>
<td>Women and investigators were blinded.</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>Unclear risk</td>
<td>Double-blinded study, but unclear if outcome assessors were blinded</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>High risk</td>
<td>64 women (6%) were lost to follow-up and excluded, intention-to-treat analyses performed</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Unclear risk</td>
<td>Figures change with serial publications, particularly for secondary outcomes, and results are not reported separately for the individual 4 groups</td>
</tr>
<tr>
<td>Other bias</td>
<td>Unclear risk</td>
<td>Limited methodological details provided.</td>
</tr>
</tbody>
</table>
### Methods

Randomisation: unclear about sequence generation.
Allocation concealment: states a list was prepared according to the randomisation sequence in blocks of 20, tablets were bottled in identical coded bottles, eligible women were given the next numbered bottle.
Blinding of outcome assessment: women and research assistants who assessed the study outcomes were unaware of the intervention groups.
Documentation of exclusion: 49 women lost to follow-up (multivitamin group: 23, placebo group: 26), no post-randomisation exclusions.
Use of placebo control: placebo given.

### Participants

8428 women were randomised in the study. Pregnant women between 12 and 27 weeks who had a negative test for HIV infection and planned to stay in the city until delivery and for 1 year thereafter recruited through antenatal clinics in Dar es Salaam. 8468 women were enrolled, however 40 women were then found to be ineligible. 8428 women were randomly assigned to receive either a multivitamin (n = 4214) or placebo (n = 4214) from the time of enrolment until 6 weeks after delivery. 6 women died before delivery and 43 were lost to follow-up by the time of delivery.

### Interventions

The supplements included 20 mg of vitamin B1, 20 mg of vitamin B2, 25 mg of vitamin B6, 100 mg of niacin, 50 mcg of vitamin B12, 500 mg of vitamin C, 30 mg of vitamin E, and 0.8 mg of folic acid.

The active tablets and placebo were similar in shape, size, and colour.
All women, irrespective of the assigned study regimen, were given daily doses of iron (60 mg of elemental iron) and folic acid (0.25 mg). They were also given malaria prophylaxis in the form of sulfadoxine-pyrimethamine tablets at 20 weeks and 30 weeks of gestation.

### Outcomes

1. Low birthweight (< 2500 g).
2. Preterm delivery (before 37 weeks’ gestation).
3. Fetal death.
4. Birthweight below 2000 g.
5. Extremely preterm delivery (before 34 weeks).
6. Small-for-gestational age (birthweight below the 10th percentile for gestational age).
7. Fetal death and death in the first 6 weeks of life.

### Notes

Women’s risk of spontaneous and recurrent miscarriage was unclear.
Women’s nutritional status is also unclear.
Intention-to-treat analyses performed.
Compliance: average compliance was 88%, no difference in compliances between the 2 groups.
Location: Tanzania.

### Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Unclear risk</td>
<td>Generation of sequence not reported, except that there were blocks of 20 in the se-</td>
</tr>
</tbody>
</table>
### Fawzi 2007 (Continued)

<table>
<thead>
<tr>
<th>Bias Type</th>
<th>Risk Level</th>
<th>Description</th>
</tr>
</thead>
<tbody>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>Identical coded bottles prepared according to the randomisation list, eligible women were assigned the next numbered bottle</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>Low risk</td>
<td>Women and outcome assessors were blinded to allocation.</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>Low risk</td>
<td>“Research assistants who assessed the study outcomes were unaware of the intervention groups.” pg1424</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>Low risk</td>
<td>49 (1%) women lost to follow-up, balanced across groups, analyses by intention-to-treat</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Low risk</td>
<td>All pre-specified outcomes appear to be reported.</td>
</tr>
<tr>
<td>Other bias</td>
<td>Low risk</td>
<td>The study appears to be free of other sources of bias.</td>
</tr>
</tbody>
</table>

### Fleming 1968

**Methods**

Randomisation and allocation concealment: quasi-randomised, alternate women were allocated to receive folic acid or placebo according to the order in which they attended antenatal clinic. No other methodological details were given.

Blinding of outcome assessment: women and investigators were blinded to the treatment allocation, until after the completion of the trial.

Documentation of exclusion: 21 women (28%) excluded from the analysis.

Use of placebo control: control tablet containing iron given.

**Participants**

75 women were recruited into the trial. Women were eligible if they were primigravida, less than 26 weeks’ pregnant (range of gestation 10 to 26 weeks), with haematocrit value (PCV) 27% or more, and who had not received treatment so far as was known. Women with Hb SC, Hb SS, Hb CC were excluded. Alternate patients were allocated to group A (placebo) or B (folic acid). 75 women were included (40 in group A and 35 in group B) initially; however, only 26 in group A and 28 in group B completed the trial. 16 women (10 in group A and 8 in group B) defaulted from the trial, 3 (2 in group A and 1 in group B) were anaemic on the second visit warranting folic acid treatment, 1 in group A self-medicated with folic acid and 1 in group A ‘aborted’.

**Interventions**

All women received antimalarials and iron supplements as per the standard antenatal care at the hospital.

Women in group B received 5 mg folic acid tablets on each attendance, which was fortnightly initially and weekly in the last trimester.
Group A received “one tablet of lactose base and colouring matter in the same manner”

### Outcomes

1. PCV and reticulocyte index.
2. Serum folic acid concentration and ‘megaloblastic score’.
3. Malarial infection.
4. Maternal morbidity (pyelonephritis, pre-eclamptic toxaemia, septicaemia, puerperal psychosis).
5. Prematurity.
6. Birthweight (mean birthweight but no standard deviation).
7. Fetal mortality.

### Notes

Results not reported as intention-to-treat; however, where possible, the review authors included data in the review as intention-to-treat. Unclear of women’s risk of spontaneous and recurrent miscarriage. 16 women in the trial showed evidence of folic acid deficiency at trial entry. Sample-size calculation: none reported. No intention-to-treat analyses performed. Compliance: no compliance information reported specifically; however, women were “seen to swallow” the tablets at their fortnightly and weekly visits. Location: Nigeria. Time frame: unclear.

### Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>High risk</td>
<td>Quasi-randomised, alternate allocation.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>High risk</td>
<td>Quasi-randomised, alternate allocation.</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>Low risk</td>
<td>Women and investigators blinded.</td>
</tr>
<tr>
<td>All outcomes</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>Low risk</td>
<td>“The identity of the tablets was not known to investigators until after the completion of the trial.” pg426</td>
</tr>
<tr>
<td>All outcomes</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>High risk</td>
<td>21 women (28%) excluded from the analysis.</td>
</tr>
<tr>
<td>All outcomes</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Unclear risk</td>
<td>Results not reported as intention to treat; however, where possible, the review authors included data in the review as intention to treat</td>
</tr>
<tr>
<td>Other bias</td>
<td>Unclear risk</td>
<td>Limited methodological details provided.</td>
</tr>
</tbody>
</table>
### Methods

Randomisation and allocation concealment: women were “randomly allocated to one of five groups using a random number table”, no other details given.

Blinding of outcome assessment: women and investigators were blinded to the treatment allocation, until after the completion of the trial.

Documentation of exclusion: 18 women (9%) were excluded due to anaemia at enrolment, ‘defaulting’, or being ‘mentally subnormal’, these women were replaced by other women chosen by an investigator. A further 42 women were excluded before delivery and another 30 failed to attend the postnatal clinic, birth outcomes were available for 160 women (80%).

Use of placebo control: no placebo control.

### Participants

228 women met the eligibility criteria; however 200 pregnant women were recruited into the study. Women were allocated to 1 of 5 groups; 40 women were allocated to each group.

Eligible women included:
1. Hausa women living in Zaria and planning to deliver in Zaria;
2. pregnant for the first time;
3. at less than 24 weeks’ gestation, as estimated by the height of the fundus uteri;
4. the wives of unskilled or semiskilled men.

Women were excluded if they had already taken any antimalarial treatment or haematinics during the pregnancy, or had the following complications: hydatiform mole, Hb SC disease, overt anaemia or proteinuria.

The mean gestational age of women at enrolment was 18.5 weeks.

### Interventions

Women were allocated to 1 of 5 groups:
- group 1: no active treatment (control);
- group 2: antimalarials only (600 mg chloroquine/day + 100 mg proguanil/day);
- group 3: iron + antimalarials (60 mg iron/day + 600 mg chloroquine/day + 100 mg proguanil/day);
- group 4: folic acid + antimalarials (1 mg folic acid/day + 600 mg chloroquine/day + 100 mg proguanil/day);
- group 5: iron + folic acid + antimalarials (1 mg folic acid/day + 60 mg iron/day + 600 mg chloroquine/day + 100 mg proguanil/day).

### Outcomes

**Maternal outcomes**
2. Gestation age.
3. Mode of delivery.
4. Complications of pregnancy (abortion, hypertension, pre-eclampsia or eclampsia, hydramnios, abdominal pain).

**Infant outcomes**
1. Fetal distress.
2. Birthweight.
3. Apgar score at 2 minutes.
4. Fetal complications.

**Laboratory outcomes**
1. Hb concentration, red cell indices and WBC at first attendance, 28 weeks, 36 weeks, at delivery (form mother and infant) and 6 weeks postpartum.
Not all outcomes were reported for each individual treatment group. Miscarriage was reported for the combined groups 4 and 5, therefore for the purpose of this review the groups 4 and 5 are combined (folic acid + iron) and compared with group 2 and group 3 (iron + antimalarials). The authors reported that 8 women had hypertension without other signs, 21 women had pre-eclampsia and 6 developed eclampsia, with no association between these outcomes and treatment group. No other details were provided, including the breakdown of these outcomes by treatment group.

Notes

Women’s risk of spontaneous and recurrent miscarriage was unclear
Women were at high risk of anaemia. Information about other nutritional indices was not provided
Intention-to-treat analyses not performed, however, where possible, the review authors included data in the review as intention-to-treat
Compliance: 72 women (36%) were classed as defaulters.
Location: Nigeria.
Timeframe: unclear.

Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Unclear risk</td>
<td>A random number table was used but no details provided of how it was generated</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Unclear risk</td>
<td>No details provided about the allocation.</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>Low risk</td>
<td>Neither the researchers nor the patients were aware of the treatment allocation until after the completion of the study</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>Low risk</td>
<td>“Neither the researchers nor the patients were aware of the treatment allocation until after the completion of the study,” pg 214</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>High risk</td>
<td>228 women met the entry criteria, but only 200 were included in the trial. 18 women were excluded and replaced by other women</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>High risk</td>
<td>Not all outcomes are reported by treatment group. In serial publications up to 70% of the data were excluded</td>
</tr>
<tr>
<td>Other bias</td>
<td>Unclear risk</td>
<td>Limited methodological details provided.</td>
</tr>
</tbody>
</table>
Randomisation and allocation concealment: women were randomly assigned to receive either 400 mg of vitamin C daily or not in addition to their standard antenatal vitamin. Randomisation was obtained by a computer-generated, block design sequence to receive vitamin C or not in a 1:1 ratio. No other methodological details given.

Blinding of outcome assessment: unclear, no details given.

Documentation of exclusion: 16 women (4%) did not complete the trial after randomisation.

Use of placebo control: no placebo control.

400 women 4 to 12 gestational weeks of pregnancy confirmed serologically by B-HCG reagent test along with referred last menstrual period not exceeding the past 3 months, aged at least 18. Women with referred pregnancy of more than 3 months by last menstrual period, concomitant HIV infection status, active or recent (< 2 weeks) sexually transmitted disease infection, medical record of any severe organ disease such as heart, liver or renal failure at the time of assessment, diagnosis of pregnancy during inpatient admission for any other reason, recent history of multivitamin supplementation (< 12 weeks) for any reason, except for pregnancy, and patients incapable to read and write.

Chewable tablet of synthetic form of L-ascorbic acid or vitamin C 400 mg administered daily 2 tablet 2 times a day, from first trimester until delivery. Comparison received no vitamin C in addition to their standard antenatal vitamin.

All women received ferrous sulphate 200 mg, folic acid 5 mg and vitamin B-complex 60 mg once daily tablets. Nutritional counselling was provided to all women.

2. Overall hospitalisations rate.
3. Weight gain during pregnancy (normal < 16 kg).
4. Term pregnancy (≥ 37 gestational weeks).
5. Preterm delivery.
6. miscarriage (< 24 gestational weeks).
7. Low birthweight (< 2500 g).
8. Gestational systolic blood pressure.

Women’s risk of spontaneous or recurrent miscarriage is unclear, as is their dietary intake.

Sample-size calculation based on at least 30% of women not hospitalised during pregnancy in the control group and at least 50% of women not hospitalised in the intervention group.

Analyses were not based on intention-to-treat.

Compliance: no details of any compliance assessments were given.

Location: Kyeibuza, Uganda.

Timeframe: August 2007 and January 2009.

**Risk of bias**

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>“Patients presented to the health centre, those who met the inclusion criteria were randomly assigned...randomiza-</td>
</tr>
</tbody>
</table>
### Hans 2010  *(Continued)*

<table>
<thead>
<tr>
<th>Bias Type</th>
<th>Risk</th>
<th>Comments</th>
</tr>
</thead>
<tbody>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Unclear</td>
<td>No description for allocation concealment in the text. Probably not done</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>Unclear</td>
<td>No information on participant or personnel blinding provided</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>Unclear</td>
<td>No information on outcome assessor blinding provided.</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>Unclear</td>
<td>Drop-out rates for each group were not clearly described and no information was provided on the reason of exclusion and attrition</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Unclear</td>
<td>Although expected outcomes were reported, the study protocol was not available to see if all prespecified outcomes were reported as planned</td>
</tr>
<tr>
<td>Other bias</td>
<td>Low risk</td>
<td>The study appears to be free of other sources of bias.</td>
</tr>
</tbody>
</table>

### Hemmi 2003

**Methods**

Randomisation and allocation concealment: unclear, “patients were randomly assigned to the control group or the study group”. No other methodological details given.

Blinding of outcome assessment: unclear, no details given.

Documentation of exclusion: 28 women (19%) in the control group were excluded, no details given for the exclusion.

Use of placebo control: no placebo control.

**Participants**

150 women were recruited into the study. Women with a luteal phase defect, as described by a peak serum P level < 120 mg/mL in the mid-luteal phase measured at 3 time points, were eligible and invited to participate. Luteal phase defects were ascertained in 2 consecutive menstrual cycles, and the third cycle was the intervention cycle. Women receiving IVF-ET treatment were excluded. 313 women were considered for enrolment in the study, 150 (48%) were randomised. 28 women were withdrawn from the control group, leaving 122 women in the study, who were allocated to vitamin C (n = 76) or control (n = 46). 5 women in the control group and 19 women in the vitamin C group became pregnant during the study period.
**Interventions**

Women in the intervention group took 750 mg vitamin C per day from the first day of the third menstrual cycle until a urinary pregnancy test was positive. Pregnancy rate was checked up until 6 months after the study cycle was started. Women in the control group received no supplementation and no treatment was given in the third cycle.

**Outcomes**

1. Serum P concentrations.
2. Serum E2 (oestrogen) concentrations.
3. Pregnancy rate.

**Notes**

Women's risk of spontaneous or recurrent miscarriage was unclear according to criteria specified in the review.

Their dietary intake of vitamin C is unknown.

No sample-size calculation was reported.

Analyses were not based on intention to treat.

Compliance: no details of any compliance assessments were given.

Country: Japan.


The denominators used for this trials are the number of women randomised and with a confirmed pregnancy (i.e. 19 for the vitamin group and 5 for the control group).

**Risk of bias**

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors' judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Unclear risk</td>
<td>Methodological details unclear.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Unclear risk</td>
<td>Methodological details unclear.</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias) All outcomes</td>
<td>Unclear risk</td>
<td>Methodological details unclear.</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias) All outcomes</td>
<td>Unclear risk</td>
<td>No methodological details are given.</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias) All outcomes</td>
<td>High risk</td>
<td>28 women (19%) in the control group excluded.</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Unclear risk</td>
<td>No details of exclusion of women in the control group given.</td>
</tr>
<tr>
<td>Other bias</td>
<td>High risk</td>
<td>No placebo control.</td>
</tr>
</tbody>
</table>
### ICMR 2000

| Methods | Randomisation and allocation concealment: unclear, “containers of vitamin or placebo capsules were given a random number” and “the key to random numbers was kept at the ICMR Headquarters”. No other methodological details were given. Blinding of outcome assessment: “double blind” mentioned in the text, but no details given. Documentation of exclusion: 187 women (40%) were excluded from the analysis. Use of placebo control: placebo control. |
| Participants | 466 women were recruited into the study. Women who had previously given birth to a child with an open NTD, and planned to have another child were eligible and invited to participate. This was regardless of their parity, number of previous births with an NTD, age, consanguinity, and socioeconomic status. Women who had previously given birth to a child with closed spina bifida, or with a history of diabetes or abnormal fasting and post-prandial blood sugar, history of epilepsy, congenital anomalies indicative of a genetic syndrome in the previous NTD, history of vitamin intake in the 3 months prior to enrolment, and pregnancy were excluded. 466 women were enrolled and randomised to either vitamin (n = 231) or placebo (n = 235), of these women, 90 were lost to follow-up immediately and 71 did not conceive until the final follow-up. Of the remaining 305 women who were known to become pregnant (vitamin n = 152, placebo n = 153), pregnancy outcomes were unknown for 26 women. In the paper, 279 of the initial 466 women were included in the analysis; however, in this review results are presented for main outcomes on an intention-to-treat basis (i.e. n = 466). |
| Interventions | The folic acid containing multivitamin included 120 mg ferrous sulphate, 240 mg calcium phosphate, 4000 IU vitamin A, 400 IU vitamin D, 2.5 mg vitamin B1, 2.5 mg vitamin B2, 2 mg vitamin B6, 15 mg nicotinamide, 40 mg vitamin C, 4 mg folic acid, 10 mg zinc. The placebo tablets contained the following trace elements: 120 mg ferrous sulphate and 240 mg calcium phosphate. Both capsules were identical in appearance and women were provided with the tablets from at least 28 days before conception and continuing until at least the second missed menstrual period. |
| Outcomes | 1. Recurrence of NTDs. 2. Live births. 3. Stillbirths. 4. Spontaneous and induced abortion. 5. Multiple birth. |
| Notes | The risk profile of women in the trial for spontaneous and recurrent miscarriage is unclear, as is the dietary intake of participants. Sample-size calculation performed, assuming a 20% drop out rate. The trial was terminated after publication of the MRC trial in 1991. Compliance: compliance was assessed at 3-monthly visits, by checking a diary card maintained by the woman and the number of capsules returned. If the total number of missed days in 3 months did not exceed 10 days, and the total number of missed days at a stretch did not exceed 3, compliance was taken as satisfactory. Women not meeting the above criteria were excluded if they became pregnant in that particular quarter. No compliance data are specifically reported. Analyses not based on intention-to-treat. |
ICMR 2000  (Continued)

| Country: India. Time frame: 1988 to 1991. The denominators used for this trial are based on the number of women randomised (i.e. 231 for the vitamin group and 235 for the placebo group). There was not enough information to accurately confirm the number of women that did or did not become pregnant due to the large number of losses to follow-up |

<p>| Risk of bias |</p>
<table>
<thead>
<tr>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Unclear risk</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Unclear risk</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias) All outcomes</td>
<td>Unclear risk</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias) All outcomes</td>
<td>Unclear risk</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias) All outcomes</td>
<td>High risk</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Unclear risk</td>
</tr>
<tr>
<td>Other bias</td>
<td>Unclear risk</td>
</tr>
</tbody>
</table>

Jauniaux 2004

| Methods | Randomisation and allocation concealment: randomised controlled trial, but no other information provided Blinding of outcome assessment: unclear. Documentation of exclusions: unclear, no information provided Use of placebo: placebo control. |
| Participants | Women with a history of 2 or more early pregnancy losses, with no identifiable cause for the losses |
| Interventions | Vitamin C 1000 mg and vitamin E 400 IU versus placebo. |
| Outcomes | Miscarriage. |
Notes

Updated 27/11/2013: the trial was stopped in October 2009 due to poor recruitment and lack of funding.
Women's risk of spontaneous and recurrent miscarriage is unclear.
Sample-size calculation: not done.
No intention-to-treat analyses: not done.
Compliance: unclear.
Location: UK.

Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors' judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Unclear risk</td>
<td>Could not be assessed because the trial was stopped.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Unclear risk</td>
<td>Could not be assessed because the trial was stopped.</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>Unclear risk</td>
<td>Could not be assessed because the trial was stopped.</td>
</tr>
<tr>
<td>All outcomes</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>Unclear risk</td>
<td>Could not be assessed because the trial was stopped.</td>
</tr>
<tr>
<td>All outcomes</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>Unclear risk</td>
<td>Could not be assessed because the trial was stopped.</td>
</tr>
<tr>
<td>All outcomes</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Unclear risk</td>
<td>Could not be assessed because the trial was stopped.</td>
</tr>
<tr>
<td>Other bias</td>
<td>Unclear risk</td>
<td>Could not be assessed because the trial was stopped.</td>
</tr>
</tbody>
</table>
**Katz 2000**

| Methods | Randomisation and allocation concealment: cluster-randomised. 270 centres in the Salarhi district, Nepal, were involved which included 30 subdistricts each with 9 wards. Each ward was assigned to 1 of 3 treatment groups. "Wards were assigned by a random draw of numbered chits, blocked on subdistrict." Blinding of outcome assessment: women and study investigators were not aware of the treatment codes. Maternal mortality was assessed by study investigators blinded to treatment allocation, no details were given for other outcomes. Documentation of exclusions: 157 (1%) women were lost to follow-up and excluded. Use of placebo: placebo control. |
| Participants | 15,832 women were recruited into the study. All married women of child bearing age in the Salarhi district, Nepal, were eligible and invited to participate in the study. Women migrating into the study area, or women that were never pregnant or refused participation, or women who migrated before being pregnant, were excluded from the analysis. Eligible women were identified from census data and marriage registers. 44,646 women were recruited, of which 1136 (2.5%) were excluded as they either emigrated before becoming pregnant, died or refused consent. During the study period 15,832 women identified themselves as being pregnant, and 157 women were lost to follow-up in the postpartum period. Results are reported for 17,373 pregnancies, allocated to the following groups: vitamin A (n = 6070), beta-carotene (n = 5650) or placebo (n = 5653). Denominators for the treatment groups vary for the measures of early infant mortality, due to losses to follow-up after birth. |
| Interventions | The 3 treatment groups consisted of a weekly single oral supplement of either: 1. 23,300 IU preformed vitamin A as retinyl palmitate; 2. 42 mg of all trans beta-carotene; 3. placebo. All capsules contained mg dl-alpha-tocopherol as an antioxidant. Women took the tablets prior to conception, during pregnancy and postpartum, for a total of 3.5 years. |
| Outcomes | 1. Fetal loss, defined as any reported miscarriage, stillbirth or maternal death during pregnancy. The outcomes were based on self-reports, and women who reported to be pregnant for >= 6 weeks but then no longer reported being pregnant were considered to have had a miscarriage. Serial publications also reported neonatal death. |
| Notes | Women's risk profile for spontaneous or recurrent miscarriage was unclear, as was their dietary intake of vitamin A. Compliance: women were distributed the capsules in their home on a weekly basis, receipt of capsules was noted only if the distributor observed the woman swallowing the capsule. Over half of the women who became pregnant during the study received over 80% of their intended supplements, and 75% of pregnant women received at least half of their eligible doses. There were serial publications of this study causing the study numerators and denominators to vary between published versions, and multiple pregnancy figures reported did not include higher order pregnancies. Sample-size calculation performed. Partial intention-to-treat analyses, and the risk ratios and confidence intervals were adjusted to account for any cluster-design effect. |
The denominators used for this trial are the number of women randomised who identified themselves as pregnant (i.e. 6070 for the vitamin A group, 5650 for the beta-carotene group and 5653 for the placebo group).

### Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Unclear risk</td>
<td>Cluster-randomised, unclear how sequence was generated.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Unclear risk</td>
<td>Each ward was assigned to the treatment groups based on 'a random draw of numbered chits, blocked on subdistrict'</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>Low risk</td>
<td>Women and investigators blinded to treatment allocation.</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>Unclear risk</td>
<td>“This committee and the data analysts were unmasked to the treatment codes, but the codes were made available to study investigators only at the end of the trial.”</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>Unclear risk</td>
<td>157 women (1%) were lost to follow-up and excluded, partial intention-to-treat analysis performed</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Unclear risk</td>
<td>Denominators vary in several publications of this trial.</td>
</tr>
<tr>
<td>Other bias</td>
<td>High risk</td>
<td>Some women were pregnant more than once during the study period, however the denominators reported are the total number of pregnancies during the study period, not the total number of women randomised, which incorrectly assumes that each data point included is independent from the next</td>
</tr>
</tbody>
</table>
### Methods
Randomisation and allocation concealment: block randomisation, stratified by hospital, using “consecutively numbered, opaque, sealed envelopes”
Blinding of outcome assessment: women and study investigators were initially blinded to the treatment allocation, however the tablet preparations were changed after 55 women were randomised and after this only participants were blinded
Documentation of exclusion: 3 women (1%) were lost to follow-up and excluded
Use of placebo control: 3 treatment regimens were assessed, no placebo control

### Participants
354 women were recruited into the study. Women with a previous NTD defined as anencephalus, iniencephalus, encephalocoele, and spina bifida aperta, who were not pregnant when contacted but were planning a future pregnancy, were eligible and invited to participate. Women were identified from case registers at the participating hospitals. Women with conditions likely to result in impaired absorption from the gastrointestinal tract were excluded.

435 women were approached, of which 354 (84%) consented and were randomised to either F (n = 115), MV (n = 119) or MF (n = 120). 16 women did not become pregnant, and 75 women withdrew; however, their pregnancy outcome status was known, and 18 of these women subsequently became pregnant after withdrawing. 3 women were lost to follow-up. 281 women (93 in the F group, 93 in the MF group and 95 in the MV group) became pregnant in the study period and their pregnancy outcome was known

### Interventions
Indistinguishable trial tablets were initially made by Beecham and Glaxo, however Beecham withdrew their support after 55 women had been randomised. After this time a commercially available pregnavite Forte F was used (MF tablet) and Antigen Pharmaceuticals produced a white multivitamin tablet without folic acid. This was associated with a loss of blinding. Women were randomised to 1 of 3 treatments:
1. folic acid alone (F);
2. multivitamin with folic acid (MF);
3. multivitamin with no folic acid (MV).

The F and MF resulted in a daily dose of 0.3 mg folic acid. The MF and MV resulted in a daily dose of 4000 IU vitamin A, 400 IU calciferol, 1.5 mg thiamine hydrochloride, 1.5 mg riboflavin, 1 mg pyridoxine hydrochloride, 15 mg nicotinamide, 40 mg ascorbic acid, 480 mg calcium phosphate, and 252 mg ferrous sulphate. Women took the tablets for at least 2 months prior to conception and until the date of the 3rd missed period

### Outcomes
1. Recurrence risk of NTDs.
2. Spontaneous abortion.
3. Ectopic pregnancy.
4. Livebirth.
5. Stillbirth.
6. Congenital malformations excluding NTDs.

### Notes
The trial was stopped after there were poor recruitment rates and birth rates. A sample-size calculation required 462 women to show a reduction in NTDs from 5% to 1%. Data from 106 women who were already pregnant at time of recruitment are also included.
The risk profile of women in the trial for spontaneous and recurrent miscarriage is unclear, as is their dietary intake.
Compliance: compliance was assessed on tablet counts and blood tests; however, the results are not presented.
Intention-to-treat analyses were performed. Location: Republic of Ireland. Timeframe: December 1981 to January 1988. The denominators used for this trial are the number of women randomised who became pregnant in the study period and their pregnancy outcome was known (i.e. 93 in the F group, 93 in the MF group and 95 in the MV group)

### Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>Block randomisation stratified by hospital site.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>Consecutively numbered, opaque sealed envelopes used.</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias) All outcomes</td>
<td>Low risk</td>
<td>Only participants were blinded.</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias) All outcomes</td>
<td>High risk</td>
<td>Only participants were blinded.</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias) All outcomes</td>
<td>Low risk</td>
<td>3 women (1%) lost to follow-up and excluded. Intention-to-treat analyses performed</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Unclear risk</td>
<td>Compliance data not reported.</td>
</tr>
<tr>
<td>Other bias</td>
<td>High risk</td>
<td>The trial was stopped after there were poor recruitment rates and birth rates</td>
</tr>
</tbody>
</table>
### Methods

Randomised controlled trial of vitamin A, iron and folic acid supplementation versus iron and folic acid only, during pregnancy, to improve infant outcomes born to women infected with HIV in Malawi.

Randomisation and allocation concealment: "treatment assignment was determined by use of a computer's random-number generator" and "mothers were assigned an original study identification number at enrolment and were given the next sequentially numbered opaque bottle with supplements". "Treatment assignment was concealed by pre packing study supplements in sequentially numbered series assigned to study identification numbers."

Blinding of outcome assessment: unclear, not specifically stated, but participants were blind to their treatment allocation.

Documentation of exclusion: 63 (9%) women were lost to follow-up and 14 (2%) pairs of twins were excluded.

Use of placebo control: control tablets containing iron and folic acid were given.

### Participants

Pregnant women between 18 and 29 weeks' gestation and infected with HIV. The average gestation of participants was 23 weeks. 693 women were enrolled and allocated to either vitamin A (n = 340) or control (n = 357), of which pregnancy outcomes were known for 623 women. 63 women were lost to follow-up and 14 sets of twins were excluded due to their higher risk of low birthweight and infant mortality.

### Interventions

All women received orally administered daily doses of 30 mg iron and 400 mcg folic acid during the study. Women in the intervention group received 10,000 IU vitamin A (3 mg retinol equivalent) orally, in addition to the iron and folic acid supplements. Women were asked to take the tablets from enrolment until delivery. Tablet counts were conducted every 4 weeks. All women received 30 mg retinol equivalents at 6 weeks postpartum, according to standard postpartum care in Malawi.

### Outcomes

1. Infant Hb level at 6 weeks and 12 months of age.
2. Percentage of infants with anaemia at 6 weeks of age and at 12 months, defined as a Hb level of < 110 g/L.
4. Percentage of infants < 2500 g at birth.
5. Weight and length at 6 weeks, 14 weeks and 6 months of age.
6. Transmission of HIV to the infant, infant mortality at < 6 weeks of age, at 12 months and at 24 months.
7. Stillbirth and spontaneous abortion (undefined).

### Notes

Women's risk of spontaneous and recurrent miscarriage is unclear, although may be increased due to their HIV status.

50% of women in the vitamin A group and 51% of women in the control group had deficient levels of vitamin A (defined as plasma vitamin A < 0.70 umol/L) at trial entry. Sample-size calculation performed.

No intention-to-treat analyses were performed.

Compliance: more than 95% of women in both groups took > 90% of study supplements, as ascertained by tablet counts.

Location: Malawi.

Timeframe: November 1995 to December 1996.
### Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>Computer-generated random number list.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>Sequentially number opaque bottles used.</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias) All outcomes</td>
<td>Unclear risk</td>
<td>Not specifically stated but women were blinded.</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias) All outcomes</td>
<td>Unclear risk</td>
<td>Unclear of outcome assessors were unaware of treatment allocation</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>Unclear risk</td>
<td>63 women (9%) lost to follow-up and 14 pairs of twins (2%) excluded. No intention-to-treat analyses performed</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Low risk</td>
<td>All pre-specified outcomes appear to be reported.</td>
</tr>
<tr>
<td>Other bias</td>
<td>Unclear risk</td>
<td>Insufficient information to assess whether an important risk of other bias exists</td>
</tr>
</tbody>
</table>

### McCance 2010

#### Methods

Randomisation and allocation concealment: participants were randomly allocated in a 1:1 ratio to receive 1000 mg vitamin C and 400 IU vitamin E. A randomisation sequence generated in advance by Victoria Pharmaceuticals using PRISYM ID software (version 1.0009) was used. The randomisation sequence was stratified by centre with balanced blocks of 8 patients and was held by Victoria Pharmaceuticals. Individual sealed envelopes containing treatment allocations were given to trial pharmacists in every centre allowing treatment groups to be revealed in a clinical emergency.

Blinding of outcome assessment: diagnosis was independently confirmed by 3 senior clinicians, who were unaware of treatment allocation.

Documentation of exclusion: only 1 loss to follow-up was reported.

Use of placebo control: matched placebo control.

#### Participants

762 pregnant women between 8 and 22 weeks’ gestation with type-1 diabetes attending 25 antenatal metabolic clinics across Northern Ireland, Scotland, and northwest England. Participants were women with type 1 diabetes preceding pregnancy, presentation between 8 weeks’ and 22 weeks’ gestation, singleton pregnancy, and age 16 years or older. Women with chronic hypertension were included in the trial.

Women were excluded if they did not give consent, were enrolled in another research study, were being treated with warfarin, or were known to misuse drugs.

Women taking...
McCance 2010  (Continued)

<table>
<thead>
<tr>
<th>Interventions</th>
<th>1000 mg vitamin C and 400 IU vitamin E versus matched placebo started between 8 and 22 weeks’ gestation and taken until delivery</th>
</tr>
</thead>
</table>
| Outcomes      | 1. Pre-eclampsia.  
2. Placental and endothelial function (established by PAI-1 to PAI-2 ratio).  
3. Gestational hypertension.  
4. Birthweight (centile as calculated from customised birthweight charts).  
5. Miscarriage.  
7. Obstetric complications and other adverse outcomes.  
8. Fetal malformation.  
9. Gestational age at delivery.  
10. Admission to a neonatal care unit. |
| Notes         | Women’s risk profile for spontaneous and recurrent miscarriage: women with chronic hypertension were included.  
Multivitamin supplementation at randomisation was reported at trial entry. Other information about nutrition status are not provided.  
Sample-size calculation: based on 40% reduction in pre-eclampsia.  
Modified intention to treat was used for analysis of the primary endpoint.  
Compliance: unused tablets and capsules were collected during delivery admission or at the 6-week postnatal trial visit, or were returned in postage prepaid envelopes.  
Location: Northern Ireland, Scotland, northwest England.  

**Risk of bias**

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>Computer-generated randomisations by Victoria Pharmaceuticals using PRISYM ID software (version1.0009)</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>Supplements were identical in appearance. The randomisation sequence was stratified by centre with balanced blocks of 8 patients, and was held by Victoria Pharmaceuticals.</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>Low risk</td>
<td>Treatment allocation was masked from all trial personnel and participants until trial completion</td>
</tr>
<tr>
<td>All outcomes</td>
<td></td>
<td>Diagnosis were independently confirmed by 3 senior clinicians, who were unaware of treatment allocation</td>
</tr>
</tbody>
</table>

Vitamin supplementation for preventing miscarriage (Review)
Copyright © 2016 The Cochrane Collaboration. Published by John Wiley & Sons, Ltd.
### McCance 2010 (Continued)

| Incomplete outcome data (attrition bias) | Low risk | Only 1 loss to follow-up in placebo group, but the reason is unclear |
| Selective reporting (reporting bias) | High risk | Fewer outcomes were stated in the trial registration. |
| Other bias | Low risk | The study appears to be free of other sources of bias. |

### MRC 1991

#### Methods

Randomisation and allocation concealment: third party randomisation, “randomisation was carried out through the Clinical Trials Service Unit in Oxford”. Randomisation was stratified by centre

Blinding of outcome assessment: women, caregivers and study investigators were blinded to the treatment allocation

Documentation of exclusion: 164 women (9%) excluded.

Use of placebo control: placebo control.

#### Participants

1817 women were recruited into the study. Women who had a previous pregnancy affected by a NTD, and were planning another pregnancy and not already taking supplements were eligible for the study. Women whose affected child had Meckel’s syndrome and those women with epilepsy were excluded. 1817 women were randomised to either F (n = 449), MV (n = 453), MF (n = 461) or P (n = 454), of which, 1195 were informative pregnancies that is, where the outcome of NTD or not was definitely known (F n = 298, MV n = 302, MF n = 295, P n = 300). Results for pregnancy loss are reported for both informative and not informative pregnancies. 164 women were excluded as they may have been pregnant at the time of randomisation.

#### Interventions

Women were randomised into 1 of 4 groups:

1. 4 mg, 240 mg di-calcium phosphate and 120 mg ferrous sulphate (F);
2. 4000 IU vitamin A, 400 IU calciferol, 1.5 mg thiamine hydrochloride, 1.5 mg riboflavin, 1 mg pyridoxine hydrochloride, 15 mg nicotinamide, 40 mg ascorbic acid, 240 mg di-calcium phosphate and 120 mg ferrous sulphate (MV);
3. folic acid combined with the multivitamins specified above (MF);
4. placebo containing 240 mg di-calcium phosphate and 120 mg ferrous sulphate only (P).

Women took the tablets prior to conception and attended the site every 3 months to collect additional supplies and again during the 12th week of pregnancy. No special dietary advice was given to women.

#### Outcomes

1. NTD and other birth defects.
2. Spontaneous abortions.
3. Ectopic pregnancy.
4. Termination or pregnancy.
5. Livebirth.
7. Multiple pregnancy.
8. Subsequent publications report on blood folic acid and zinc concentrations.

Notes

The trial was stopped early after there were 1195 informative pregnancies, according to prespecified stopping rules. The aim of the study was to obtain information on at least 2000 informative pregnancies unless a sufficiently clear result emerged sooner. Women's risk profile for spontaneous and recurrent miscarriage was unclear, as was their nutritional status.

Compliance: compliance based on self-reports, and data were available for women with an informative pregnancy only, where 79 (6%) women reported they stopped taking their capsules before their last scheduled visit.

Intention-to-treat analyses are reported in this review including not informative pregnancies (i.e. n = 1817).

Location: multi-national study co-ordinated from the UK.


The denominators used for this trial are the number of women randomised, i.e. (449 for the F group, 453 for the MV group, 461 for the MF and 454 for the P group). There was no information provided about any women randomised that did not become pregnant in the study period.

### Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors' judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>Third party randomisation, &quot;randomisation was carried out through the Clinical Trials Service Unit in Oxford&quot;</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>Third party randomisation, &quot;randomisation was carried out through the Clinical Trials Service Unit in Oxford&quot;</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>Low risk</td>
<td>Women, caregivers and investigators blinded to treatment allocation</td>
</tr>
<tr>
<td>All outcomes</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>Unclear risk</td>
<td>No details are given if outcome assessment was blinded.</td>
</tr>
<tr>
<td>All outcomes</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>Low risk</td>
<td>164 women (9%) excluded, intention-to-treat analyses performed</td>
</tr>
<tr>
<td>All outcomes</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Unclear risk</td>
<td>No information provided about any women randomised that did not become pregnant in the study period</td>
</tr>
<tr>
<td>Other bias</td>
<td>High risk</td>
<td>The trial was stopped early after there were 1195 informative pregnancies, according to prespecified stopping rules</td>
</tr>
</tbody>
</table>

---

Vitamin supplementation for preventing miscarriage (Review)

Copyright © 2016 The Cochrane Collaboration. Published by John Wiley & Sons, Ltd.
### Methods

Randomisation and allocation concealment: 1 of the authors 'randomly allocated 1200 participant numbers by computer into 2 groups in permuted blocks of 50'. Every identification number was allocated a supplement container, which was then packed by a team member not otherwise involved in the trial. After enrolment, another author allocated participants sequential identification numbers with the corresponding supplement containers.

Blinding of outcome assessment: double-blind stated but no other details given.

Documentation of exclusion: 61 women (5%) withdrew or were lost to follow-up, however data on miscarriage were reported for those who withdrew due to miscarriage.

Use of placebo control: control of iron and folic acid supplements given which looked identical to the intervention supplements.

### Participants

1200 women were recruited into the study. Women were eligible if they were: less than 20 completed weeks, had a singleton pregnancy, no notable fetal abnormality, no existing maternal illness of a severity that could compromise the outcome of pregnancy, and lived in an area of Dhanusha or the adjoining district of Mahottari accessible for home visits.

Maternal illnesses that led to exclusion were: recently treated recurrent cysticercosis, need for chlorpromazine or anticoagulant drugs with changing doses, and symptomatic mitral stenosis or multivalvular heart disease. Fetal exclusions were: twin pregnancies, anencephaly, occipital meningocele, encephalocele, duodenal atresia and a grossly dilated pelvicalyceal system.

### Interventions

**Intervention group:** vitamin A 800 mcg, vitamin E 10 mg, vitamin D 5 mcg, vitamin B1 1.4 mg, vitamin B2 1.4 mg, niacin 18 mg, vitamin B6 1.9 mg, vitamin B12 2.6 mcg, folic acid 400 mcg, vitamin C 70 mg, iron 30 mg, zinc 15 mg, copper 2 mg, selenium 65 mcg, and iodine 150 mcg.

**Control group:** iron 60 mg and folic acid 400 mcg.

Supplementation began at a minimum of 12 weeks' gestation and continued until delivery.

### Outcomes

1. Birthweight.
2. Gestational duration.
3. Infant length and head circumference.
4. Miscarriage defined as cessation of confirmed pregnancy before 23 weeks' gestation.
5. Stillbirth defined as delivery of an infant showing no signs of life (movement, breathing, or heartbeat) after 23 weeks' gestation.
6. Early neonatal death defined as death of a live born infant in the first 7 days after birth.
7. Late neonatal death as death of a live born infant after 7 but within 28 days.

### Notes

Women's risk of spontaneous and recurrent miscarriage was unclear. Women's nutritional status is also unclear, however, women are presumable at high risk of under-nutrition as the paper states that in Nepal 'deficiencies of several micronutrients have been well described in individual studies and in a national sample'.

Intention-to-treat analyses performed.

Compliance: median 'adherence' was 98% in the control group and 97% in the intervention group.

Location: Nepal.
Osrin 2005  (Continued)

Timeframe: August 2002 to October 2003.

Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors' judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Unclear risk</td>
<td>1 of the authors allocated participants with sequential identification numbers, but unclear if this person was involved in the recruitment of participants</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias) All outcomes</td>
<td>Unclear risk</td>
<td>Double-blind stated in the text but no other details given.</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias) All outcomes</td>
<td>High risk</td>
<td>The allocation code was broken for the analysis. Pg 956.</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias) All outcomes</td>
<td>Low risk</td>
<td>61 women (5%) withdrew or were lost to follow-up, however data on miscarriage were reported for those who withdrew due to miscarriage. Intention-to-treat analyses performed</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Low risk</td>
<td>All pre-specified outcomes appear to be reported.</td>
</tr>
<tr>
<td>Other bias</td>
<td>Low risk</td>
<td>The study appears to be free of other sources of bias.</td>
</tr>
</tbody>
</table>

People’s League 1942

Methods
Randomisation and allocation concealment: “women enrolled at the antenatal clinic were divided into two main groups by placing them alternatively on separate lists”
Blinding of outcome assessment: unclear, no information given on blinding of participants, carers or outcome assessors
Documentation of exclusion: 622 women (11%) were excluded.
Use of placebo control: no placebo given.

Participants
5644 women were recruited into the study. All women attending the antenatal clinics and who were less than or equal to 24 weeks’ gestation and who were in 'good health' were eligible for the study. Women who were more than 24 weeks’ gestation and women who suffered from any disease or physical abnormality were excluded from the study. After enrolment, women who had twin births and who miscarried at an early stage were
also excluded.

5644 women were initially enrolled in the study of which 5022 (89%) remained in the study. Of the 622 (11%) women withdrawn from the trial, 494 were evacuated from the London area (due to World War 2), 39 women had twin births and 89 women miscarried at an early stage. 5022 women remained in the study and were allocated to either multivitamins (n = 2510) or control (n = 2512). Women were further divided into primiparae and multiparae, and various age groups.

### Interventions

Women allocated to the treatment group were given daily vitamin C 100 mg, ferrous iron 0.26 g, calcium 0.26 g, minute quantities of iodine, manganese and copper, adsorbate of vitamin B1 containing all factors of the B complex and halibut liver oil 0.36 g containing vitamin A (52,000 IU per g) and vitamin D (2500 IU per g).

Women allocated to the control group received no placebo.

### Outcomes

1. Toxaemia classified into subgroups based on: hypertension only, albuminuria with or without hypertension, or hypertension with albuminuria (pre-eclampsia).
3. Length of gestation (categorised as less than 40 weeks, 40 weeks, and greater than 40 weeks).
4. Percentage of women breastfeeding.
5. Stillbirth.
6. Neonatal mortality (defined as death before 8 days).
7. Birthweight (pounds) (only reported for primiparae and multiparae separately).

### Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation</td>
<td>High risk</td>
<td>Quasi-randomisation using alternate separate lists.</td>
</tr>
<tr>
<td>Allocation concealment</td>
<td>High risk</td>
<td>No allocation concealment.</td>
</tr>
<tr>
<td>Blinding of participants and personnel</td>
<td>Unclear risk</td>
<td>No information about blinding provided.</td>
</tr>
<tr>
<td>Blinding of outcome assessment</td>
<td>Unclear risk</td>
<td>No information about blinding provided.</td>
</tr>
</tbody>
</table>
### Methods
Randomisation and allocation concealment: the randomisation (computer-generated sequence) was blocked—i.e., balanced—by centre in groups of 2 to 10 individuals.
Blinding of outcome assessment: none of the trial staff or any other person involved in the trial knew the allocated treatment of any woman until after completion of the study.
Documentation of exclusion: 9 (0.4%) women were excluded.
Use of placebo control: placebo control.

### Participants
2404 women with clinical risk factors for pre-eclampsia.
Inclusion criteria:
- gestational age 14+ - 21+ weeks;
- one or more of the following risk factors: pre-eclampsia in the pregnancy preceding the index pregnancy, requiring delivery before 37 completed weeks’ gestation, diagnosis of HELLP syndrome in any previous pregnancy, eclampsia in any previous pregnancy; essential hypertension requiring medication, type 1 or type 2 diabetes, multiple pregnancy; abnormal uterine artery doppler waveform; primiparity with BMI at first antenatal appointment of 30 kg/m² or more.
Exclusion criteria:
- women unable or unwilling to give written informed consent or women who were being treated with warfarin. Women taking vitamin supplements that contained doses of vitamin C of 200 mg or more or of vitamin E of 40 IU or more daily were excluded.

### Interventions
Women were assigned to 1000 mg vitamin C and 400 IU vitamin E (RRR α-tocopherol; n = 1199) or matched placebo (n = 1205) daily from the second trimester of pregnancy until delivery.

### Outcomes
**Primary outcomes**
1. Pre-eclampsia,

**Secondary outcomes**:
1. Low birthweight (< 2.5 kg).
2. Small size for gestational age.
3. Preterm birth (≤ 37+ weeks’ gestation).
4. Gestational age at delivery.
5. Smaller than 10th centile for gestation.
6. Use of health-care resources.

### Notes
Women risk status for spontaneous and recurrent miscarriage: history of chronic hypertension, BMI, pre-eclampsia, multiple pregnancy, diabetes, and other risk factors reported.
Dietary intake at trial entry: use of supplements reported.
**Intention-to-treat analyses:** not performed.

Compliance: 80% (n = 1653) of women took at least 50% of their tablets, 65% (n = 1345) took 80% or more, and 32% (n = 661) took all of their tablets; 6% (n = 125) did not take any tablets.

Sample-size calculation: expected incidence of pre-eclampsia in the placebo group of at least 15% and in the treatment group of at least 30%.

Location: 25 hospitals, UK.


---

### Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>Computer-generated random sequence Pg 1146</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Unclear risk</td>
<td>Although paper states that supplementation and placebo looked and tasted the same, (Pg 1146) there is no clear description of how women were allocated to treatment group</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias) All outcomes</td>
<td>Low risk</td>
<td>“none of the trial staff or any other person involved in the trial knew the allocated treatment of any woman until after completion of the study” Pg 1146</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias) All outcomes</td>
<td>Unclear risk</td>
<td>Unclear from text if outcome assessors were blinded</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias) All outcomes</td>
<td>Low risk</td>
<td>Very low loss to follow-up rates. In the supplementation group 3 (0.25%) losses to follow-up and in the placebo group 6 (0.5%)</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Unclear risk</td>
<td>A large number of outcomes reported in the publication, but not pre-specified in the registered trial</td>
</tr>
<tr>
<td>Other bias</td>
<td>Low risk</td>
<td>The study appears to be free of other sources of bias.</td>
</tr>
</tbody>
</table>
Methods

Randomisation and allocation concealment: community-based, individually-randomised, placebo-controlled and double-blinded study. Pregnant women were randomly allocated in a 1:1:1:1 ratio in blocks of 12 based on a list of treatment numbers derived from a pseudo-random number generated with SAS software.

Blinding of outcome assessment: all investigators, field and laboratory staff and participants were blinded to the treatment code until all field data had been collected and preliminary data analysis by coded groups had been completed.

Documentation of exclusions: 75 women (3.5%) were excluded.

Use of placebo control: placebo control.

Participants

2,173 women at a gestational age of 17 weeks were included in the study. Women at a gestational age of >=17 weeks were not eligible.

Interventions

Women were allocated to one of the three intervention groups:
1. 2,400 retinol equivalents of vitamin A as retinyl palmitate,
2. 20 mg of zinc sulfate, or
3. the same dose of vitamin A and zinc sulfate.

Comparison group: placebo.

All capsules also contained 2 mg dl-alpha-tocopherol as antioxidant and 350 mg of soyabean oil, 20 mg of beeswax and 8 mg of lecithin as capsule filler.

Outcomes

1. Birthweight.
2. Birth length.
3. Neonatal morbidity.
4. Infant mortality.

Notes

Women's risk of spontaneous and recurrent miscarriage was unclear.

Women's nutritional status is unclear.

Intention-to-treat analyses not performed.

Sample-size calculation not performed.

Compliance: consumption of 70% of supplements.

Location: Indonesia.


Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>Pregnant women were randomly allocated in a 1:1:1:1 ratio in blocks of 12 based on a list of treatment numbers derived from a computer-generated pseudo-random number.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Unclear risk</td>
<td>Treatment allocations was prepared and held at the University of Newcastle ....Supplements were coded with treatment numbers and women were assigned a treatment number in sequence based on date they were recruited.</td>
</tr>
</tbody>
</table>
**Prawirohartono 2011**  
(Continued)

<table>
<thead>
<tr>
<th>Risk</th>
<th>Description</th>
<th>Details</th>
</tr>
</thead>
<tbody>
<tr>
<td>Low risk</td>
<td><strong>Blinding of participants and personnel</strong> (performance bias)</td>
<td>All investigators, field and laboratory staff and participants were blinded to the treatment code until all field data had been collected and preliminary data analysis by coded groups had been completed.</td>
</tr>
<tr>
<td>Low risk</td>
<td><strong>Blinding of outcome assessment</strong> (detection bias)</td>
<td>All investigators, field and laboratory staff and participants were blinded to the treatment code until all field data had been collected and preliminary data analysis by coded groups had been completed.</td>
</tr>
<tr>
<td>Low risk</td>
<td><strong>Incomplete outcome data</strong> (attrition bias)</td>
<td>Drop-out rate appears balanced between groups and reasons for loss to follow-up were provided.</td>
</tr>
<tr>
<td>Low risk</td>
<td><strong>Selective reporting</strong> (reporting bias)</td>
<td>Study protocol unavailable but outcomes predefined outcomes were reported.</td>
</tr>
<tr>
<td>Unclear risk</td>
<td><strong>Other bias</strong></td>
<td>Overall, 70% of supplements were consumed, but it is unclear in text where there was more or less compliance.</td>
</tr>
</tbody>
</table>

**Roberfroid 2008**

**Methods**

Randomisation and allocation concealment: the randomisation scheme was generated by a computer program in permuted blocks of 4. Randomisation numbers were sealed in opaque envelopes. At each inclusion, the consulting physician opened the next sealed envelope and transmitted the randomisation number to a pharmacist managing the allocation sequence and the packaging of drugs at a central location. Blinding of outcome assessment: the consulting physicians, pharmacist and women were blinded to allocation. Documentation of exclusions: 107 women were lost to follow-up (however their pregnancy outcome was reported). Post randomisation 26 twins were excluded (multivitamin group: 15; iron/folic acid group: 11 twins (including 1 set of triplets). Only singleton pregnancies were included in the analysis because fetal loss and anthropometric measures at birth in multiple pregnancies are not primarily nutrition-related. 3 women died before delivery and 1 woman underwent a therapeutic abortion. Use of placebo control: no placebo.

**Participants**

1374 women were recruited to participate, however 52 women were randomly assigned twice for consecutive pregnancies, resulting in data for 1426 pregnancies. Women had a pregnancy confirmed by urine testing and were randomly assigned to receive either IFA.
(n = 712) or UNIMMAP (n = 714) daily until 3 months after delivery. Women were recruited between 5 to 36 weeks' gestation; 34.6% (n = 493) of the participants were recruited in the first trimester of pregnancy, mean gestational age at enrolment was 17.3 weeks (SD 7.8 weeks).

**Interventions**

UNIMMAP: vitamin A 800 mcg, vitamin D 200 IU, vitamin E 10 mg, vitamin B-1 1.4 mg, vitamin B-2 1.4 mg, niacin 18 mg, folic acid 400 mcg, vitamin B-6 1.9 mg, vitamin B-12 2.6 mcg, vitamin C 70 mg, zinc 15 mg, iron 30 mg, copper 2 mg, selenium 65 mcg, iodine 150 mcg.

IFA (control): folic acid 400 mcg, Iron 60 mg.

In a case of maternal illness, appropriate treatments were provided according to national guidelines. Severely anaemic women (Hb < 70 g/L, without dyspnoea) received ferrous sulphate (200 mg) + folic acid (0.25 mg) twice daily, for 3 months, regardless of their allocation group. All participants also received 400 mg albendazole in the second and third trimesters. If malaria occurred despite chemoprophylaxis, quinine (300 mg, 3 times/day) was given for 5 days. Vitamin A (200,000 IU) was given to all women after delivery, in accordance with national recommendations.

**Outcomes**

1. Gestational duration.
2. Birthweight, birth length, and Rohrer ponderal index at birth (weight(g)X100/length3(cm)).
3. Low birthweight (< 2500 g).
4. Small-for-gestational age (birthweight below the 10th percentile).
5. Large-for-gestational age (birthweight above the 90th percentile of the study population).
6. Thoracic circumference, head circumference, mid upper arm circumference.
7. Hb concentration in mothers during the third trimester, Hb and sTfR concentrations in cord blood.
8. Preterm birth (< 37 weeks’ gestation).
9. Stillbirth (delivery of an infant showing no sign of life after a gestational age of 28 weeks).

**Notes**

Women's risk of spontaneous and recurrent miscarriage was unclear. 18% of women in each group had experienced a previous fetal loss.

Women's nutritional status is unclear, although women are presumably at risk as the purpose of the trial is to correct MMN deficiencies.

Intention-to-treat analyses not performed, however the review included details of losses to follow-up where the outcome was known.

Compliance: unclear, states that there was no difference in compliance between the 2 groups.

Location: Burkino Faso.

Timeframe: March 2004 to October 2006.

---

**Risk of bias**

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors' judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>
**Roberfroid 2008** (Continued)

<table>
<thead>
<tr>
<th>Bias Type</th>
<th>Risk Level</th>
<th>Details</th>
</tr>
</thead>
<tbody>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>Randomisation numbers were kept in sealed opaque envelopes.</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias) All outcomes</td>
<td>Low risk</td>
<td>Consulting physicians, pharmacist and women were blinded to the intervention</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>Unclear risk</td>
<td>Authors claim the study had double-blind design, but it is unclear if the assessors were blinded</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>High risk</td>
<td>Data were reported for singletons only.</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Unclear risk</td>
<td>As above - data only reported for singletons.</td>
</tr>
<tr>
<td>Other bias</td>
<td>High risk</td>
<td>Some women were pregnant more than once during the study period, however the denominators reported are the total number of pregnancies during the study period, not the total number of women randomised, which incorrectly assumes that each data point included is independent from the next</td>
</tr>
</tbody>
</table>

**Roberts 2010**

<table>
<thead>
<tr>
<th>Methods</th>
<th>Details</th>
</tr>
</thead>
<tbody>
<tr>
<td>Randomisation and allocation concealment: women were randomly assigned to receive capsules containing a combination of 1000 mg of vitamin C (ascorbic acid) and 400 IU of vitamin E (RRR-alpha-tocopherol acetate) or matching placebo (mineral oil). The simple urn method, with stratification according to clinical centre, was used by the data co-ordinating centre to create a randomisation sequence. No further methodological details are provided</td>
<td>Blinding of outcome assessment: medical charts were reviewed by at least 3 reviewers who were unaware of the treatment assignments</td>
</tr>
<tr>
<td>Documentation of exclusions: 183 women (1.8%) were lost to follow-up and for 2 women had been removed after randomisation Use of placebo control: placebo given.</td>
<td></td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Participants</th>
<th>Details</th>
</tr>
</thead>
<tbody>
<tr>
<td>10,154 pregnant women who had a singleton fetus with a gestational age of less than 16 weeks 0 days at the time of screening attending 16 clinical centres and the independent data coordinating centre of the MFMU Network. Women were eligible for inclusion if they had not had a previous pregnancy that lasted beyond 19 weeks 6 days</td>
<td>Women were not eligible if they had elevated systolic or diastolic blood pressure, proteinuria, were taking or had taken antihypertensive medication, or were taking more than</td>
</tr>
</tbody>
</table>
Robert 2010  (Continued)

<table>
<thead>
<tr>
<th>Interventions</th>
<th>A combination of 1000 mg of vitamin C (ascorbic acid) and 400 IU of vitamin E daily administered from enrolment until delivery. The control group received placebo</th>
</tr>
</thead>
</table>

<table>
<thead>
<tr>
<th>Outcomes</th>
<th>1. Pregnancy-associated hypertension. 2. Serious adverse outcomes in the mother or her fetus or neonate. Pre-eclampsia. 3. Other maternal and neonatal outcomes.</th>
</tr>
</thead>
</table>

<table>
<thead>
<tr>
<th>Notes</th>
<th>Women’s risk of spontaneous and recurrent miscarriage was unclear  Use of prenatal vitamins or multivitamins, daily dose of vitamin C and E were reported at trial entry  Sample size calculation was based on a 30% reduction in the rate of the primary outcome  Intention-to-treat analysis was performed.  Compliance: monthly, participants returned unused study drugs from the previous month  Location: UK.  Timeframe: July 2003 to February 2008.</th>
</tr>
</thead>
</table>

<table>
<thead>
<tr>
<th>Risk of bias</th>
<th></th>
</tr>
</thead>
<tbody>
<tr>
<td>Bias</td>
<td>Authors’ judgement</td>
</tr>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Unclear risk</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias) All outcomes</td>
<td>Low risk</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias) All outcomes</td>
<td>Low risk</td>
</tr>
</tbody>
</table>
Incomplete outcome data (attrition bias)

| All outcomes | Low risk | Vitamin C and E = 95/5088 missing and in placebo = 90/5066 missing. Reasons for missing similar between groups |

Selective reporting (reporting bias)

| Low risk | Relevant outcomes reported as pre-specified in the protocol. |

Other bias

| Low risk | The study appears to be free of other sources of bias. |

### Rumbold 2006

#### Methods

Randomisation and allocation concealment: computer-generated random number list with balanced variable blocks and stratification for collaborating centre and gestational age (< 18 weeks versus 18 weeks or more), allocation occurred via a central telephone randomisation service. The treatment packs contained 4 sealed, opaque, white plastic bottles of either the antioxidants vitamin C and vitamin E or the placebo and were prepared by a researcher not involved in recruitment or clinical care.

Blinding of outcome assessment: women, caregivers and investigators were blinded to allocation.

Documentation of exclusion: no losses to follow-up.

Use of placebo control: placebo given.

#### Participants

1877 women were recruited into the study. Eligible women included those: with a nulliparous singleton pregnancy, between 14 and 22 weeks of gestation and with normal blood pressure at the first measurement in pregnancy and again at trial entry.

Women who had any of the following were excluded: known multiple pregnancy, known potentially lethal fetal anomaly, known thrombophilia, chronic renal failure, antihypertensive therapy, or specific contraindications to vitamin C or E therapy such as haemochromatosis or anticoagulant therapy.

Women were allocated to the vitamin C and E group (n = 935) or placebo (n = 935).

#### Interventions

Women allocated to the vitamin C and E group took 4 coated tablets of a combination of 250 mg of vitamin C (as ascorbic acid) and 100 IU of vitamin E (as d-alpha-tocopherol succinate) each day from trial entry until delivery (total daily dose of vitamin C: 1000 mg; vitamin E: 400 IU).

Women were advised not to take any other antioxidant supplements, although a multivitamin preparation that provided a daily intake of no more than 200 mg of vitamin C or 50 IU of vitamin E was permitted.

#### Outcomes

1. Pre-eclampsia.
2. A composite measure of death or serious outcomes in the infant.
4. Serious infant complications occurring before hospital discharge.
5. For women included a composite of any of the following until 6 weeks postpartum: death, pulmonary edema, eclampsia, stroke, thrombocytopenia, renal insufficiency, respiratory distress syndrome, cardiac arrest, respiratory arrest, placental abruption, abnormal liver function, preterm prelabour rupture of membranes, major
<table>
<thead>
<tr>
<th>Risk of bias</th>
<th>Authors' judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>Computer-generated random number list.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>Allocation occurred via a central telephone randomisation service. Tablets were provided in sealed opaque bottles</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>Low risk</td>
<td>Women, caregivers and investigators were blinded.</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>Unclear risk</td>
<td>No information provided about blinding of outcome assessment</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>Low risk</td>
<td>No losses to follow-up.</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Low risk</td>
<td>All pre-specified outcomes reported.</td>
</tr>
<tr>
<td>Other bias</td>
<td>Low risk</td>
<td>The study appears to be free of other sources of bias.</td>
</tr>
</tbody>
</table>

Notes: Women were at low risk of spontaneous and recurrent miscarriage based on the review criteria. The majority of women participating had a baseline dietary intake of vitamin C and E above the Australian recommended daily amount. Intention-to-treat analyses performed. Compliance: there was no difference in compliance between the vitamin group (67%) and the placebo group (70%). Location: Australia. Timeframe: December 2001 and January 2005.
**Methods**

- Generation of random number sequence: a computer-generated random number sequence
- Randomisation and allocation concealment: central allocation (randomisation by an independent third party who had no conflict of interest in the study)
- Blinding of outcome assessment: treatment allocations were blinded to both the investigator and the patient until the study was finished
- Documentation of exclusion: none reported.
- Use of placebo control: no, comparisons were between antioxidants versus iron and folic acid

**Participants**

- 60 women between 8 and 12 weeks’ gestation were eligible for randomisation (supplementation group: n = 29; folic acid group: n = 31)
- Setting: at the antenatal clinic of the Department of Obstetrics and Gynecology, University of Indonesia between March 2003 and June 2004
- Eligibility criteria: pregnant women with low antioxidant status
- Exclusion criteria:
  1. history or current use of anti-hypertensive medication or diuretics;
  2. use of vitamins C > 150 mg and/or E > 75 IU per day;
  3. known placental abnormalities;
  4. current pregnancy as a result of in vitro fertilisation;
  5. regular use of platelet active drugs or non-steroidal anti-inflammatory drugs;
  6. known fetal abnormalities;
  7. documented uterine bleeding within a week of screening;
  8. uterine malformations;
  9. history of medical complications.

**Interventions**

- Supplementation group: received antioxidant supplements daily - vitamins A (1000 IU), B6 (2.2 mg), B12 (2.2 mcg), C (200 mg), E (400 IU), folic acid (400 mcg), N-acetylcysteine (200 mg), Cu (2 mg), Zn (15 mg), Mn (0.5 mg), Fe (30 mg), calcium (800 mg), and selenium (100 mcg)
- Folic acid group: received Fe 30 mg and folic acid 400 mcg daily
- Timing of the intervention: early pregnancy (8 to 12 weeks).

**Outcomes**

- 1. Pre-eclampsia.
- 3. Hypertension.
- 4. Intrauterine growth restriction.
- 5. Intrauterine fetal death.

**Notes**

- Women’s risk of spontaneous and recurrent miscarriage was unclear
- Participating women had low antioxidant status at enrolment, as defined as superoxide dismutase level below 164U/mL. No nutritional information provided
- Intention-to-treat analyses performed.
- Compliance: unclear, no information reported.
- Location: Indonesia.

*Risk of bias*
### Rumiris 2006  (Continued)

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors' judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>Computer-generated random number sequence.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>Central allocation (randomisation by an independent third party who had no conflict of interest in the study)</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>Low risk</td>
<td>Treatment allocations were blinded to both the investigator and the patient until the study was finished</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>Unclear risk</td>
<td>Authors claim the study had a double-blind design, but it is unclear if the assessors were blinded. Treatment allocations were blinded to both the investigator and the patient until the study was finished</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>Low risk</td>
<td>No missing data.</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Low risk</td>
<td>All pre-specified outcomes were reported, no apparent evidence of selective reporting</td>
</tr>
<tr>
<td>Other bias</td>
<td>Unclear risk</td>
<td>At baseline, the control group appears to have a 2 mmHg higher systolic blood pressure than the intervention group, this figure was of borderline statistical significance, P = 0.059</td>
</tr>
</tbody>
</table>

### Rush 1980

#### Methods
Randomisation and allocation concealment: unclear, women were allocated to groups based on "random assignment". Randomisation was stratified on pre-pregnancy weight, weight gain during pregnancy, previous low birthweight infant and protein intake. No other methodological details given
Blinding of outcome assessment: unclear, women were allocated to 2 forms of treatment or control, where both treatments were given as a canned beverage and the control group were given standard oral multivitamins. No information is given on blinding of outcome assessors
Documentation of exclusion: 237 women (22%) were excluded.
Use of placebo control: no placebo, the control group received standard prenatal multivitamin supplements

#### Participants
1051 women were recruited into the study. Women eligible were black, English speaking, and not greater than 30 weeks' gestation. They also had 1 of the following criteria: low pre-pregnant weight (under 110 pounds at conception); low weight gain at the time of
Rush 1980 (Continued)

Recruitment; at least 1 previous low birthweight infant; a history of protein intake of less than 50 g in the 24 hours preceding recruitment. Women were not eligible if they were known to be seeking a termination, had specific chronic health disorders, if they admitted to recent use of narcotics or heavy use of alcohol, or weighed >= 140 pounds at conception.

The mean gestation at enrolment ranged from 16-18 weeks for the treatment groups. 1225 women were invited to join the study, of which 1051 (84%) consented. Of these, 237 (22%) were excluded and 814 women (77%) remained active in the study until delivery and were allocated to 1 of 3 groups: supplement (n = 263), complement (n = 272) or control (n = 279).

Interventions

Women were randomised to 1 of 3 groups:
1. high protein supplement (daily 40 g animal protein, 470 calories, 1000 mg calcium, 100 mg magnesium, 60 mg iron, 4 mg zinc, 2 mg copper, 150 mcg iodine, 6000 IU vitamin A, 400 IU vitamin D, 30 USPU vitamin E, 60 mg vitamin C, 3 mg vitamin B1, 15 mg vitamin B2, 15 mg niacin, 2.5 mg vitamin B6, 1 mg pantothenic acid, 200 mcg biotin, 350 mcg folic acid, 8 mcg vitamin B12);
2. balanced protein-energy complement (6 g animal protein, 250 mg calcium, 12 mg magnesium, 40 mg iron, 0.084 mg zinc, 0.15 mg copper, 100 mcg iodine, 4000 IU vitamin A, 400 IU vitamin D, 60 mg vitamin C, 3 mg vitamin B1, 15 mg vitamin B2, 10 mg niacin, 3 mg vitamin B6, 1 mg pantothenic acid, 350 mcg folic acid, 3 mcg vitamin B12);
3. control (250 mg calcium, 0.15 mg magnesium, 117 mg iron, 0.85 mg zinc, 0.15 mg copper, 100 mcg iodine, 4000 IU vitamin A, 400 IU vitamin D, 60 mg vitamin C, 3 mg vitamin B1, 2 mg vitamin B2, 10 mg niacin, 3 mg vitamin B6, 1 mg pantothenic acid, 350 mcg folic acid, 3 mcg vitamin B12).

Women received the high protein or balanced protein-energy supplements in the format of a drink. Women in the control group received a standard oral prenatal multivitamin supplement.

Outcomes

1. Total weight gain, average weight gain and early weight gain during pregnancy.
2. Duration of gestation (presented as cumulative rates of delivery from life tables for each treatment group).
3. Preterm birth < 37 weeks.
4. Fetal death (before < 20 weeks’ gestation and >= 20 weeks’ gestation).
5. Neonatal death (according to gestation at delivery).
7. Somatic measures of infant growth at 1 year of age.
8. Psychological measures at 1 year of age.

Notes

Women’s risk of spontaneous and recurrent miscarriage is unclear, as there is no information about concurrent medical conditions or other risk factors for miscarriage. Women in the trial had a low caloric intake at trial entry, and unexpectedly, an adequate protein intake. No other specific nutritional information is reported.
Sample-size calculation reported: 250 women were required in each treatment group to show a 125 g difference in birthweight. A 25% loss to follow-up was incorporated into the sample size.
Intention-to-treat analyses not performed.
There were 9 sets of twins amongst the 3 treatment groups.
Compliance: "on average, about three quarters of the prescribed amount of beverage was probably ingested".
Location: New York City, USA.

<table>
<thead>
<tr>
<th>Risk of bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Unclear risk</td>
<td>No methodological details given beyond reporting of 'random assignment'</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Unclear risk</td>
<td>No methodological details given beyond reporting of 'random assignment'</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>Unclear risk</td>
<td>No information provided about blinding of participants and personnel. Unlikely as participants were given canned beverages or multivitamins</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>Unclear risk</td>
<td>No information provided about blinding of outcome assessment</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>Unclear risk</td>
<td>237 women (22%) excluded, no intention-to-treat analysis.</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Unclear risk</td>
<td>Unclear if all pre-specified outcomes reported.</td>
</tr>
<tr>
<td>Other bias</td>
<td>Unclear risk</td>
<td>Limited methodological details provided.</td>
</tr>
</tbody>
</table>

**Schmidt 2001**

Methods
Randomisation and allocation concealment: unclear, women were “randomly assigned on an individual basis, to double-blind, weekly supplementation until delivery”
Blinding of outcome assessment: unclear, double-blind stated in text but no details given
Documentation of exclusion: 42 women (17%) were lost to follow-up and excluded
Use of placebo control: control tablets containing iron and folic acid were given

Participants
243 women were recruited into this study. Pregnant women between 16 and 20 weeks' gestation, aged between 17 and 35 years old, with a parity < 6 and Hb level between 80-140 g/L, were eligible for this study. Women were randomised to receive either vitamin A plus iron and folic acid (n = 122) or iron and folic acid only (n = 121). Of these 22 (18%) and 20 (17%) women in vitamin A plus iron and folic acid and the iron and folic acid groups respectively, dropped out between enrolment and the follow-up at 4 months.
Interventions

Women were randomised to a weekly supplementation with 120 mg ferrous sulfate and 500 mcg folic acid, with or without vitamin A (2400 retinol equivalents). Women were asked to take the trial tablets from between 16 and 20 weeks’ gestation until birth.

Outcomes

1. Stillbirth.
2. Concentrations of Hb, serum ferritin and serum transferrin receptors, at or near term.
3. Concentrations of iron and vitamin A in breast milk.
4. Hb and serum vitamin A concentrations in the mother and infant at 4 months postpartum.
5. General health, growth and development measures in the first year of life.

Notes

Women risk status for spontaneous and recurrent miscarriage is unclear.
At baseline, between 13% and 17% of women had marginal vitamin A deficiency 44% to 50% of women were anaemic.
Sample-size calculation performed allowing for a 50% drop-out during the study period.
Intention-to-treat analyses were not performed.
Compliance: adherence to the tablet intake was assessed through interview during a postnatal home visit, which revealed that the median tablet intake was 50 tablets (i.e. 25 weeks), while only 17% of the subjects took more than 90 tablets.
Location: Indonesia.
Serial publications of this data report different denominators.

Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Unclear risk</td>
<td>No information provided about sequence generation.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Unclear risk</td>
<td>Women were 'randomly assigned on an individual basis' but no other details given</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>Unclear risk</td>
<td>Double-blind used in the text but no details provided.</td>
</tr>
<tr>
<td>All outcomes</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>Unclear risk</td>
<td>Authors claim the study had double-blind design, but it is unclear who were blinded. No further information was available</td>
</tr>
<tr>
<td>All outcomes</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>Unclear risk</td>
<td>42 women (17%) were lost to follow-up and excluded, no intention-to-treat analyses performed</td>
</tr>
<tr>
<td>All outcomes</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Unclear risk</td>
<td>Serial publications of this study report different denominators</td>
</tr>
</tbody>
</table>
Other bias | Unclear risk | Limited methodological details provided.
---|---|---

### Spinnato 2007

**Methods**
- Generation of random number sequence: the randomisation sequence was constructed by the data co-ordinating centre (DCC) as permuted blocks of random size, stratified by clinical centre, and implemented via a program residing on the clinical centres study computer.
- Randomisation and allocation concealment: central allocation.
- Blinding of outcome assessment: all clinicians and clinical investigators were blinded to group assignment.
- Documentation of exclusion: none reported.
- Use of placebo control: placebo control.

**Participants**
- 739 eligible women between 12\(\frac{0}{7}\) and 19\(\frac{6}{7}\) weeks of gestation were enrolled in the study (treatment: 371; placebo: 368).
- Setting: 4 Brazilian clinical centres: 1 primary clinical centre (Recife) and 3 additional clinical sites (Campinas, Botucatu, and Porto Alegre); each site’s major teaching hospital serves a primarily urban low-income population.
- Eligibility criteria: women between 12\(\frac{0}{7}\) and 19\(\frac{6}{7}\) weeks of gestation and diagnosed with nonproteinuric chronic hypertension or a prior history of pre-eclampsia in their most recent pregnancy that progressed beyond 20 weeks’ gestation.
- Exclusion criteria: multifetal gestation, allergy to vitamin C or vitamin E, requirement for aspirin or anticoagulant medication, 24-hour urinary protein ≥ 300 mg, pre-pregnancy diabetes mellitus, known fetal anomaly incompatible with life.
- Loss to follow-up: 32 women (treatment 16; placebo 16).

**Interventions**
- Intervention group: daily treatment with both vitamin C (1000 mg) and E (400 IU) until delivery or until the diagnosis of pre-eclampsia.
- Control group: daily placebo until delivery or until the diagnosis of pre-eclampsia.
- Timing of the intervention: between 12\(\frac{0}{7}\) and 19\(\frac{6}{7}\) weeks of gestation.

**Outcomes**
- 1. Pre-eclampsia (women were followed through the 14th day postpartum for the occurrence of pre-eclampsia).
- 2. Severity of pre-eclampsia.
- 4. Abruptio placentae.
- 5. Premature rupture of membranes.
- 6. Preterm birth.
- 7. Small-for-gestational age.
- 8. Low birthweight infant.

**Notes**
- 25 inclusion/exclusion criteria violations (23 enrolled outside 12-19 weeks’ gestation; 2 twin gestations - 1 lost to spontaneous abortions, 1 delivered liveborn in treatment group); all 25 women remained in their assigned study groups.
- 26 women had early treatment terminations (treatment 19; placebo 7), but remained in follow-up.
- Women’s risk of spontaneous and recurrent miscarriage was unclear.
Women’s nutritional status is also unclear. Intention-to-treat analyses performed. Compliance: average compliance was 85%, and similar between treatment groups.

Location: Brazil.

<table>
<thead>
<tr>
<th>Risk of bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>Computer-generated sequence number.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>Central allocation.</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias) All outcomes</td>
<td>Low risk</td>
<td>All clinicians and clinical investigators were blinded to group assignment</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias) All outcomes</td>
<td>Unclear risk</td>
<td>Authors claim the study had double-blind design, but it is unclear who was blinded. No further information was available</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias) All outcomes</td>
<td>Low risk</td>
<td>Small numbers of missing data, balanced across groups (32 women; treatment 16; placebo 16)</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Low risk</td>
<td>All pre-specified outcomes were reported, no apparent evidence of selective reporting</td>
</tr>
<tr>
<td>Other bias</td>
<td>Low risk</td>
<td>The study appears to be free of other sources of bias.</td>
</tr>
</tbody>
</table>

Steyn 2003

Methods

Randomisation and allocation concealment: “randomisation was undertaken by computer-generated numbers”. Roche Pharmaceutical supplied numbered containers with either vitamin C or matching placebo, and they retained the study code until completion of the study. No other methodological details given.

Blinding of outcome assessment: “double blind” stated, Roche Pharmaceuticals retained the code until completion of the study.

Documentation of exclusion: none reported.

Use of placebo control: placebo control.

Participants

200 women were recruited into the study. Women with a history of a previous mid-trimester abortion (defined as spontaneous expulsion of the uterine contents between 13 and 26 weeks’ gestational age), or previous preterm labour, and less than 26 weeks’ gestation were eligible and invited to participate. Women with iatrogenic causes of their
previous preterm labour such as previous induction of labour before term for severe pre-eclampsia, were excluded. 203 consecutive women were approached, of which 200 (98.5%) consented and were randomised to either vitamin C (n = 100) or placebo (n = 100). No losses to follow-up were reported.

Interventions
Twice daily tablet of either 250 mg vitamin C or placebo, from trial entry until 34 weeks' gestation. All women were tested for bacterial vaginosis and all women with positive cultures for *Mycoplasma hominis* (and between 22 and 32 weeks' gestation) were treated with erythromycin for 7 days.

Outcomes
1. Preterm labour, defined as spontaneous onset of labour and delivery before 37 completed weeks.
2. The secondary outcome was perinatal outcome, a composite endpoint including birthweight, gestational age at delivery, perinatal mortality, duration of admission in the neonatal intensive care unit and neonatal complications. The age of fetal viability was considered to be 28 weeks' gestation.

Notes
Results are from an interim analysis performed when 100 participants were recruited into each arm. Recruitment was stopped after the interim analysis revealed few differences between the 2 groups. Unclear if there was a sample-size calculation performed. Women's risk profile spontaneous and recurrent miscarriage is unclear, although they are clearly at high risk of preterm birth. It is also unclear if multiple births were included. 6% of women had an inadequate dietary intake of vitamin C, defined as an intake < 67% of the recommended dietary allowance (70 mg per day).

 Compliance: women were requested to bring the containers to each visit and the remaining tablets were counted to improve and control compliance; however, no compliance data were reported.

Country: South Africa.
Timeframe: unclear.

### Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors' judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>Computer-generated sequence number.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>Roche pharmaceuticals supplied numbered study containers and kept the study code until completion of the study</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>Unclear risk</td>
<td>Authors claim the study had double-blind design, but it is unclear who was blinded. Allocation was double-blind and Roche Pharmaceuticals retained the code until completion of the study</td>
</tr>
</tbody>
</table>
### Steyn 2003

<table>
<thead>
<tr>
<th>Bias Type</th>
<th>Risk</th>
<th>Note</th>
</tr>
</thead>
<tbody>
<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>Unclear risk</td>
<td>Authors claim the study had double-blind design, but it is unclear who was blinded. Allocation was double-blind and Roche Pharmaceuticals retained the code until completion of the study.</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>Low risk</td>
<td>No losses to follow-up reported.</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Unclear risk</td>
<td>Recruitment stopped after an interim analysis.</td>
</tr>
<tr>
<td>Other bias</td>
<td>High risk</td>
<td>Results are from an interim analysis performed when 100 participants were recruited into each arm.</td>
</tr>
</tbody>
</table>

### Summit 2008

#### Methods
- Generation of random number sequence: computer-generated number; 262 clustered unit of randomisation (all pregnant women served by the same midwife received supplements with the same midwife identification number)
- Randomisation and allocation concealment: central allocation
- Blinding of outcome assessment: all study scientists and personnel, government staff, and enrollees were unaware of the allocation
- Documentation of exclusion: 1748 loss to follow-up before delivery (IFA: 853; MMN: 895); 1128 loss to follow-up after delivery (IFA: 553; MMN: 575). 10,549 pregnant women excluded post-randomisation because of trial termination (IFA group: 5057; MMN group: 5492)
- Use of placebo control: no placebo, comparisons were between multiple micronutrients and iron and folic acid

#### Participants
- 41,839 pregnant women of any gestational age living on Lombok, Nusa Tenggara Barat Province, Indonesia. Women were allocated to IFA (n = 20,543) or MMN (n = 21,296)

#### Interventions
- MMN group: the MMN was the UNIMMAP formulation containing 30 mg iron (ferrous fumarate) and 400 mcg folic acid along with 800 mcg retinol (retinyl acetate), 200 IU vitamin D (ergocalciferol), 10 mg vitamin E (alpha tocopherol acetate), 70 mg ascorbic acid, 1.4 mg vitamin B1 (thiamine mononitrate), 18 mg niacin (niacinamide), 1.9 mg vitamin B6 (pyridoxine), 2.6 mcg vitamin B12 (cyanocobalamin), 15 mg zinc (zinc gluconate), 2 mg copper, 65 mcg selenium, and 150 mcg iodine - 1 capsule daily up to 3 months after birth
- IFA group: the IFA contained 30 mg iron (ferrous fumarate) and 400 mcg folic acid - 1 capsule daily up to 3 months after birth
- Timing of the intervention: any time during pregnancy.

#### Outcomes
- 1. Early infant mortality (deaths until 90 days postpartum).
- 3. Fetal loss (abortions and stillbirths).
4. Low birthweight.

Notes

Women’s risk of spontaneous and recurrent miscarriage was unclear. Women’s nutritional status is also unclear. However, 30% of women in each group had an mid upper arm circumference < 23.5cm, which was used as an indicator of women being undernourished. Intention-to-treat analyses performed.

Compliance: median compliance was 85%, there was no difference between treatment groups in compliance.

Location: Indonesia.

Timeframe: July 1, 2001 to April 1, 2004.

---

Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>Computer-generated number.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>Central allocation.</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>Low risk</td>
<td>All study scientists and personnel, government staff, and enrollees were unaware of the allocation</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>Low risk</td>
<td>The code to indicate which strip was IFA or MMN was known only by the manufacturing production manager and a quality control officer from UNICEF. All study scientists and personnel, government staff, were unaware of the allocation. Blinding is unlikely to be broken</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>High risk</td>
<td>Loss to follow-up: 1748 loss to follow-up before delivery (IFA: 853; MMN: 895); 1128 loss to follow-up after delivery (IFA: 553; MMN: 575) Post-randomisation exclusion: 10, 549 pregnant women excluded because of trial termination (IFA group: 5057; MMN group: 5492)</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Low risk</td>
<td>All pre-specified outcomes were reported, no apparent evidence of selective reporting</td>
</tr>
<tr>
<td>Other bias</td>
<td>Low risk</td>
<td>The study appears to be free of other sources of bias.</td>
</tr>
</tbody>
</table>
Sunawang 2009

**Methods**

Randomisation and allocation concealment: 160 clusters were randomly assigned to 4 blocks of 40 clusters each. 2 of the 4 blocks received either MMN of IFA. No details of how blocks were selected. 432 women in 40 clusters were allocated to the group receiving multiple micronutrients and 411 in the other 40 clusters were allocated to the iron-folic acid group. No further details on how allocation was done.

Blinding of outcome assessment: single-blind design, no further details provided.

Documentation of exclusion: 93 women (11%) were excluded.

Use of placebo control: no placebo, the control group received iron-folic acid supplementation.

**Participants**

843 women residing in Indramayu District in the West Java Province of Indonesia who were between 12 to 20 weeks of gestation. Only women intending to remain in the study location until giving birth were recruited. Women suffering from potentially confounding illnesses, including diabetes, coronary heart disease, and tuberculosis, were excluded from the study.

**Interventions**

MMN supplements containing 15 micronutrients vs IFA (60 mg of elemental iron as ferrous sulfate and 0.25 mg of folic acid). Supplementation was from time of enrolment at 12 to 20 weeks of gestation and continued to 30 days postpartum.

**Outcomes**

1. Birthweight.
2. Anthropometry.
3. Hb, serum ferritin, serum zinc, and serum retinol.
5. Side effects of supplementation.
7. Fetal loss.

**Notes**

Women’s risk of spontaneous and recurrent miscarriage is unclear.

Women’s nutritional status is also unclear.

Sample size calculation not done.

Intention-to-treat analysis performed.

Compliance: Data on compliance was collected from weekly visits. Mean adherence in the intervention group 68% and control group 71%.

Location: Indonesia.


**Risk of bias**

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Unclear risk</td>
<td>160 clusters were randomly assigned to 4 blocks of 40 clusters each. 2 of the 4 blocks received either MMN of IFA. No details of how blocks were selected. pg s489 pg 5</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Unclear risk</td>
<td>432 women in 40 clusters were allocated to the group receiving multiple micronutrients and 411 in the other 40 clusters were</td>
</tr>
<tr>
<td></td>
<td>Risk</td>
<td>Notes</td>
</tr>
<tr>
<td>--------------------------------</td>
<td>-------</td>
<td>----------------------------------------------------------------------</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>High risk</td>
<td>Authors claim the study had a single-blind design, but it is unclear who was blinded. Probably not done. pg s489 pgh 8</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>High risk</td>
<td>No details of blinded outcome assessment provided.</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>Unclear risk</td>
<td>23/432 missing in MMN and 33/411 missing in IFA. Reasons for drop-out are similar, but IFA had 5 twin births while MMN cluster had none. Fig. 1</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Low risk</td>
<td>All pre-specified outcomes were reported as described.</td>
</tr>
<tr>
<td>Other bias</td>
<td>Low risk</td>
<td>The study appears to be free of other sources of bias.</td>
</tr>
</tbody>
</table>

**Tofail 2008**

**Methods**

Generation of random number sequence: a computer-generated register was used for randomisation. Randomisation was performed in blocks of 12 food supplementation was randomly allocated but not blinded. micronutrient capsules looked identical and women were allocated to 1 of 3 types of micronutrient supplements in a 2 by 3 design Blinding of outcome assessment: allocation code was not broken until after the analysis Documentation of exclusion: 845 losses to follow-up (19%). Use of placebo control: no placebo, the control group received IFA (Fe30Fol) supplementation

**Participants**

4436 pregnant women from rural Bangladesh within 6-8 weeks of conception (not more than 14 weeks). Women with Hb < 80 g/L were ineligible for participation

**Interventions**

MMNs containing 15 recommended micronutrients (30 mg of iron, 400 µg of folic acid). The micronutrient supplements were offered to the enrolled pregnant women at the 14weeks' clinic visit up to 3 months after delivery. Intervention was divided into early or late food supplementation Early or usual food supplementation plus 1, 2 or 3:
1) 30 mg iron(fumarate) + 400 mcg folic acid (Fe30Fol)
2) 60 mg iron (fumarate) + 400 mcg folic acid (Fe60Fol)
3) 30 mg iron (fumarate), 400 mcg folic acid, 800 mcg RE vitamin A (retinyl acetate), 5 mcg vitamin D (cholecalciferol), 10 mg vitamin E (a-tocopherol acetate), 70 mg vitamin C, 1.4 mg thiamine (mononitrate), 1.4 mg riboflavin, 18 mg niacin, 1.9 mg vitamin B-6 (pyridoxine hydrochloride), 2.6 mcg vitamin B-12 (cyanocobalamin), 15 mg zinc (sulfate), 2 mg copper (sulfate), 65 mcg selenium (sodium selenite), and 150 mcg iodine
(potassium iodide) (MMNs)
All women received food supplement which consisted of roasted rice powder, roasted
pulse powder, molasses, and soybean oil and had a total energy content of 2500 kJ.
Early (immediately after detection of pregnancy, around 9 weeks) enrolment in food
supplementation or usual (at the time of their choosing, around 20 weeks)

### Outcomes

**Infant outcomes:**
1. Blood Hb and plasma ferritin, zinc, retinol, vitamin B-12, and folate at 6 months.
2. Low birthweight.
4. Morbidities (diarrhoea, lower acute respiratory infections).
5. Birth anthropometry.
8. Kidney function

**Maternal outcome:**
1. Hb at 30 weeks.

### Notes

Women’s risk of spontaneous and recurrent miscarriage was unclear
Women’s nutritional status is also unclear.
Sample size calculation: not done.
Intention-to-treat analysis performed.
Compliance: the number of micronutrient pills taken in the first 30 weeks of pregnancy was 79+/−34 in the Fe30 group, 78+/−34 in the Fe60 group, and 75+/−33 in the MM group
Location: Bangladesh.

### Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>A computerised tracking system assigned women to 1 of the 6 groups in blocks of 12 women were randomly allocated. “Computer generated register of study identity numbers with random assignment of food groups (“E” or “U”) and micronutrient groups (from 12 possible pill bottle number codes) was used for randomization.” Randomisation was probably done. Eneroth 2011 pg 221</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>Micronutrient capsules looked identical.</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias) All outcomes</td>
<td>Low risk</td>
<td>Each micronutrient group had been given 4 different number codes to decrease the risk of unblinding. Pg 2054</td>
</tr>
</tbody>
</table>
### Blinding of outcome assessment (detection bias)

| All outcomes | Low risk | Randomisation codes were safely kept at the administrative office of ICDDR,B and were not broken until after performing the intention-to-treat analyses. Pg 2054 |

### Incomplete outcome data (attrition bias)

| All outcomes | Low risk | 845 of 4436 losses to follow-up (19%) before birth; reasons were available. Numbers of women lost to follow-up did not differ among the treatment groups |

### Selective reporting (reporting bias)

| Low risk | All the study’s pre-specified outcomes have been reported. |

### Other bias

| Low risk | The study appears to be free of other sources of bias. |

### Van den Broek 2006

#### Methods

- Generation of random number sequence: using a random-generation procedure
- Randomisation and allocation concealment: the supplements in vitamin A and placebo treatments allocated were prepared in identical capsules and were packaged in bottles according to the randomisation schedule (sealed envelopes) by midwives who were not involved in the trial conduct
- Blinding of outcome assessment: neither the women nor the midwives involved in treatment allocation revealed the randomisation schedule to anyone involved in the conduct of the trial
- Documentation of exclusion: 77 loss to follow-up before assessment at 26-28 weeks (5000 IU vitamin A: 26; 10,000 IU vitamin A: 26; placebo: 25). Additional 93 loss to follow-up before assessment at 36-38 weeks (5000 IU vitamin A: 34; 10,000 IU vitamin A: 28; placebo: 31)
- Use of placebo control: placebo control.

#### Participants

- 700 women with singleton pregnancies at 12-24 weeks' gestation measured by ultrasound scan (5000 IU vitamin A: 234; 10,000 IU vitamin A: 234; placebo: 232)
- Setting: the antenatal clinic at the Namitambo rural Health Centre in southern Malawi, central Africa
- Eligibility criteria: (Hb) < 11.0 g/dL by HemoCue screening method at first antenatal visit, singleton pregnancy with gestational age > 12 weeks and ≤ 24 weeks measured by ultrasound scan, no fetal abnormality detectable by ultrasound at time of booking, residing in the catchment area of the health centre
- Exclusion criteria: women > 24 weeks' gestation, or twin pregnancy

#### Interventions

- Intervention group 1: 5000 IU vitamin A daily until delivery.
- Intervention group 2: 10,000 IU vitamin A daily until delivery.
- Comparison group: placebo daily until delivery.
- Timing of the intervention: supplementation started as early as possible after 12 weeks of pregnancy
- All women received iron tablets daily (60 mg elemental iron as ferrous sulphate with 0.5mg of vitamin C).
Outcomes

1. Anaemia status (no anaemia ([Hb] ≥ 11.0 g/dL), anaemia ([Hb] < 11.0 g/dL) or severe anaemia ([Hb] < 8.0 g/dL).
2. Hb concentration (Coulter counter value), iron status (determined by serum ferritin and serum transferring receptor concentration).
3. Evidence of infection (assessed by serum CRP, peripheral malaria parasitaemia and HIV status).
4. Vitamin A status (determined by serum retinol and the MRDR).

Notes

Women’s risk of spontaneous and recurrent miscarriage was unclear.
Women’s nutritional status is also unclear.
Intention-to-treat analyses performed.
Compliance: unclear, no information provided.
Location: Malawi.

Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>Random-generation procedure used.</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>The vitamin A and placebo treatments allocated were prepared in identical capsules and packaged in bottles according to the randomisation schedule (sealed envelopes) by midwives who were not involved in the trial conduct.</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>Low risk</td>
<td>Neither the women nor the midwives involved in treatment allocation revealed the randomisation schedule to anyone involved in the conduct of the trial.</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>Unclear risk</td>
<td>No information provided about blinding of outcome assessment.</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias) All outcomes</td>
<td>High risk</td>
<td>77 loss to follow-up before assessment at 26-28 weeks (5000 IU vitamin A: 26; 10,000 IU vitamin A: 28; placebo: 25). Additional 93 loss to follow-up before assessment at 36-38 weeks (5000 IU vitamin A: 34; 10,000 IU vitamin A: 28; placebo: 31).</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Low risk</td>
<td>All pre-specified outcomes were reported, no apparent evidence of selective reporting.</td>
</tr>
</tbody>
</table>
Other bias | Low risk | The study appears to be free of other sources of bias.

### Villar 2009

**Methods**
- Generation of random number sequence: no sequence generation details available
- Randomisation and allocation concealment: central allocation (randomisation was performed by the statisticians of the British VIP Trial)
- Blinding of outcome assessment: "double blind" stated.
- Documentation of exclusion: 10 women (treatment 6; placebo 4), and 29 infants (treatment 13, placebo 16) were lost to follow-up
- Use of placebo control: placebo control.

**Participants**
- 1365 women between 14-22 gestational age agreed to participate and were randomised (vitamins group: 687; placebo group: 678)
- Setting: antenatal clinics located in Nagpur, India; Lima and Trujillo, Peru; Cape Town, South Africa; and Ho Chi Minh City, Viet Nam which served populations with low social-economic status and had evidence of overall low nutritional status, between October 2004 and December 2006
- Eligibility criteria: pregnant women considered high risk for pre-eclampsia (chronic hypertension, renal disease, pre-eclampsia-eclampsia in the pregnancy preceding the index pregnancy requiring delivery before 37 weeks' gestation, HELLP syndrome in any previous pregnancy, pre-gestational diabetes, primiparous with a BMI > 30 kg/m², history of medically-induced preterm delivery, abnormal uterine artery Doppler waveforms and women with antiphospholipid syndrome), multifetal gestation. Women ingesting medications with aspirin-like compounds were not excluded
- Exclusion criteria: women ingesting vitamin supplements that contained ≥ 200 mg of vitamin C and/or ≥ 50 IU of vitamin E and women receiving warfarin

**Interventions**
- Intervention group: received 1000 mg vitamin C and 400 IU of vitamin E daily until delivery
- Comparison group: received placebo daily until delivery.
- Timing of the intervention: between 14 and 22 weeks' gestation

**Outcomes**
- 1. Pre-eclampsia.
- 2. Eclampsia.
- 3. Placental abruption.
- 4. Low birthweight (< 2500 g).
- 5. Small-for-gestational age (< 10th centile of the WHO recommended standard).
- 6. Intrauterine or neonatal death before hospital discharge.
- 7. Preterm delivery (< 37 weeks).
- 8. Early preterm delivery (< 34 weeks).
- 9. Very low birthweight (< 1500 g).
- 10. ≥ 7 days in the neonatal intensive care unit.
- Pre-eclampsia information was unavailable for 14 women in the vitamins and 9 in the placebo group.
- There were data from 81 supplemented (11.8%) and 100 placebo-treated (14.7%)
women with multiple pregnancies, for whom newborn outcomes were considered separately

Notes

Women’s risk of spontaneous and recurrent miscarriage was unclear. Women at high risk of pre-eclampsia were included but data on fetal loss was not reported separately for this group.

No specific information on women’s nutritional status is included; however, the paper states that the trial was conducted in populations with ‘documented low nutritional status’

Intention-to-treat analyses performed.

Compliance: median compliance was 87%, and was similar between the treatment groups.

Location: Antenatal clinics in India, Peru, South Africa and Vietnam.

Timeframe: October 2004 and December 2006.

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>Randomisation sequence blocked by centre in groups of 2 to 10 individuals</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>Central allocation (randomisation was performed by the statisticians of the British VIP Trial)</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias) All outcomes</td>
<td>Low risk</td>
<td>Women and investigators blinded to allocation.</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias) All outcomes</td>
<td>Low risk</td>
<td>Identifying characteristics and randomisation code were kept in a secure and separate database in the server unavailable to the research team, until all data analyses were completed</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias) All outcomes</td>
<td>Low risk</td>
<td>Small numbers of missing data, balanced across groups. Women: 10 (treatment 6; placebo 4). Infants: 29 (treatment 13; placebo 16).</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Unclear risk</td>
<td>Perinatal death was reported instead of prespecified neonatal death</td>
</tr>
<tr>
<td>Other bias</td>
<td>Low risk</td>
<td>The study appears to be free of other sources of bias.</td>
</tr>
</tbody>
</table>
### Methods

Generation of random number sequence: sectors were used as the unit of randomisations for supplement allocation. 3 sets of 3 identical coins on which the numbers 1, 2 or 3 were written were placed into a container, mixed and removed randomly, without replacement, and the 3-digit code of each sector was read aloud sequentially. Study supplements were identical in colour, taste and external appearance. Treatment codes were assigned by the Nutrilite Health Institute and were kept in a sealed envelope in a locked cabinet at both Nutrilite and Johns Hopkins University. Masking was further ensured by having a senior administrative staff (not involved in field activities) recode the supplement bottles with sector-specific permanent stickers bearing codes from 001 to 596.

Blinding of outcome assessment: investigators were all blinded to the allocation codes until the end of the parent trial.

Documentation of exclusion: 628 women (1%) were excluded.

Use of placebo control: placebo control.

### Participants

102,769 women from 596 sectors (60,294 pregnancy identified and 59,666 pregnancies included) 19 rural unions in the northwest Districts of Gaibandha and Rangpur in rural Bangladesh. Women were recruited from the first trimester of pregnancy through 12 weeks (84th day) after pregnancy termination. Eligible women included pregnant or postpartum, lactationally amenorrhoeic women were placed on a ‘waiting list’, and only became eligible for pregnancy surveillance once their menses resumed. Married women, women entering the study area within 4 months of marriage were also eligible to join the cohort under pregnancy surveillance. And when identified as pregnant, enrolled. Women who never got pregnant during the study period, permanently moved from the study area, were sterilised, reported menopause, divorce or death of husband, refused to participate or participation status unknown, died before detecting a pregnancy, had a pregnancy outcome after October 12, 2006, reported a last menstrual period after January 5, 2006 or had an unknown date of a last menstrual cycle were excluded.

### Interventions

1. Vitamin A (consisting of 7000 mcg retinol equivalents, or 23,300 IU, of VA palmitate in soybean oil with a small amount of vitamin E as an antioxidant)
2. β-carotene at 42 mg-an amount equivalent to 7000 mcg REs administered weekly from the time of pregnancy enrolment until 3 months postpartum.
3. Control group received placebo (consisting of soybean oil with a small amount of vitamin E as an antioxidant)

All 3 supplements contained 5 IU vitamin E in oil.

### Outcomes

1. All-cause pregnancy-related mortality.
2. Fetal loss due to miscarriage or stillbirth.
3. Infant mortality under 3 months of age.
5. Infant infectious morbidity.
7. Fetal and infant growth.
8. Prematurity.
10. Postnatal infant growth to 3 months of age.

### Notes

Women’s risk of spontaneous and recurrent miscarriage was unclear.

Women’s nutritional status: a 1-week history of diet at trial entry and 12 weeks after pregnancy was reported.
Sample size estimation: based on a 35% or greater reduction in all-cause mortality
Primary outcomes were compared on an intent-to-treat basis.
Compliance: adherence to supplementation was comparable across groups, with approximately 80% of women having directly consumed (under staff supervision) at least 64% of eligible supplements
Location: Bangladesh.
Timeframe: August 2001 to January 2007.

### Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>Sectors were randomised in blocks of 9. 3 sets of 3 identical coins on which the numbers 1, 2 or 3 were written were placed into a container, mixed and removed randomly, without replacement, and the 3-digit code of each sector was read aloud sequentially. Pg 7 (Labrique)</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>Study supplements were identical in colour, taste and external appearance. Supplements were originally shipped in identically-labelled, 100-count white opaque plastic bottles distinguished only by the code numbers 1, 2 or 3-listed on the label. A senior administrative staff (not involved in field activities) recoded the supplement bottles with sector-specific permanent stickers bearing codes from 001 to 596. Treatment codes were assigned by the Nutrilite Health Institute and were kept in a sealed envelope in a locked cabinet at both Nutrilite and Johns Hopkins University. Pg 7-8 (Labrique)</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias) All outcomes</td>
<td>Low risk</td>
<td>Study participants, interviewers, field supervisors and investigators remained masked to treatment assignments until the end of the trial. Pg 8 (Labrique)</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias) All outcomes</td>
<td>Low risk</td>
<td>Investigators remained masked to treatment assignments until the end of the trial. Pg 8 (Labrique)</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias) All outcomes</td>
<td>Low risk</td>
<td>Number of missing data among each group are small and balanced with detailed reason</td>
</tr>
<tr>
<td>Selective reporting (reporting bias) All outcomes</td>
<td>Low risk</td>
<td>Registration of trial and outcomes reported as described previously</td>
</tr>
</tbody>
</table>
Other bias | Low risk | The study appears to be free of other sources of bias.
--- | --- | ---

**West 2014**

**Methods**

Generation of random number sequence: 596 sectors of comparable size were used as units of randomisations. "We stratified the contiguous list of 596 sectors into 74 blocks of 8 each (for the first 592 sectors), plus a 75th block of 4 sectors. A sector-supplement code key with A or B was created by flip of a coin, reflecting assignment to iron-folic acid or MM supplementation, and duplicated, and each of 2 copies was sealed into an envelope by an uninvolved colleague at Johns Hopkins. Computer program randomised sectors within blocks to 1 of 2 codes such that each permutation had an equal probability of being chosen. The resulting 2 lists of sectors were securely transmitted to field headquarters. 1 envelope with the code key was securely transmitted to the supplement producer and the other sealed in an envelope and secured at Johns Hopkins. At no time during the trial did study investigators or field or data management staff have access to the key."

Blinding of outcome assessment: outcome assessors were blinded to the treatment allocation

Documentation of exclusion: 1351 pregnant women (3%) were excluded

Use of placebo control: no placebo, the control group received IFA supplementation

**Participants**

Participants in the study included 44,567 pregnant women and 28,518 infants in rural setting in Bangladesh. Pregnant women were recruited around ~ 10 weeks’ gestation. All women under 45 years of age who are married and living with their husbands as residents of the JiVitA study area at the time of an initial population enumeration, and those who enter the cohort as newlyweds, were eligible for pregnancy surveillance and, when identified as pregnant, enrolled. Women who permanently moved, were sterilised or menopausal, died, and did not get pregnant were excluded

**Interventions**

Weekly supplementation from enrolment (early pregnancy) to 3 months postpartum

Multiple-micronutrient: vitamin A (770 mcg retinol equivalents), vitamin D (5 mcg), vitamin E (15 mg), thiamin (1.4 mg), riboflavin (1.4 mg), niacin (1.4 mg), vitamin B12 (2.5 mg), vitamin B6 (1.9 mg), vitamin C (85 mg), zinc (12 mg), iodine (220 mcg), copper (1000 mcg), and selenium (60 mcg). The control group received IFA supplement (standard care)

**Outcomes**

1. Infant survival: determine the efficacy of a standard MMN supplement given to women daily during pregnancy through 12 weeks postpartum in lowering neonatal and infant mortality through 6 months of age by > 15% compared to the mortality of infants whose mothers receive daily iron + folic acid

2. Fetal and newborn outcomes: assess the efficacy of the MM intervention in reducing the:

   a) stillbirth rate by 23% or more, from an expected 30 to 24 or fewer stillbirths per 1000 births (i.e. live + still births);
   
   b) rate of preterm birth (live birth delivered < 37 weeks’ gestation) by 10% or more, from an expected 20% to 18% or fewer of all live births;
   
   c) prevalence of low birthweight (< 2500 g) by 5% or more, from an expected 40% to 38% or fewer of all live births; and
3. Other infant outcomes (to 6 months of age): assess the efficacy of the MM supplement intervention in improving:
   a) linear and ponderal growth, including its ability to protect lean body mass, that could lead to reduced prevalences of stunting, wasting and underweight status in infancy;
   b) infant morbidity, including diarrhoea and acute lower respiratory infection, in the first 3 months of life;
   c) micronutrient intake, represented by measured breast milk micronutrient concentrations;
   d) micronutrient status, and reducing prevalences of multiple deficiencies
4. Maternal outcomes: assess the efficacy of either MM supplement intervention in influencing among mothers the:
   a) prevalence of infectious morbidity, based on history, testing and signs, during pregnancy and in the 1st 6 months postpartum;
   b) rates of potentially fatal ("near miss") obstetric complications

Notes
Women’s risk of spontaneous and recurrent miscarriage: previous fetal loss and infant death reported
Women’s nutritional status is unclear.
Sample size estimation: live-born infants based on a 15% or greater reduction in 6-month mortality; pregnancies based on a 30% loss from induced abortion, miscarriage, and stillbirth
Intent-to-treat analysis of all outcomes.
Compliance: adherence was high, with half the women in both groups estimated to consume a median of approximately 95% of all distributed supplements; 80% in both groups consumed more than 80% of their intended tablets
Location: Bangladesh.
Timeframe: 2008 to 2012.

Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors' judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>“We used an in-house program (VBScript, Microsoft) that recognized 70 possible permutations for n = 8 sectors and k = 2 supplement allocations and 6 for the last block of n = 4 sectors. Using this program, we randomised sectors within blocks to 1 of 2 codes such that each permutation had an equal probability of being chosen”. Pg 2650</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>“Supplements were identical in appearance. Tablets were packed into opaque plastic bottles, affixed with codes A or B representing supplement content, and shipped to the field where logistics staff, uninvolved in the study, relabeled bottles with sector numbers(001-596) according to the ran-</td>
</tr>
</tbody>
</table>
**Blinding of participants and personnel (performance bias)**

| All outcomes | Low risk | Double-blinded and blinding of the participants is unlikely to have been broken |

**Blinding of outcome assessment (detection bias)**

| All outcomes | Low risk | At no time during the trial did study investigators or field or data management staff have access to the key. Pg 2651 |

**Incomplete outcome data (attrition bias)**

| All outcomes | Low risk | Numbers of missing outcome data in both group are small and balanced with similar reasons. Pg 2651, fig 1 |

**Selective reporting (reporting bias)**

| Low risk | Study protocol was available and all of the study’s pre-specified outcomes were reported in studies |

**Other bias**

| Low risk | The study appears to be free of other sources of bias. |

---

**Methods**

Generation of random number sequence: participants were randomised according to a computer-generated random number sequence. “Participants were assigned on an individual basis to antioxidant or placebo supplementation and remained in the same allocation group throughout the pregnancy and 2 weeks postpartum.” No details about how allocation concealment was done

Blinding of outcome assessment: unclear.

Documentation of exclusion: 6 women (5.5%) were excluded.

Use of placebo control: no placebo, the control group received non-antioxidant multiple micronutrients

**Participants**

110 women residing in Cipto Mangunkusumo National Hospital, Jakarta, Indonesia between 8 and 12 weeks of gestation. To be eligible, women had to have normal blood pressure at their first visit in pregnancy and again at trial entry. Women with multiple pregnancy, fetal anomaly, thrombophilia, infections, mola hydatidosa, chronic renal failure, uncontrolled hypertension, placental abnormalities, documented uterine bleeding within a week of screening, uterine malformation and history of medical and metabolic complication, such as heart disease or diabetes were excluded

**Interventions**

Antioxidant multiple micronutrient (MMN) supplement mixed into milk administered from trial entry until 2 weeks postpartum. Control received non-antioxidant (MMN) supplement and all women received 40 g of milk powder

**Outcomes**

1. Pre-eclampsia.
2. Fetal growth restriction.
3. HELLP syndrome.
Women’s risk of spontaneous and recurrent miscarriage: history of pre-eclampsia reported
Women’s nutritional status is unclear.
Sample size calculation: based on an expected incidence of pre-eclampsia in the control
group of at least 29% and in the treatment group of at least 30%
Analyses were performed on an intention-to-treat basis.
Compliance: unclear, no information reported.
Location: Jakarta, Indonesia.
Timeframe: June 2001 to December 2009.

Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>Participants were &quot;...randomized according to a computer-generated random number sequence&quot; pg 1153, pg2</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Unclear risk</td>
<td>“Participants were assigned on an individual basis to antioxidant or placebo supplementation.” No details of how allocation was done. pg 1153, pg2</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias) All outcomes</td>
<td>Low risk</td>
<td>“Treatment allocations were blinded to both investigator and the patient...” pg 1153, pg2</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias) All outcomes</td>
<td>Unclear risk</td>
<td>“Investigators blinded to the sample background” but not sure if they blinded outcomes of interest in current review</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias) All outcomes</td>
<td>Low risk</td>
<td>3/52 drop-out from supplementation group and 3/58 drop-out from control group. drop-out rate balanced in both groups</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Low risk</td>
<td>Primary and secondary outcomes reported as planned in the protocol</td>
</tr>
<tr>
<td>Other bias</td>
<td>Low risk</td>
<td>The study appears to be free of other sources of bias.</td>
</tr>
</tbody>
</table>
**Methods**

<table>
<thead>
<tr>
<th></th>
<th>Generation of random number sequence: double-blinded, multicentre trial. Randomisation was performed through an electronic data management platform, which enabled randomisation and data entry over the Internet. They were randomly allocated at a ratio of 1:1 to antioxidant supplementation (vitamins C and E) group or to placebo group through an electronic data management platform. Blinding of outcome assessment: outcome assessors were blinded to the treatment allocation. Documentation of exclusion: 277 women (10.5%) were excluded. Use of placebo control: placebo control.</th>
</tr>
</thead>
</table>

**Participants**

<table>
<thead>
<tr>
<th></th>
<th>2460 women participated in the study. Women were eligible for the trial if they were between 12 and 18 completed weeks of pregnancy on the basis of last menstrual period and confirmed by early ultrasound examination. The exclusion criteria were: (1) women who regularly consumed supplements 200 mg/day for vitamin C and/or 50 IU/day for vitamin E; (2) women who took warfarin; (3) women who had known fetal abnormalities; (4) women who had a history of medical complications; (5) women with repeated spontaneous abortion; (6) women who used an illicit drug during the current pregnancy</th>
</tr>
</thead>
</table>

**Interventions**

<table>
<thead>
<tr>
<th></th>
<th>Women were provided either with vitamins C and E or placebo. Total daily dose of vitamin C was 1000 mg, and that of vitamin E was 400 IU</th>
</tr>
</thead>
</table>

**Outcomes**

|---|---|

**Notes**

<table>
<thead>
<tr>
<th></th>
<th>Women’s risk of spontaneous and recurrent miscarriage: history of pre-eclampsia and gestational hypertension reported as well as obstetric history (abortion, stillbirth, low birthweight, preterm birth). Women’s nutritional status: use of supplements reported.</th>
</tr>
</thead>
</table>
Sample size calculation: based on an expected 4% and 15% incidence of the primary outcome in the low- and high-risk strata, respectively. Analyses were performed on an intention-to-treat basis.
Compliance: 85.5% in the vitamin group, 86.5% in the placebo group.
Location: multicentre trial in Canada (17 centres) and Mexico (10 centres).

### Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>They [women] were randomly allocated at a ratio of 1:1 to antioxidant supplementation (vitamins C and E) group or to placebo group through an electronic data management platform. Pg 239.e3</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Unclear risk</td>
<td>Although paper states that “Women in the placebo group were advised to take capsules that were identical in appearance to the active treatment capsules”, no details were provided on how they were allocated to treatment groups</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias) All outcomes</td>
<td>Low risk</td>
<td>None of the trial staff or any other person involved in the trial knew the treatment allocation for any women until after completion of the trial analysis.” Pg 239.e3 and 4</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias) All outcomes</td>
<td>Low risk</td>
<td>“None of the trial staff or any other person involved in the trial knew the treatment allocation for any women until after completion of the trial analysis.” Pg 239.e4</td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias) All outcomes</td>
<td>Unclear risk</td>
<td>A higher loss to follow-up was seen in the Mexican centres although the loss was balanced between treatment and placebo groups</td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Low risk</td>
<td>Outcomes reported according to the trial registration.</td>
</tr>
<tr>
<td>Other bias</td>
<td>Low risk</td>
<td>The study appears to be free of other sources of bias.</td>
</tr>
</tbody>
</table>
### Methods

Generation of random number sequence: cluster-randomised study wherein villages-not individuals-were randomly assigned to 1 treatment group or the other. The 2 supplements looked different but a coding system was adopted. “…packaged the supplements in boxes with identical labelling except for the supplement code”… “the code letter did not distinguish which supplement was used”

- **Blinding of outcome assessment:** unclear if outcome assessors were unaware of participants’ allocation to treatment or control group
- **Documentation of exclusion:** 768 women (20.9%) were excluded.
- **Use of placebo control:** no placebo, the control group received iron-folic acid supplementation

### Participants

| 3670 women from Maradi, rural Niger. Women who lived in 1 of the selected villages and who had experienced amenorrhoea for less than 12 weeks were eligible for participation. Exclusion criteria included women with night blindness and/or clinical signs of severe anaemia |

### Interventions

Daily multiple micronutrients consisting vitamin A 800 mcg, vitamin D 200 IU, vitamin E 10 mg, vitamin C 70 mg, vitamin B1 1.4 mg, vitamin B2 1.4 mg, vitamin B3 18 mg, vitamin B6 1.9, vitamin B12 2.6 mg, folic acid 400 mcg, iron 30 mg, zinc 15 mg, copper 2 mg, selenium 65 mcg, and iodine 150 mcg from enrolment until delivery. The control group received iron/folic acid

### Outcomes

Maternal outcomes:
1. Birth assistance.
2. Conditions of delivery.
5. Stillbirths.

Infant outcome:
1. Birthweight.

### Notes

- Women’s risk of spontaneous and recurrent miscarriage is unclear
- Women’s nutritional status is unclear.
- Sample size calculation: based on reduction of 25% in low birthweight
- Intention-to-treat analysis not performed.
- Compliance: at subsequent visits, the remaining capsules were counted, and the number of missing capsules was replenished for another month and noted in the particular booklet. Compliance with treatment in the intervention group was 79.2% ± 18.1% and in the control group 78.4% ± 18.5%
- Location: Maradi, Niger.

### Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
</table>
| Random sequence generation (selection bias) | Unclear risk       | “Villages-not individuals-were randomly assigned to one treatment group or the
### Zagre 2007

(Continued)

<table>
<thead>
<tr>
<th>Risk</th>
<th>Description</th>
</tr>
</thead>
<tbody>
<tr>
<td>Low</td>
<td>“Packaged the supplements in boxes with identical labeling except for the supplement code.”</td>
</tr>
<tr>
<td>Low</td>
<td>Health workers and midwives did not distinguish which supplement was used</td>
</tr>
<tr>
<td>Unclear</td>
<td>“Data collectors were informed that each supplement came in two sizes and colors, so that the code letter did not distinguish which supplement was used.”</td>
</tr>
<tr>
<td>Unclear</td>
<td>Proportion of women lost to follow-up was significantly different across groups 335/1777 in IFA and 290/L,893 in MMN. Detailed reasons for missing data not provided</td>
</tr>
<tr>
<td>Unclear</td>
<td>Many outcomes were assessed but report was limited to only 2 outcomes</td>
</tr>
<tr>
<td>High</td>
<td>There were differences in the baseline characteristics of participants. Women in control group tended to be poorer and less educated and more women in intervention group used more preventive measures against malaria and had larger households</td>
</tr>
</tbody>
</table>

### Zeng 2008

**Methods**

- Generation of random number sequence: randomisation of villages was stratified by county with a fixed ratio of treatments (1:1:1) and blocking of 15. The randomisation schedule was generated off site with a pseudo-random number generator. Allocation concealment was described “a treatment colour code was assigned to each village based on the treatment allocation schedule. The treatment codes were opened only once all data had been collected”
- Blinding of outcome assessment: unclear.
- Documentation of exclusion: 133 women (2.3%) were lost to follow-up, 279 women (4.8%) stopped taking the supplement and refused to continue to participate, and 601 women (10.3%) experienced fetal loss.
- Use of placebo control: no placebo, the control group received folic acid alone.

**Participants**

5828 women from Shaanxi Province of north west China between < 12 - 28 weeks' gestational age. Eligibility included all women resident in the counties who became pregnant between August 2002 and January 2006.

Women were ineligible if they:
1. were already taking supplements;  
2. had a serious illness;  
3. had an abnormal reproductive history;  
4. were planning;  
5. to work outside the area; or  
were more than 28 weeks pregnant.

### Interventions

Supplementation with MMN; IFA (60 mg of iron and 400 mcg of folic acid) or folic acid alone (400 mcg of folic acid) from enrolment until delivery

### Outcomes

1. Birthweight.  
2. Duration of gestation.  

### Notes

Women’s risk of spontaneous and recurrent miscarriage is unclear.  
Women’s nutritional status is unclear.  
Sample size calculation: based on a 25% reduction in low birthweight between either iron-folic acid or multiple micronutrient and folic acid (control) groups  
Compliance: the number of remaining capsules was reported by the village doctor who visited the women every 2 weeks. The level of compliance with the supplementation was high in all treatment groups  
Location: Shaanxi Province, China.  
Timeframe: August 2002 to January 2006.

### Risk of bias

<table>
<thead>
<tr>
<th>Bias</th>
<th>Authors’ judgement</th>
<th>Support for judgement</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random sequence generation (selection bias)</td>
<td>Low risk</td>
<td>“Randomisation of villages was stratified by county with a fixed ratio of treatments (1:1:1) and blocking of 15........ The randomisation schedule was generated off site with a pseudo-random number generator......” pg 2, pgh 3</td>
</tr>
<tr>
<td>Allocation concealment (selection bias)</td>
<td>Low risk</td>
<td>“A treatment colour code was assigned to each village based on the treatment allocation schedule. The treatment codes were opened only once all data had been collected.” pg 2, pgh 3</td>
</tr>
<tr>
<td>Blinding of participants and personnel (performance bias)</td>
<td>Low risk</td>
<td>“A treatment colour code was assigned to each village based on the treatment allocation schedule. The treatment codes were opened only once all data had been collected.” pg 2, pgh 3</td>
</tr>
<tr>
<td></td>
<td>Risk</td>
<td>Description</td>
</tr>
<tr>
<td>--------------------------------------</td>
<td>--------</td>
<td>-----------------------------------------------------------------------------</td>
</tr>
<tr>
<td>Blinding of outcome assessment (detection bias)</td>
<td>Unclear risk</td>
<td>No information provided about blinding of outcome assessors.</td>
</tr>
<tr>
<td>All outcomes</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Incomplete outcome data (attrition bias)</td>
<td>High risk</td>
<td>121/1666 missing in folic acid, 67/1537 missing in iron-folic acid, 88/1494 missing in multiple micronutrients in birthweight analysis. No reasons for missing data. Amount of missing data from other outcomes unknown. Fig 1; table 4</td>
</tr>
<tr>
<td>All outcomes</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Selective reporting (reporting bias)</td>
<td>Low risk</td>
<td>Study protocol is available and all primary outcomes were reported as described in the protocol</td>
</tr>
<tr>
<td>Other bias</td>
<td>Unclear risk</td>
<td>From the 531 clusters, 13 clusters were excluded due to no birth outcomes in the excluded clusters. The number of clusters excluded was unbalanced across the intervention groups which may have been due to important baseline differences</td>
</tr>
</tbody>
</table>

B-HCG: Beta human chorionic gonadotropin  
BMI: body mass index  
F: folic acid  
Hb: haemoglobin  
HbCC: haemoglobin C disease  
HbSC: haemoglobin SC disease  
HbSS: haemoglobin sickle cell disease  
HELLP syndrome: haemolysis, elevated liver enzymes, low platelet count syndrome  
HIV-1: Human Immunodeficiency Virus-1  
HOFPP: Hungarian Optimal Family Planning Programme  
IQR: interquartile range  
IFA: iron and folic acid  
IU: international units  
IVF-ET: in vitro fertilization and embryo transfer  
mcg: micrograms  
mg/mL: milligrams per millilitre  
MF: multivitamins with folic acid  
mg: milligrams  
MMN: multiple micronutrient  
MRDR: modified relative dose-response  
MV: multivitamins without folic acid  
MRC: Medical Research Council  
NTD: neural tube defect  
P: progesterone  
PAI-1: plasminogen activator inhibitor-1  
PAI-2: plasminogen activator inhibitor-2
PCV: packed cell volume  
sTfR: Soluble transferrin receptor  
UK: United Kingdom  
UNIMMAP: United Nations International Multiple Micronutrient Preparation  
USA: United States of America  
WBC: white blood cell

**Characteristics of excluded studies [ordered by study ID]**

<table>
<thead>
<tr>
<th>Study</th>
<th>Reason for exclusion</th>
</tr>
</thead>
<tbody>
<tr>
<td>Baumslag 1970</td>
<td>Onset of supplementation was &gt; 20 weeks' gestation. Women were supplemented with either iron, iron and folic acid or iron, folic acid and vitamin B12 from “after the 24th week of pregnancy”</td>
</tr>
<tr>
<td>Biswas 1984</td>
<td>Unclear of the gestational age at which women entered the trial</td>
</tr>
<tr>
<td>Blot 1981</td>
<td>Onset of supplementation was &gt; 20 weeks' gestation. Supplementation with either iron and folic acid or iron alone occurred “at the end of the 6th month of pregnancy”. Unclear if women were randomised to the treatment groups</td>
</tr>
<tr>
<td>Chanarin 1968</td>
<td>Onset of supplementation was &gt; 20 weeks' gestation. Women were given a folic acid supplement after the 20th week of pregnancy. Abortion was reported according to folic acid status at 15 weeks, prior to supplementation</td>
</tr>
<tr>
<td>Chelchowska 2004</td>
<td>No relevant clinical outcomes, reports biochemical markers of antioxidant status only</td>
</tr>
<tr>
<td>Christian 2003</td>
<td>Data for main outcomes of interest were not reported in a form that could be included in the analysis. Perinatal death was reported which included stillbirths (gestational age ≥ 28 weeks) and death among infants in the first 7 days of life. Miscarriage data (defined as fetal loss &lt; 28 weeks' gestation age) was not provided</td>
</tr>
<tr>
<td>Colman 1974</td>
<td>Onset of supplementation was &gt; 20 weeks' gestation. Women were supplemented “during the final month of pregnancy”. Outcomes reported included folic acid red cell and serum folic acid concentration and haemoglobin concentration</td>
</tr>
<tr>
<td>Correia 1982</td>
<td>No main outcomes of interest reported. Outcomes presented in the study were fetal weight (birthweight) and placental weight. Women’s risk of spontaneous and recurrent miscarriage is unclear</td>
</tr>
<tr>
<td>Coutsoudis 1999</td>
<td>Onset of supplementation was &gt; 20 weeks' gestation. Women were given vitamin A and beta-carotene “during the third trimester of pregnancy”</td>
</tr>
<tr>
<td>Dawson 1962</td>
<td>Onset of supplementation was &gt; 20 weeks' gestation. Women were supplemented with folic acid &quot;on or after the 28th week&quot;. Group allocation was not done randomly. Reported outcomes include incidence of folic acid deficiency and megaloblastic anaemia, and haemoglobin concentration</td>
</tr>
<tr>
<td>Study (Year)</td>
<td>Onset of supplementation and outcomes reported</td>
</tr>
<tr>
<td>-------------------</td>
<td>-----------------------------------------------</td>
</tr>
<tr>
<td>Edelstein 1968</td>
<td>Onset of supplementation was &gt; 20 weeks' gestation. Supplementation was started at the 28th week of pregnancy. Outcomes reported included serum folate activity and serum folate, urinary formiminoglutamic acid, serum vitamin B12, mean haemoglobin and haematocrit values.</td>
</tr>
<tr>
<td>Ferguson 1955</td>
<td>Only 24 (9%) of the 269 women in the trial began to participate before 15 weeks' gestation and outcomes not reported separately according to gestation at enrolment.</td>
</tr>
<tr>
<td>Feyi-Waboso 2005</td>
<td>Onset of supplementation was 20 or more weeks' gestation.</td>
</tr>
<tr>
<td>Fletcher 1971</td>
<td>No inclusion/exclusion criteria reported, unclear of gestational age at enrolment to the study, reports combined outcomes for &quot;antepartum and threatened or complete abortion&quot; and &quot;stillbirth or neonatal death or congenital malformation&quot; (not reported separately).</td>
</tr>
<tr>
<td>Giles 1971</td>
<td>Onset of supplementation was &gt; 20 weeks' gestation for a large proportion of the participants. 4 groups in the study, 2 of which involved supplementation after 20 weeks' gestation. Results were not reported separately between groups.</td>
</tr>
<tr>
<td>Hampel 1974</td>
<td>Unclear of the gestational age at which women entered the trial.</td>
</tr>
<tr>
<td>Hankin 1966</td>
<td>No main outcomes reported. Supplementation was from &quot;approximately 20 weeks&quot;, no clinically relevant outcomes, outcomes relating to vitamin C status in plasma and breast milk reported.</td>
</tr>
<tr>
<td>Hekmatdoost 2011</td>
<td>The intervention assessed the effectiveness of different forms of the same vitamin - folate (folic acid vs 5-methyltetrahydrofolate (MTHF)) against each other. There was no placebo or control group.</td>
</tr>
<tr>
<td>Hibbard 1969</td>
<td>No main outcomes reported. Biochemical measures of blood folate status reported.</td>
</tr>
<tr>
<td>Hunt 1984</td>
<td>All women received a multivitamin in addition to the zinc supplement or placebo.</td>
</tr>
<tr>
<td>Huybregts 2009</td>
<td>Both groups received a multivitamin supplement (same vitamin content in each group).</td>
</tr>
<tr>
<td>Kaestel 2005</td>
<td>Women were recruited until late pregnancy and onset of supplementation occurred after 20 weeks.</td>
</tr>
<tr>
<td>Laurence 1981</td>
<td>No main outcomes or pregnancy loss outcomes reported. Miscarriage reported in those women where there was a neural tube defect, but not in all women according to treatment group.</td>
</tr>
<tr>
<td>Lin 2010</td>
<td>Intervention assessed effect of nutritional supplement besides vitamins. No main outcomes reported.</td>
</tr>
<tr>
<td>Lira 1989</td>
<td>No main outcomes reported. Biochemical measures of iron and folate status reported.</td>
</tr>
<tr>
<td>Lumeng 1976</td>
<td>Unclear gestational age at enrolment, 5 women were excluded due to abortion, premature labour, inadequate dietary records or missing more than 3 prenatal visits. Exclusions were not reported by group allocation. Outcomes related to maternal and fetal plasma levels of pyridoxal 5'-phosphate and coenzyme saturation of aspartate aminotransferase and alanine aminotransferase in maternal erythrocytes were reported.</td>
</tr>
<tr>
<td>Study Year</td>
<td>Description</td>
</tr>
<tr>
<td>------------</td>
<td>-------------</td>
</tr>
<tr>
<td>1981</td>
<td>Onset of supplementation was &gt; 20 weeks’ gestation. Women were supplemented with vitamin D “throughout the 3rd trimester”</td>
</tr>
<tr>
<td>1987</td>
<td>No clinical outcomes reported. Maternal plasma concentrations of trophoblastic protein SP1 were reported</td>
</tr>
<tr>
<td>1965</td>
<td>Onset of supplementation was &gt; 20 weeks’ gestation. Women were supplemented with either iron or iron and folic acid, or iron, folic acid and vitamin B12. Supplementation was started after the 24th week of pregnancy</td>
</tr>
<tr>
<td>2002</td>
<td>No main outcomes reported. Women were enrolled in either early or late pregnancy. Biochemical measures of biotin status reported</td>
</tr>
<tr>
<td>2002</td>
<td>Unclear if this is a cohort study or randomised trial. Women in this study were participating in a randomised placebo-controlled trial of calcium supplementation, and completed a dietary assessment at 12-21 weeks’ gestation and 29-31 weeks’ gestation. Unclear whether all women took a standard prenatal multivitamin or just women in the placebo group. Results are presented according to “teens”, “twins” and “singleton” pregnancies, not according to whether women took the supplement or not. Outcomes reported included dietary intakes of vitamin C and E (with and without the contribution of the prenatal vitamin supplement)</td>
</tr>
<tr>
<td>1966</td>
<td>Onset of supplementation was &gt; 20 weeks’ gestation. Women supplemented with oral vitamin K1 “several days before delivery”</td>
</tr>
<tr>
<td>2014</td>
<td>Study was a food intervention trial.</td>
</tr>
<tr>
<td>1985</td>
<td>Unclear about content of vitamin supplements. Women were supplemented with high or low ‘bulk’ dietary supplements with vitamins added; however, the vitamin supplements added were not specified</td>
</tr>
<tr>
<td>1984</td>
<td>Unclear of gestation at enrolment to the trial. No pregnancy loss outcomes reported.</td>
</tr>
<tr>
<td>2001</td>
<td>No main outcomes reported. Women enrolled between 18 and 28 weeks’ gestation, no clinical outcomes reported, only haemoglobin and plasma erythropoietin concentrations</td>
</tr>
<tr>
<td>2002</td>
<td>Both groups received a multivitamin (same vitamin content in both groups)</td>
</tr>
<tr>
<td>1981</td>
<td>Non-randomised study of periconceptional multivitamin supplementation for the prevention of neural tube defects</td>
</tr>
<tr>
<td>1993</td>
<td>No main outcomes reported. Anaemic pregnant women were enrolled between 16 and 24 weeks’ gestation. The only clinical outcome reported was the percentage of women with anaemia following treatment with a combination of vitamin A and iron or placebo</td>
</tr>
<tr>
<td>Study</td>
<td>Results/Outcomes</td>
</tr>
<tr>
<td>------------------</td>
<td>------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------</td>
</tr>
<tr>
<td>Tanumihardjo 2002</td>
<td>No main outcomes reported. Mean gestation at enrolment was 17.6 weeks, no clinical outcomes reported, markers of vitamin A and iron status reported</td>
</tr>
<tr>
<td>Taylor 1982</td>
<td>No main outcomes reported. Haematological features and serum ferritin were determined. No reports on pregnancy loss</td>
</tr>
<tr>
<td>Thauvin 1992</td>
<td>No main outcomes reported. Women were supplemented from 3 months’ gestation, data on pregnancy outcomes including spontaneous abortion were collected but not reported</td>
</tr>
<tr>
<td>Trigg 1976</td>
<td>Unclear of gestation at enrolment to the trial.</td>
</tr>
<tr>
<td>Ulrich 1999</td>
<td>Non-randomised study. Observational cohort study of folic acid users, randomised to different doses of folic acid, but no controls</td>
</tr>
<tr>
<td>Villamor 2002</td>
<td>No main outcomes reported. Women enrolled between 12 and 27 weeks’ gestation, no pregnancy loss or main outcomes reported, only reports measures of weight gain during pregnancy</td>
</tr>
<tr>
<td>Vutyanovich 1995</td>
<td>No main outcomes reported. Women were enrolled in the study if they were less than 17 weeks’ gestation; however, no pregnancy loss or main outcomes were reported, only measures of nausea and vomiting</td>
</tr>
<tr>
<td>Wehby 2012</td>
<td>This trial involved only one vitamin compared at different doses</td>
</tr>
<tr>
<td>Young 2015</td>
<td>No relevant outcomes were reported in this study.</td>
</tr>
</tbody>
</table>

**Characteristics of studies awaiting assessment** [ordered by study ID]

**Adu-Afarwuah 2015**

<table>
<thead>
<tr>
<th>Methods</th>
<th>Partially double-blind, individually-randomised controlled trial</th>
</tr>
</thead>
<tbody>
<tr>
<td>Participants</td>
<td>1320 pregnant women around 20 gestational weeks.</td>
</tr>
<tr>
<td>Interventions</td>
<td>Women received 1 of 3 supplements daily until delivery: 60 mg Fe + 400 μg folic acid (IFA), or 1-2 recommended dietary allowances of 18 micronutrients including 20 mg Fe (MMN), or SQ-LNS with the same nutrient levels as in MMN, plus 4 additional minerals as well as macronutrients contributing 118 kcal (LNS)</td>
</tr>
<tr>
<td>Outcomes</td>
<td>Haemoglobin concentration (g/L) and 2 markers of iron status, zinc protoporphyrin (ZPP, μmol/mol heme) and transferrin receptor (TfR, mg/L) were assessed</td>
</tr>
</tbody>
</table>
### Adu-Afarwuah 2015  
(Continued)

| Notes | Only study abstract available and the composition of the supplement is not clear. Need to see full text |

### Agarwal 2012a

<table>
<thead>
<tr>
<th>Methods</th>
<th>Individually-randomised controlled trial.</th>
</tr>
</thead>
<tbody>
<tr>
<td>Participants</td>
<td>50 women of recurrent of abortions (more than 2).</td>
</tr>
<tr>
<td>Interventions</td>
<td>Women received Vitamin B12 1500 mcg, Vitamin B6 10 mg and folic acid 5 mg daily throughout pregnancy</td>
</tr>
</tbody>
</table>
| Outcomes | 1. Abortion.  
2. Pre-eclampsia.  
3. IUGR.  
4. Preterm labour. |
| Notes | Only study abstract available, comparison group also received same intervention but duration is different |

### Frenzel 1956

<table>
<thead>
<tr>
<th>Methods</th>
<th>Unclear.</th>
</tr>
</thead>
<tbody>
<tr>
<td>Participants</td>
<td>Unclear.</td>
</tr>
<tr>
<td>Interventions</td>
<td>Unclear.</td>
</tr>
<tr>
<td>Outcomes</td>
<td>Unclear.</td>
</tr>
<tr>
<td>Notes</td>
<td>A copy of the paper could not be located.</td>
</tr>
</tbody>
</table>

### Prado 2015

<table>
<thead>
<tr>
<th>Methods</th>
<th>Randomised cohorts.</th>
</tr>
</thead>
<tbody>
<tr>
<td>Participants</td>
<td>Pregnant women and infants.</td>
</tr>
<tr>
<td>Interventions</td>
<td>Various formulations of lipid-based nutrient supplements (LNS) during pregnancy and infancy or infancy alone</td>
</tr>
</tbody>
</table>
| Outcomes | 1. Association between linear growth and language development  
2. Association between 18-month language scores and  
3. 18-month length-for-age z-score (LAZ)  
4. LAZ at age 6 or 9 months and  
5. The change in LAZ between infancy (6 or 9 months) and 18 months |
| Notes | Abstract only, no specification what LNS is in abstract. |

IFA: iron and folic acid
### Characteristics of ongoing studies [ordered by study ID]

**Johns 2004**

<table>
<thead>
<tr>
<th>Trial name or title</th>
<th>The effect of antioxidant supplementation on women with threatened miscarriage</th>
</tr>
</thead>
<tbody>
<tr>
<td>Methods</td>
<td>Randomised controlled trial.</td>
</tr>
<tr>
<td>Participants</td>
<td>580 women who present with first trimester bleeding.</td>
</tr>
<tr>
<td>Interventions</td>
<td>Vitamin C 1000 mg and Vitamin E 400 IU versus placebo.</td>
</tr>
<tr>
<td>Outcomes</td>
<td>1. Incidence of miscarriage.</td>
</tr>
<tr>
<td></td>
<td>2. Late miscarriage.</td>
</tr>
<tr>
<td></td>
<td>3. Preterm labour.</td>
</tr>
<tr>
<td></td>
<td>4. Preterm pre-labour rupture of the membranes.</td>
</tr>
<tr>
<td></td>
<td>5. Fetal growth restriction.</td>
</tr>
<tr>
<td></td>
<td>6. Pre-eclampsia.</td>
</tr>
<tr>
<td>Starting date</td>
<td>01/03/2004.</td>
</tr>
</tbody>
</table>
| Contact information | **Dr Jemma Johns**  
UCLH/UCL Research & Development Governance Committee  
Research and Development Directorate  
University College London Hospitals NHS Trust  
1st Floor, Maple House  
149 Tottenham Court Road |
| Notes               | Listed as completed.                                                         |
## DATA AND ANALYSES

### Comparison 1. Vitamin C plus vitamin E versus placebo

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 Total fetal loss</td>
<td>7</td>
<td>18949</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.14 [0.92, 1.40]</td>
</tr>
<tr>
<td>2 Early or late miscarriage</td>
<td>4</td>
<td>13346</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>0.90 [0.65, 1.26]</td>
</tr>
<tr>
<td>3 Stillbirth</td>
<td>7</td>
<td>21442</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.31 [0.97, 1.76]</td>
</tr>
<tr>
<td>4 Congenital malformations</td>
<td>5</td>
<td>8334</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.17 [0.84, 1.62]</td>
</tr>
<tr>
<td>5 Any adverse effects of vitamin supplementation sufficient to stop supplementation</td>
<td>1</td>
<td>739</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.16 [0.39, 3.41]</td>
</tr>
</tbody>
</table>

### Comparison 2. Vitamin C versus no supplement/placebo

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 Total fetal loss</td>
<td>2</td>
<td>224</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.28 [0.58, 2.83]</td>
</tr>
<tr>
<td>2 Early or late miscarriage</td>
<td>2</td>
<td>224</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.17 [0.52, 2.65]</td>
</tr>
<tr>
<td>3 Stillbirth</td>
<td>1</td>
<td>200</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>3.0 [0.12, 72.77]</td>
</tr>
</tbody>
</table>

### Comparison 3. Vitamin C plus multivitamins versus placebo plus multivitamins or multivitamins alone

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 Total fetal loss</td>
<td>1</td>
<td>406</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.32 [0.63, 2.77]</td>
</tr>
<tr>
<td>2 Early or late miscarriage</td>
<td>2</td>
<td>790</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.19 [0.79, 1.79]</td>
</tr>
</tbody>
</table>

### Comparison 4. Vitamin A plus iron and folate versus iron and folate

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 Total fetal loss</td>
<td>3</td>
<td>1640</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>1.01 [0.61, 1.66]</td>
</tr>
<tr>
<td>2 Early or late miscarriage</td>
<td>2</td>
<td>1397</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>0.86 [0.46, 1.62]</td>
</tr>
<tr>
<td>3 Stillbirth</td>
<td>3</td>
<td>1640</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>1.29 [0.57, 2.91]</td>
</tr>
</tbody>
</table>
### Comparison 5. Vitamin A versus placebo

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 Total fetal loss</td>
<td>3</td>
<td>52480</td>
<td>Risk Ratio (Random, 95% CI)</td>
<td>1.05 [0.90, 1.23]</td>
</tr>
<tr>
<td>2 Early of late miscarriage</td>
<td>1</td>
<td>39668</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>0.98 [0.92, 1.04]</td>
</tr>
<tr>
<td>3 Stillbirth</td>
<td>1</td>
<td>39668</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>0.95 [0.86, 1.06]</td>
</tr>
</tbody>
</table>

### Comparison 6. Beta-carotene versus placebo

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 Total fetal loss</td>
<td>2</td>
<td>51163</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>1.02 [0.98, 1.07]</td>
</tr>
<tr>
<td>2 Early or late miscarriage</td>
<td>1</td>
<td>39860</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>1.00 [0.94, 1.06]</td>
</tr>
<tr>
<td>3 Stillbirth</td>
<td>1</td>
<td>39860</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>1.09 [0.98, 1.20]</td>
</tr>
</tbody>
</table>

### Comparison 7. Vitamin A or beta-carotene versus placebo

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 Total fetal loss</td>
<td>1</td>
<td>17373</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>1.05 [0.91, 1.21]</td>
</tr>
</tbody>
</table>

### Comparison 8. Vitamin A (with/without multivitamins) versus multivitamins or placebo

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 Total fetal loss</td>
<td>1</td>
<td>1074</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>0.80 [0.53, 1.21]</td>
</tr>
<tr>
<td>2 Early or late miscarriage</td>
<td>1</td>
<td>1075</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>0.76 [0.37, 1.55]</td>
</tr>
<tr>
<td>3 Stillbirth</td>
<td>1</td>
<td>1075</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>1.04 [0.60, 1.79]</td>
</tr>
</tbody>
</table>
Comparison 9. Multivitamin plus iron and folic acid versus iron and folic acid

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 Total fetal loss</td>
<td>10</td>
<td>94948</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>0.96 [0.93, 1.00]</td>
</tr>
<tr>
<td>2 Early or late miscarriage</td>
<td>10</td>
<td>94948</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>0.98 [0.94, 1.03]</td>
</tr>
<tr>
<td>3 Stillbirth</td>
<td>10</td>
<td>79851</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>0.92 [0.85, 0.99]</td>
</tr>
<tr>
<td>4 Congenital malformation</td>
<td>1</td>
<td>1200</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.0 [0.14, 7.08]</td>
</tr>
</tbody>
</table>

Comparison 10. Multivitamin without folic acid versus no multivitamin/folic acid

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 Total fetal loss</td>
<td>1</td>
<td>907</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>0.49 [0.34, 0.70]</td>
</tr>
<tr>
<td>2 Early or late miscarriage</td>
<td>1</td>
<td>907</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>0.89 [0.59, 1.34]</td>
</tr>
<tr>
<td>3 Stillbirth</td>
<td>1</td>
<td>907</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>0.14 [0.01, 2.77]</td>
</tr>
<tr>
<td>4 Congenital malformations</td>
<td>1</td>
<td>907</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.60 [0.53, 4.86]</td>
</tr>
</tbody>
</table>

Comparison 11. Multivitamin with/without folic acid versus no multivitamin/folic acid

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 Total fetal loss</td>
<td>1</td>
<td>1368</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>0.91 [0.65, 1.27]</td>
</tr>
<tr>
<td>2 Early or late miscarriage</td>
<td>1</td>
<td>1368</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>0.95 [0.67, 1.34]</td>
</tr>
<tr>
<td>3 Stillbirth</td>
<td>1</td>
<td>1368</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>0.33 [0.06, 1.98]</td>
</tr>
<tr>
<td>4 Congenital malformations</td>
<td>1</td>
<td>1368</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.99 [0.75, 5.26]</td>
</tr>
</tbody>
</table>

Comparison 12. Multivitamin plus folic acid versus folic acid

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 Total fetal loss</td>
<td>3</td>
<td>5012</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>1.04 [0.88, 1.23]</td>
</tr>
<tr>
<td>2 Early or late miscarriage</td>
<td>3</td>
<td>5012</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>0.97 [0.80, 1.18]</td>
</tr>
<tr>
<td>3 Stillbirth</td>
<td>3</td>
<td>4316</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>1.32 [0.93, 1.88]</td>
</tr>
<tr>
<td>4 Congenital malformations</td>
<td>2</td>
<td>1096</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.71 [0.72, 4.04]</td>
</tr>
</tbody>
</table>
### Comparison 13. Multivitamin without folic acid versus folic acid

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 Total fetal loss</td>
<td>2</td>
<td>1090</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>0.90 [0.62, 1.30]</td>
</tr>
<tr>
<td>2 Early or late miscarriage</td>
<td>2</td>
<td>1090</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>0.89 [0.61, 1.31]</td>
</tr>
<tr>
<td>3 Stillbirth</td>
<td>2</td>
<td>1090</td>
<td>Risk Ratio (Random, 95% CI)</td>
<td>0.99 [0.04, 22.90]</td>
</tr>
<tr>
<td>4 Congenital malformations</td>
<td>2</td>
<td>1090</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.61 [0.67, 3.85]</td>
</tr>
</tbody>
</table>

### Comparison 14. Multivitamin with/without folic acid versus folic acid

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 Total fetal loss</td>
<td>2</td>
<td>1644</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>0.96 [0.70, 1.33]</td>
</tr>
<tr>
<td>2 Early or late miscarriage</td>
<td>2</td>
<td>1644</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>0.96 [0.70, 1.33]</td>
</tr>
<tr>
<td>3 Stillbirth</td>
<td>2</td>
<td>1644</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>0.79 [0.15, 4.10]</td>
</tr>
<tr>
<td>4 Congenital malformations</td>
<td>2</td>
<td>1644</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.66 [0.76, 3.63]</td>
</tr>
</tbody>
</table>

### Comparison 15. Multivitamin with/without vitamin A versus vitamin A or placebo

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 Total fetal loss</td>
<td>1</td>
<td>1074</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>0.60 [0.39, 0.92]</td>
</tr>
</tbody>
</table>

### Comparison 16. Multivitamin versus control

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 Total fetal loss</td>
<td>1</td>
<td>5021</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>0.83 [0.58, 1.17]</td>
</tr>
<tr>
<td>2 Stillbirth</td>
<td>1</td>
<td>5021</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>0.83 [0.58, 1.17]</td>
</tr>
</tbody>
</table>
### Comparison 17. Multivitamin plus vitamin E versus multivitamin without vitamin E or control

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>Total fetal loss</td>
<td>1</td>
<td>823</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>0.92 [0.46, 1.83]</td>
</tr>
<tr>
<td>Early or late miscarriage</td>
<td>1</td>
<td>823</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>1.04 [0.26, 4.13]</td>
</tr>
<tr>
<td>Stillbirth</td>
<td>1</td>
<td>823</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>0.88 [0.39, 1.98]</td>
</tr>
</tbody>
</table>

### Comparison 18. Multivitamin plus folic acid versus no multivitamin/folic acid

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>Total fetal loss</td>
<td>3</td>
<td>6883</td>
<td>Risk Ratio (Random, 95% CI)</td>
<td>1.00 [0.75, 1.34]</td>
</tr>
<tr>
<td>Early or late miscarriage</td>
<td>3</td>
<td>6883</td>
<td>Risk Ratio (Random, 95% CI)</td>
<td>0.99 [0.72, 1.38]</td>
</tr>
<tr>
<td>Stillbirth</td>
<td>3</td>
<td>6883</td>
<td>Risk Ratio (Fixed, 95% CI)</td>
<td>1.04 [0.51, 2.10]</td>
</tr>
<tr>
<td>Congenital malformations</td>
<td>2</td>
<td>5777</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.69 [0.81, 3.53]</td>
</tr>
</tbody>
</table>

### Comparison 19. Folic acid plus multivitamin versus no folic acid/multivitamin

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>Total fetal loss</td>
<td>3</td>
<td>6883</td>
<td>Risk Ratio (M-H, Random, 95% CI)</td>
<td>1.00 [0.75, 1.34]</td>
</tr>
<tr>
<td>Early or late miscarriage</td>
<td>3</td>
<td>6883</td>
<td>Risk Ratio (M-H, Random, 95% CI)</td>
<td>0.99 [0.72, 1.38]</td>
</tr>
<tr>
<td>Stillbirth</td>
<td>3</td>
<td>6883</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.03 [0.51, 2.09]</td>
</tr>
<tr>
<td>Congenital malformations</td>
<td>2</td>
<td>5777</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.69 [0.81, 3.53]</td>
</tr>
</tbody>
</table>

### Comparison 20. Folic acid without multivitamin versus no folic acid/multivitamin

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>Total fetal loss</td>
<td>1</td>
<td>903</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>0.95 [0.64, 1.40]</td>
</tr>
<tr>
<td>Early or late miscarriage</td>
<td>1</td>
<td>903</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>0.97 [0.65, 1.44]</td>
</tr>
<tr>
<td>Stillbirth</td>
<td>1</td>
<td>903</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>0.67 [0.11, 4.02]</td>
</tr>
<tr>
<td>Congenital malformations</td>
<td>1</td>
<td>903</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.42 [0.45, 4.43]</td>
</tr>
</tbody>
</table>

Vitamin supplementation for preventing miscarriage (Review)

Copyright © 2016 The Cochrane Collaboration. Published by John Wiley & Sons, Ltd.
### Comparison 21. Folic acid with/without multivitamin versus no folic acid/multivitamin

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 Total fetal loss</td>
<td>1</td>
<td>1364</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>0.97 [0.69, 1.35]</td>
</tr>
<tr>
<td>2 Early or late miscarriage</td>
<td>1</td>
<td>1364</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>0.99 [0.70, 1.39]</td>
</tr>
<tr>
<td>3 Stillbirth</td>
<td>1</td>
<td>1364</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>0.67 [0.15, 2.96]</td>
</tr>
<tr>
<td>4 Congenital malformations</td>
<td>1</td>
<td>1364</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.90 [0.71, 5.04]</td>
</tr>
</tbody>
</table>

### Comparison 22. Folic acid plus multivitamin versus multivitamin

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 Total fetal loss</td>
<td>2</td>
<td>1102</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.15 [0.80, 1.67]</td>
</tr>
<tr>
<td>2 Early or late miscarriage</td>
<td>2</td>
<td>1102</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.16 [0.80, 1.69]</td>
</tr>
<tr>
<td>3 Stillbirth</td>
<td>2</td>
<td>1102</td>
<td>Risk Ratio (M-H, Random, 95% CI)</td>
<td>1.00 [0.04, 22.55]</td>
</tr>
<tr>
<td>4 Congenital malformations</td>
<td>2</td>
<td>1102</td>
<td>Risk Ratio (M-H, Random, 95% CI)</td>
<td>0.93 [0.28, 3.12]</td>
</tr>
</tbody>
</table>

### Comparison 23. Folic acid without multivitamin versus multivitamin

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 Total fetal loss</td>
<td>2</td>
<td>1090</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.12 [0.77, 1.62]</td>
</tr>
<tr>
<td>2 Early or late miscarriage</td>
<td>2</td>
<td>1090</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.12 [0.77, 1.64]</td>
</tr>
<tr>
<td>3 Stillbirth</td>
<td>2</td>
<td>1090</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>4.97 [0.58, 42.29]</td>
</tr>
<tr>
<td>4 Congenital malformations</td>
<td>2</td>
<td>1090</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>0.62 [0.26, 1.49]</td>
</tr>
</tbody>
</table>

### Comparison 24. Folic acid with or without multivitamin versus multivitamin

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 Total fetal loss</td>
<td>2</td>
<td>1644</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.14 [0.82, 1.57]</td>
</tr>
<tr>
<td>2 Early or late miscarriage</td>
<td>2</td>
<td>1642</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.09 [0.79, 1.51]</td>
</tr>
<tr>
<td>3 Stillbirth</td>
<td>2</td>
<td>1644</td>
<td>Risk Ratio (M-H, Random, 95% CI)</td>
<td>0.69 [0.02, 28.39]</td>
</tr>
<tr>
<td>4 Congenital malformations</td>
<td>2</td>
<td>1644</td>
<td>Risk Ratio (M-H, Random, 95% CI)</td>
<td>0.68 [0.19, 2.51]</td>
</tr>
</tbody>
</table>

Vitamin supplementation for preventing miscarriage (Review)  
Copyright © 2016 The Cochrane Collaboration. Published by John Wiley & Sons, Ltd.
### Comparison 25. Folic acid plus iron versus iron

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 Total fetal loss</td>
<td>1</td>
<td>75</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>0.23 [0.01, 4.59]</td>
</tr>
<tr>
<td>2 Early or late miscarriage</td>
<td>1</td>
<td>75</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>0.38 [0.02, 9.03]</td>
</tr>
<tr>
<td>3 Stillbirth</td>
<td>1</td>
<td>75</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>0.38 [0.02, 9.03]</td>
</tr>
</tbody>
</table>

### Comparison 26. Folic acid plus iron and antimalarials versus iron and antimalarials

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 Total fetal loss</td>
<td>1</td>
<td>160</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>13.0 [0.74, 226.98]</td>
</tr>
<tr>
<td>2 Early or late miscarriage</td>
<td>1</td>
<td>160</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>13.0 [0.74, 226.98]</td>
</tr>
</tbody>
</table>

### Comparison 27. Antioxidant vitamin supplementation

<table>
<thead>
<tr>
<th>Outcome or subgroup title</th>
<th>No. of studies</th>
<th>No. of participants</th>
<th>Statistical method</th>
<th>Effect size</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 Early or late miscarriage</td>
<td>1</td>
<td>110</td>
<td>Risk Ratio (M-H, Fixed, 95% CI)</td>
<td>1.12 [0.24, 5.29]</td>
</tr>
</tbody>
</table>

**WHAT’S NEW**

Last assessed as up-to-date: 6 November 2015.

<table>
<thead>
<tr>
<th>Date</th>
<th>Event</th>
<th>Description</th>
</tr>
</thead>
<tbody>
<tr>
<td>6 November</td>
<td>New citation required but conclusions have not changed</td>
<td>Conclusion unchanged.</td>
</tr>
<tr>
<td>6 November</td>
<td>New search has been performed</td>
<td>Search updated and 15 new trials included (Bhutta 2009; Hans 2010; Jauniaux 2004; McCance 2010; Prawirotomo 2011; Poston 2006; Roberts 2010; Sunawang 2009; Tofail 2008; West 2011; West 2014; Wibowo 2012; Xu 2010; Zagre 2007; Zeng 2008); an additional nine trials have been excluded (Christian 2003; Chelchowska 2004; Correia 1982; Hekmatdoost 2011; Kaestel 2005; Potdar 2014; Taylor 1982; Wehby 2012; Young 2015). Three trials are awaiting classi-</td>
</tr>
</tbody>
</table>
We have revised the outcomes to restrict the scope of the current update to look at miscarriage and miscarriage-related outcomes (Differences between protocol and review).

Six new authors joined the team for the 2015/2016 update: Erika Ota, Olukunmi Balogun, Katharina da Silva Lopes, Yo Takemoto, Mizuki Takegata and Rintaro Mori. Three authors who contributed to previous versions of this review stepped down for this update: Ning Pan, Caroline Crowther and Philippa Middleton.

### HISTORY

Protocol first published: Issue 1, 2003
Review first published: Issue 2, 2005

<table>
<thead>
<tr>
<th>Date</th>
<th>Event</th>
<th>Description</th>
</tr>
</thead>
<tbody>
<tr>
<td>27 August 2010</td>
<td>New citation required but conclusions have not changed</td>
<td>Substantive amendment and addition of a new author.</td>
</tr>
<tr>
<td>27 August 2010</td>
<td>New search has been performed</td>
<td>Search updated. 11 new studied included (Fawzi 2007; Fleming 1986; Osrin 2005; Roberfroid 2008; Rumbold 2006; Rumiris 2006; Spinnato 2007; Taylor 1982a; Summit 2008; Van den Broek 2006; Villar 2009), 3 studies excluded (Feyi-Wábosó 2005; Huybregts 2009; Shu 2002). Two new studies are awaiting classification (Chelchowska 2004b; Kubik 2004b) and two new ongoing trials were identified (Fall 2007b; Johns 2004; Sezikawa 2007).</td>
</tr>
<tr>
<td>20 September 2008</td>
<td>Amended</td>
<td>Converted to new review format.</td>
</tr>
</tbody>
</table>
CONTRIBUTIONS OF AUTHORS

Olukunmi Balogun: screened studies for inclusion/exclusion, data extraction, 'Risk of bias' assessment, data analysis, manuscript revision.

Kathrina da Silva Lopes: screened studies for inclusion/exclusion, data extraction, 'Risk of bias' assessment, manuscript revision.

Erika Ota: screened studies for inclusion/exclusion, statistical support, 'Summary of findings' tables, overall supervision.

Yo Takemoto: screened studies for inclusion/exclusion, data extraction, 'Risk of bias' assessment, manuscript revision.

Alice Rumbold: overall supervision.

Mizuki Takegata: data extraction, 'Risk of bias' assessment.

Rintaro Mori: statistical support, overall supervision.

DECLARATIONS OF INTEREST

Alice Rumbold is an investigator on the Australian Collaborative Trial of Supplements with vitamin C and vitamin E for the prevention of pre-eclampsia (Rumbold 2006). This trial is included in this review but its eligibility for inclusion, trial quality assessments and data extraction were carried out independently by two of the review authors not involved in the original trial.

Erika Ota: none known.

Olukunmi O Balogun: none known.

Katharina da Silva Lopes: none known.

Yo Takemoto: none known.

Mizuki Takegata: none known.

Rintaro Mori: Rintaro Mori's institution receives government funding from the Clinical Research Program for Child Health and Development, AMED, Japan to provide support for the PCG Satellite in Japan.

SOURCES OF SUPPORT

Internal sources

- The Grant of National Center for Child Health and Development 27B-10, 26A-5, Japan.

External sources

- Japan Agency for Medical Research and Development, Japan.

The National Research Center for Child Health and Development, Japan receives government funding (AMED No.27300101) from the Clinical Research Program for Child Health and Development, AMED, Japan to support the Cochrane Pregnancy and Childbirth Satellite in Japan.
DIFFERENCES BETWEEN PROTOCOL AND REVIEW

2011 update

We now include trials where supplementation occurred in mid-pregnancy. This was not specified in the original protocol for this review, but this was amended to be in line with other miscarriage reviews such as 'Progestogen for preventing miscarriage' (Haas 2009). We included trials where the onset of supplementation occurred both prior to and after 20 weeks' gestation, and when it could not be established whether the majority of the women started supplementation prior to 20 weeks' gestation. To overcome differences in the definition of miscarriage and stillbirth, we have used a combined outcome of total fetal loss (early or late miscarriage or stillbirth). We have still reported early or late miscarriage and stillbirth separately in addition to this combined outcome. Similarly, we specified in the original protocol that we would exclude studies reporting greater than 20% losses to follow-up. In this review we have included studies that reported more than 20% losses to follow-up and undertaken further analyses based on trial quality.

2016 update

The methods section was updated to current Cochrane Pregnancy and Childbirth standard text.

The scope of the current update has been restricted to look at miscarriage and miscarriage-related outcomes. Comparison 1 (any vitamins) and comparison 2 (sensitivity analysis) have been deleted in this update. After discussion it was decided that it did not make sense to compare any vitamin supplementation with no supplementation from either a clinical or consumer perspective.

We deleted the following primary outcomes.

For the woman

1. Placental abruption.
2. Pre-eclampsia.
3. Psychological effects (anxiety and depression) (previously included as maternal outcomes)

For the infant

1. Preterm birth (defined as birth less than 37 weeks' gestation).
2. Birthweight.
3. Small-for-gestational age (birthweight less than the third centile or the most extreme centile reported) (previously included as infant outcomes)

and secondary outcomes:

Secondary outcomes

1. Multiple pregnancy (including only trials supplementing women prior to or around the time of conception).
2. Very preterm birth (defined as less than 34 weeks' gestation).
3. Apgar score less than seven at five minutes.
4. Use of blood transfusion for the mother.
5. Anaemia (maternal and infant).
6. Placental weight.
7. Methods of feeding: breastfeeding, formula or both.
8. Subsequent fertility (subsequent pregnancy rate per couple or as defined by the authors).
9. Poor growth at childhood follow-up.
10. Disability at childhood follow-up.
13. Admission to neonatal intensive care unit.

All subgroups have been deleted from any analysis.
INDEX TERMS

Medical Subject Headings (MeSH)
Abortion, Habitual [prevention & control]; Abortion, Spontaneous [* prevention & control]; Antioxidants [administration & dosage]; Ascorbic Acid [administration & dosage]; Dietary Supplements [*adverse effects]; Folic Acid [administration & dosage]; Iron [administration & dosage]; Pre-Eclampsia [prevention & control]; Pregnancy Outcome; Pregnancy, Multiple; Prenatal Care; Randomized Controlled Trials as Topic; Stillbirth; Vitamin A [administration & dosage]; Vitamins [*administration & dosage; adverse effects]

MeSH check words
Female; Humans; Pregnancy